Error Exponent Bounds and Short-Length Coding Schemes for Distributed Hypothesis Testing (DHT)

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- 2 Considered source model
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- 4 Examples of source models
- 5 Practical short-length coding schemes for DHT
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Testing and Information Theory

• Hypothesis Testing is a standard problem in Statistics :

The probability distribution of a sequence $\mathbf{x}^n = (x_1, x_2, ..., x_n)$ is given by

$$\mathcal{H}_0: \mathbf{X}^n \sim P_{0\mathbf{X}}$$
$$\mathcal{H}_1: \mathbf{X}^n \sim P_{1\mathbf{X}}$$

• Statistician : how to optimally decide between \mathcal{H}_0 and \mathcal{H}_1 , by fully observing the data \mathbf{x}^n ?

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- Hypothesis Testing in Information Theory :

$$x^n$$
 Encoder R Decoder $H_0 | H_1$

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- Statistician : how to optimally decide between \mathcal{H}_0 and \mathcal{H}_1 , by fully observing the data \mathbf{x}^n ?
- Hypothesis Testing in Information Theory :

$$\xrightarrow{x^{n}} \text{Encoder} \xrightarrow{R} \text{Decoder} \xrightarrow{H_{0} \mid H_{1}}$$

• Information theorist : How to design the Encoder such that to optimally decide between \mathcal{H}_0 and \mathcal{H}_1 ?

Distributed Hypothesis Testing (DHT)



DHT Formulation

 $\mathcal{H}_0: (\mathbf{X}, \mathbf{Y}) \sim P_{0\mathbf{X}\mathbf{Y}}$ $\mathcal{H}_1: (\mathbf{X}, \mathbf{Y}) \sim P_{1\mathbf{X}\mathbf{Y}}$

Distributed Hypothesis Testing (DHT)



DHT Formulation

 $\mathcal{H}_0 : (\mathbf{X}, \mathbf{Y}) \sim P_{0\mathbf{X}\mathbf{Y}}$ $\mathcal{H}_1 : (\mathbf{X}, \mathbf{Y}) \sim P_{1\mathbf{X}\mathbf{Y}}$

- Encoder : $f^{(n)} : \mathcal{X}^n \longrightarrow \mathcal{M}_n = \{1, \dots, M\}$
- Rate-constraint : $\frac{\log M}{n} \leq R$,
- Decoder : $g^{(n)} : \mathcal{M}_n \times \mathcal{Y}^n \longrightarrow \mathcal{H} = \{H_0, H_1\}$

Performance criterion

• Type-I error probability

$$\alpha_{n} = \mathbb{P}\left[g^{\left(n\right)}\left(\boldsymbol{f}^{\left(n\right)}\left(\mathbf{X}^{n}\right), \mathbf{Y}^{n}\right) = H_{1} \mid H_{0} \text{ is true}\right],$$

• Type-II error probability

$$\beta_{n} = \mathbb{P}\left[g^{(n)}\left(f^{(n)}\left(\mathbf{X}^{n}\right), \mathbf{Y}^{n}\right) = H_{0} \mid H_{1} \text{ is true}\right]$$

Performance criterion

• Type-I error probability

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• Type-II error probability

$$\beta_{n} = \mathbb{P}\left[g^{(n)}\left(f^{(n)}\left(\mathbf{X}^{n}\right), \mathbf{Y}^{n}\right) = H_{0} \mid H_{1} \text{ is true}\right]$$

• Objective : For given $\alpha_n \leq \epsilon$, find the Type-II error exponent θ such that

$$\limsup_{n \to \infty} \frac{1}{n} \log \frac{1}{\beta_n} \ge \theta \tag{1}$$

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Achievable error exponent bounds for i.i.d. sources model

Quantization scheme [Ahlswede86] [HAN87]

- Ahlswede and al derived optimal error exponent for testing against independence.
- HAN improved it with joint typicality check at the encoder.

Quantize-binning scheme [SHA 94]

- Shimokowa et. al combined quantization with random binning (similar to Wyner-Ziv and Slepian-Wolf).
- It allows to exploit the correlation between the sources to reduce the compression rate
- It optimality was considered in [Rahman 2012], [Katz 2015], [Watanabe 2022]
- It was extended to various more complex setups [Salehkalaibar 18], [Sreekuma 19], [Escamilla2020]
- It has been recently improved by [Kochman2023]

However, it was limited only to i.i.d. sources model Objective : extend it to more general sources models, not necessarily i.i.d.

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Hypothesis Testing in Information Theory

2 Considered source model

③ Error exponent bounds for general sources

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We consider a more generic sources model [Han1998]

$$\mathbf{X} = \{X^n = (X_1, X_2, \cdots, X_n)\}_{n=1}^{\infty} \text{ and } \mathbf{Y} = \{Y^n = (Y_1, Y_2, \cdots, Y_n)\}_{n=1}^{\infty}$$

- \bullet The components of ${\bf X}$ and ${\bf Y}$ are not necessarily i.i.d
- Includes the previous i.i.d. models as particular instances.
- The objective is to derive more generic exponent error bounds
- We rely on Information-Spectrum approach [Han1998]

For a sequence $\{Z_n\}_{n=1}^{\infty}$

•
$$p - \limsup_{n \to \infty} Z_n = \inf \{ \alpha \mid \lim_{n \to +\infty} \mathbb{P}(Z_n > \alpha) = 0 \},\$$

•
$$p - \liminf_{n \to \infty} Z_n = \sup \{ \alpha \mid \lim_{n \to +\infty} \mathbb{P}(Z_n < \alpha) = 0 \}$$

Examples

For a pair $(\mathbf{U}^n, \mathbf{X}^n)$, sup and inf spectral mutual information :

•
$$\bar{I}(\mathbf{X}; \mathbf{U}) = p - \limsup_{n \to \infty} \frac{1}{n} \log \frac{P_{\mathbf{U}^n | \mathbf{X}^n}(\mathbf{U}^n | \mathbf{X}^n)}{P_{\mathbf{U}^n}(\mathbf{U}^n)}$$

• $\underline{I}(\mathbf{X}; \mathbf{U}) = p - \liminf_{n \to \infty} \frac{1}{n} \log \frac{P_{\mathbf{U}^n | \mathbf{X}^n}(\mathbf{U}^n | \mathbf{X}^n)}{P_{\mathbf{U}^n}(\mathbf{U}^n)}$

When \mathbf{U} and \mathbf{X} are i.i.d., we can show that

•
$$\overline{I}(\mathbf{X}; \mathbf{U}) = \underline{I}(\mathbf{X}; \mathbf{U}) = I(X; U)$$

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Theorem

$$\theta \leq \sup_{P_{\mathbf{U}|\mathbf{X}}} \left\{ \min \left\{ \theta_{test}, \theta_{bin} \right\} \right\},$$
(2)

•
$$\theta_{bin} = r - (\overline{I}(\mathbf{X}; \mathbf{U}) - \underline{I}(\mathbf{U}; \mathbf{Y})), r \leq \mathsf{R}$$

• $\theta_{test} = \underline{D} (P_{0\mathbf{U}\mathbf{Y}} || P_{1\mathbf{U}\mathbf{Y}}) + (\underline{I}(\mathbf{X}; \mathbf{U}) - \overline{I}(\mathbf{X}; \mathbf{U}))$

The error exponent is a trade-off between a binning error and a testing error.

•
$$\overline{I}(\mathbf{X}; \mathbf{U}) = p - \limsup_{n \to \infty} \frac{1}{n} \log \frac{P_{0\mathbf{U}^n | \mathbf{X}^n}(\mathbf{U}^n | \mathbf{X}^n)}{P_{0\mathbf{U}^n}(\mathbf{U}^n)}$$

• $\underline{D}(P_{0\mathbf{U}\mathbf{Y}} \| P_{1\mathbf{U}\mathbf{Y}}) = p - \liminf_{n \to \infty} \frac{1}{n} \log \frac{P_{0\mathbf{U}^n \mathbf{Y}^n}(\mathbf{U}^n, \mathbf{Y}^n)}{P_{1\mathbf{U}^n \mathbf{Y}^n}(\mathbf{U}^n, \mathbf{Y}^n)}.$

Codebook generation

- Generate $e^{n\mathsf{R}}$ sequences \mathbf{u}^n according to a fixed distribution $P_{\mathbf{U}^n|\mathbf{X}^n}$
- Distribute them uniformly in e^{nr} bins

Encoder

- Search a sequence \mathbf{u}^n such that $(\mathbf{u}^n, \mathbf{x}^n) \in T_n^{(1)}$
- Send the index of the bin to which \mathbf{u}^n belongs
- Otherwise, send a message error to the decoder to simply declare \mathcal{H}_1

$$T_n^{(1)} = \left\{ (\mathbf{x}^n, \mathbf{u}^n) \text{ s.t. } \frac{1}{n} \log \frac{P_{\mathbf{U}^n \mid \mathbf{X}^n} \left(\mathbf{u}^n \mid \mathbf{x}^n \right)}{P_{\mathbf{U}^n} \left(\mathbf{u}^n \right)} < r_0 - \varepsilon \right\}$$

Achievable coding scheme

Decoder

From the received bin index and side information \mathbf{y}^n :

- Pick $\hat{\mathbf{u}}^n$ such that $(\hat{\mathbf{u}}^n, \mathbf{y}^n) \in T_n^{(2)}$
- Decide \mathcal{H}_0 if $(\hat{\mathbf{u}}^n, \mathbf{y}^n)$ belongs to an acceptance region A_n
- Otherwise, decide \mathcal{H}_1 (also when a message error is received)

$$\begin{aligned} T_n^{(2)} &= \left\{ (\mathbf{y}^n, \mathbf{u}^n) \text{ s.t. } \frac{1}{n} \log \frac{P_{\mathbf{U}^n | \mathbf{Y}^n} (\mathbf{u}^n | \mathbf{y}^n)}{P_{\mathbf{U}^n} (\mathbf{u}^n)} < r_0 - \varepsilon \right\} \\ \mathsf{A}_n &= \left\{ (\mathbf{y}^n, \mathbf{u}^n) \text{ s.t. } \frac{P_{\mathbf{U}^n \mathbf{Y}^n} (\mathbf{u}^n, \mathbf{y}^n)}{P_{\overline{\mathbf{U}}^n \overline{\mathbf{Y}^n}} (\mathbf{u}^n, \mathbf{y}^n)} > S \right\}. \end{aligned}$$

- We can not rely on the method of types as in i.i.d. case
- Our achievability proof is information-spectrum based.
- Here $T_n^{(2)}$ and A_n are different that typical set as known in i.i.d. case

Proof outlines : Type-I error α_n analysis

•
$$E_{11} = \left\{ \nexists \mathbf{u}^n \text{ s.t. } (\mathbf{X}^n, \mathbf{u}^n) \in T_n^{(1)}, (\mathbf{Y}^n, \mathbf{u}^n) \in \mathcal{A}_n \right\},$$

•
$$E_{12} = \left\{ \exists \mathbf{u}'^n \neq \mathbf{u}^n \text{ s.t. } i(\mathbf{u}'^n) = i(\mathbf{u}^n) \text{but } (\mathbf{u}'^n, \mathbf{Y}^n) \notin \mathcal{A}_n \right\}.$$

• We show that : $\alpha_n \leq \mathbb{P}(E_{11}) + \mathbb{P}(E_{12})$

Information-Spectrum approach

• For $r_0 = \overline{I}(\mathbf{X}; \mathbf{U})$, and from the definition of $\overline{I}(\mathbf{X}; \mathbf{U})$, we show

$$\lim_{n \to \infty} \mathbb{P}\left((\mathbf{X}^n, \mathbf{U}^n) \notin T_n^{(1)} \right) = 0.$$

• When $S = \underline{D}(P_{\mathbf{U}\mathbf{Y}} || P_{\overline{\mathbf{U}\mathbf{Y}}})$ and from the definition of $\underline{D}(P_{\mathbf{U}\mathbf{Y}} || P_{\overline{\mathbf{U}\mathbf{Y}}})$, we have

$$\lim_{n\to\infty} \mathbb{P}\left(\left(\mathbf{Y}^n, \mathbf{U}^n \right) \notin \mathcal{A}_n \right) = 0.$$

• Thus, We show that : $\alpha_n \stackrel{n \to \infty}{\longrightarrow} 0$.

$$\begin{split} T_n^{(1)} &= \left\{ \left(\mathbf{x}^n, \mathbf{u}^n \right) \text{ s.t. } \frac{1}{n} \log \frac{P_{\mathbf{U}^n \mid \mathbf{X}^n \left(\mathbf{u}^n \mid \mathbf{X}^n \right)}}{P_{\mathbf{U}^n \left(\mathbf{u}^n \right)}} < r_0 - \varepsilon \right\} \\ A_n &= \left\{ \left(\mathbf{y}^n, \mathbf{u}^n \right) \text{ s.t. } \frac{P_{\mathbf{U}^n \mathbf{Y}^n} \left(\mathbf{u}^n, \mathbf{y}^n \right)}{P_{\mathbf{U}^n \mathbf{Y}^n} \left(\mathbf{u}^n, \mathbf{y}^n \right)} > S \right\}. \end{split}$$

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Proof outlines : Type-II error β_n analysis

•
$$E_{21} = \left\{ \exists \tilde{\mathbf{u}}^n \neq \mathbf{u}^n : i(\tilde{\mathbf{u}}^n) = i(\mathbf{u}^n), \mathsf{but}\left(\overline{\mathbf{Y}}^n, \tilde{\mathbf{u}}^n\right) \in \mathcal{A}_n \right\},$$

•
$$E_{22} = \left\{ (\mathbf{u}^n, \overline{\mathbf{Y}}^n) \in T_n^{(2)}, (\mathbf{u}^n, \overline{\mathbf{Y}}^n) \in \mathcal{A}_n \right\}.$$

• $\beta_n \leq \mathbb{P}(E_{21}) + \mathbb{P}(E_{22})$

Information-Spectrum approach

• When
$$r_0 = \overline{I}(\mathbf{X}; \mathbf{U}), r' = \underline{I}(\mathbf{Y}; \mathbf{U})$$
, and $S = \underline{D}(P_{\mathbf{U}\mathbf{Y}} || P_{\overline{\mathbf{U}\mathbf{Y}}})$

- $\bullet \ \beta_n \leq e^{-n\left(r \left(\overline{I}(\mathbf{X}; \mathbf{U}) \underline{I}(\mathbf{Y}; \mathbf{U})\right) \epsilon\right)} + e^{-n\left(\underline{I}(\mathbf{X}; \mathbf{U}) \overline{I}(\mathbf{X}; \mathbf{U}) + \underline{D}\left(P_{\mathbf{U}\mathbf{Y}} \| P_{\overline{\mathbf{U}\mathbf{Y}}}\right) 2\epsilon\right)}.$
- Since, $\limsup_{n\to\infty} \frac{1}{n} \log \frac{1}{\beta_n} \ge \theta$
- This shows that the error exponent θ in the theorem is achievable.

$$\begin{split} \mathsf{T}_{n}^{(2)} &= \bigg\{ (\mathbf{y}^{n}, \mathbf{u}^{n}) \text{ s.t. } \frac{1}{n} \log \frac{P_{\mathbf{U}^{n}} |\mathbf{y}^{n} (\mathbf{u}^{n} |\mathbf{y}^{n})}{P_{\mathbf{U}^{n}} (\mathbf{u}^{n})} > r' - \varepsilon \bigg\} \\ A_{n} &= \bigg\{ (\mathbf{y}^{n}, \mathbf{u}^{n}) \text{ s.t. } \frac{P_{\mathbf{U}^{n}} \mathbf{y}^{n} (\mathbf{u}^{n}, \mathbf{y}^{n})}{P_{\overline{\mathbf{U}}^{n}} \overline{\mathbf{y}^{n}} (\mathbf{u}^{n}, \mathbf{y}^{n})} > S \bigg\}. \end{split}$$

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Short-length nature of DHT

• For Binary i.i.d. sources, we compute $\beta_n = e^{-n\theta}$ as function of code length n



- For instance, for n = 100, $\beta_n = 10^{-12}$, for n = 50, $\beta_n = 10^{-6}$.
- This strongly suggests that practical schemes should focus on values of n < 50.
- We now introduce two practical coding schemes for such short sequence lengths
- So far, we only focus on Binary sources

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DHT for Binary sources model



• Source model : (X, Y) are such that Y = X + Z, where X and Z are independent and P(X = 1) = 0.2 and and P(Z = 1) = p

• Hypothesis definition :

$$\mathcal{H}_0: p = p_0,$$
$$\mathcal{H}_1: p = p_1.$$

- Decoder : Decide between \mathcal{H}_0 and \mathcal{H}_1 from \mathbf{Y}^n and a coded version of \mathbf{X}^n
- Objective : Design practical coding schemes for this setup

Implementation

• From the **Generator matrix** G of a linear block code, we calculate \mathbf{z}_q^m as :

$$\mathbf{z}_q^m = \arg\min_{\mathbf{z}^m} d\left(G\mathbf{z}^m, \mathbf{x}^n\right)$$

- At the receiver, we obtain the quantized vector $\mathbf{x}_q^n = G \mathbf{z}_q^m$
- The receiver performs the NP test over $(\mathbf{x}_q^n, \mathbf{y}^n)$

Scheme 1 : Binary Quantization

Performance²

$$\alpha_{n} = 1 - \frac{1}{N_{0}^{(q)}} \sum_{\lambda=0}^{\lambda_{q}} \sum_{\gamma=0}^{d_{\max}^{(q)}} \sum_{j=0}^{n} E_{\gamma}^{(q)} \Gamma_{\lambda,j,\gamma} p_{0}^{j} (1-p_{0})^{n-j}, \qquad (3)$$
$$\beta = \frac{1}{N_{0}^{(q)}} \sum_{\lambda=0}^{\lambda_{q}} \sum_{\gamma=0}^{d_{\max}} \sum_{j=0}^{n} E_{\gamma}^{(q)} \Gamma_{\lambda,j,\gamma} p_{1}^{j} (1-p_{1})^{n-j}, \qquad (4)$$

where for $j = \gamma + \lambda - 2u$ and $0 \le u \le \min(\gamma, \lambda) \le n$, $\Gamma_{\lambda, j, \gamma} = {\gamma \choose u} {n-\gamma \choose \lambda-u}$.

- $d_{\max}^{(q)}$ is the maximum hamming weight of words \mathbf{x}^n of a decision region C_0
- $E_{\gamma}^{(q)}$ is the number of words \mathbf{x}^n of Hamming weight γ , and $N_0^{(q)} = \sum_{\gamma=0}^{d_{\text{max}}^{(q)}} E_{\gamma}^{(q)}$
- One can optimize $E_{\gamma}^{(q)}$ to obtain a lower bound for practical DHT

^{2.} Dupraz, E., Adamou, I. S., Asvadi, R., Matsumoto, T. "APractical Short-Length Coding Schemes for Binary Distributed Hypothesis Testing", International Symposium in Information Theory (ISIT) 2024 (= >

Encoder

• Quantization : From the Generator matrix G of a linear block code, we calculate \mathbf{z}_q^m as :

$$\mathbf{z}_q^m = \arg\min_{\mathbf{z}^m} d\left(G\mathbf{z}^m, \mathbf{x}^n\right)$$

• Binning : From a parity check matrix H of another linear block code, we compute

• Send
$$\mathbf{u}^k$$
 at rate k/n $\mathbf{u}^k = H \mathbf{z}_q^m$

Decoder

Identify the vector z^m_q such that

$$\hat{\mathbf{z}}^m = rg\min_{\mathbf{z}_q^m} d\left(G\mathbf{z}_q^m, \mathbf{y}^n
ight)$$
 s.t. $H\mathbf{z}^m = \mathbf{u}^k$

- Compute $\hat{\mathbf{x}}^n = G\hat{\mathbf{z}}^m$
- Then, apply the NP test onto $(\hat{\mathbf{x}}^n, \mathbf{y}^n)$.

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Scheme 2 : Quantize-Binning

Performance

$$\alpha_n = 1 - \mathbb{P}_B(p_0) - \mathbb{P}_{\bar{B}}(p_0), \qquad (5)$$

$$\beta_n = \mathbb{P}_B(p_1) + \mathbb{P}_{\bar{B}}(p_1), \qquad (6)$$

where

$$\mathbb{P}_{B}(\delta) = \sum_{\nu=0}^{\min(d_{\max}^{(qb)}, \lambda_{qb})} \frac{E_{\nu}^{(qb)}}{\binom{n}{\nu}} \sum_{\gamma=0}^{d_{\max}^{(q)}} \frac{E_{\gamma}^{(q)}}{N_{0}^{(q)}} \sum_{j=0}^{n} \Gamma_{\nu, j, \gamma} \delta^{j} (1-\delta)^{n-j},$$
$$\mathbb{P}_{\bar{B}}(\delta) = \sum_{i=0}^{n} \left[\left(\sum_{\gamma=0}^{d_{\max}^{(q)}} \frac{E_{\gamma}^{(q)}}{N_{0}^{(q)}} \sum_{j=0}^{n} \Gamma_{i, j, w} \delta^{j} (1-\delta)^{n-j} \right) \left(\sum_{t=1}^{n} \sum_{\nu=0}^{\lambda_{qb}} \frac{E_{\nu}^{(qb)}}{\binom{n}{\nu}} \frac{A_{t}^{(qb)} \Gamma_{i, \nu, t}}{\binom{n}{i}} \right) \right]$$

- $E_{\gamma}^{(q)}$ is the number of words \mathbf{x}^n of Hamming weight γ
- the set $\{A_t^{(qb)}\}_{t\in [\![0,n]\!]}$ is the code weight distribution of the concatenated code (G and H)
- \bullet One can optimized $E_{\gamma}^{(q)}$ and $A_t^{(qb)}$ to obtain a lower bound for practical DHT

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Simulation results

- Baseline Truncation Scheme : we send l = 16 non-coded bits
- Quantization Scheme : BCH (31, 16)-code with $d_{min} = 7$. As a result m = 16 coded bits are sent.



• The quantization scheme performs better that the Truncation scheme

• The theoretical expressions are consistent with the Monte-Carlo simulations

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Simulation results

- Baseline Truncation scheme : we send l = 8 non-coded bits of \mathbf{x}^n
- Quantize-binning Scheme : BCH (31, 16)-code with $d_{min} = 7$ and Reed-Muller (16, 5)-code with $d_{min} = 8$. As a result k = 8 coded bits are sent



• The Quantize-binning scheme performs better that the Truncation scheme

• The theoretical expressions are consistent with the Monte-Carlo simulations

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Conclusion

- We derived a general expression of the Type-II error exponent for general sources that are not necessarily i.i.d.
- Our approach is information-spectrum based.
- We then proposed practical quantization scheme and quantize-binning scheme
- Our proposed schemes perform better than the baseline non-coded scheme
- We provided exact **analytical expressions** of Type-I and Type-II errors for the proposed schemes.

Current Works

- Add an empirical entropy check in our quantize-binning as in [Kochman2023]
- $\bullet\,$ Extend both theoretical and practical results to the results to the case where both ${\bf X}$ and ${\bf Y}$ are encoded

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