

Tracking Threshold Crossing Times of a Gaussian Random Walk Through Correlated Observations

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Outline

- 1 Problem formulation and background
- 2 Results
- 3 Conclusion

The Tracking Stopping Time problem

- $X = \{X_t\}_{t \geq 0}$ and stopping time τ on X
- Statistician has access to X through $Y = \{Y_t\}_{t \geq 0}$.
- Statistician wishes to find a stopping time η so that

$$\mathbb{E}|\eta - \tau|$$

is minimized.

Example: monitoring

- X_t : distance of an object from a barrier at time t
- τ : first time when $X_t = 0$
- Y_t : noisy measurements of X_t
- η : alarm time based on Y should be close to τ

Example: forecasting

- X_t : fatigue up to day t of a big manufacturing machine
- τ : first day t when X_t crosses critical fatigue threshold
- Machine replacement period: 10 days
- η : first day when new machine is operational

Wanted η close to τ because

- $\{\eta > \tau\}$: interrupted manufacturing process
- $\{\eta < \tau\}$: storage costs

⇒ TST problem with $Y_t = X_{t-10}$ if $t > 10$ and $X_t = 0$ else.

Example: change-point problem

- θ positive integers valued random variable
- $\{Y_t\}$ with $Y_t \sim P_0$ if $t < \theta$ and $Y_t \sim P_1$ if $t \geq \theta$
- Goal: find η such that $\mathbb{E}(\eta - \theta)_+$ is minimized while $\mathbb{P}(\eta < \theta) \leq \alpha$

Equivalent to TST problem:

- $\{X_t\}$ with $X_t = 0$ if $t < \theta$ and $X_t = 1$ if $t \geq \theta$.
- $\tau = \inf\{t \geq 1 : X_t = 1\}$ (i.e., $\tau = \theta$)
- Goal: find η such that $\mathbb{E}(\eta - \tau)_+$ minimized while $\mathbb{P}(\eta < \tau) \leq \alpha$

Change-point or tracking a stopping time?

C-P and TST are **not** equivalent:

for $k > t$

$$\text{C-P: } \mathbb{P}(\theta = k | Y^t = y^t, \theta > t) = \mathbb{P}(\theta = k | \theta > t)$$

$$\text{TST: } \mathbb{P}(\tau = k | Y^t = y^t, \tau > t) \neq \mathbb{P}(\tau = k | \tau > t) \text{ in general}$$

i.e., for the C-P problem,

if $\{\theta > t\}$, the first t samples Y^t are useless for predicting θ .

Bottom line

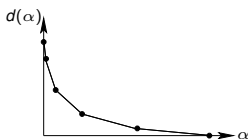
The TST problem formulation

- A useful generalization of the change-point problem.
- The change-point problem is known to be hard.
- Any hope?

A first step: algorithmic solution (Niesen and T, 2009)

Given finite alphabet process (X, Y) and $\tau \leq b$, the algorithm outputs

$$\bullet \quad d(\alpha) = \min_{\tau: \mathbb{P}(\eta < \tau) \leq \alpha} \mathbb{E}(\eta - \tau)_+ \quad \alpha \in [0, 1]$$



- the corresponding optimal stopping times.

Under conditions on (X, Y) and τ , the algorithm is $\text{poly}(b)$.

Problem formulation

Given



$$X: \quad X_0 = 0 \quad X_t = \sum_{i=1}^t V_i + st \quad t \geq 1$$

$$Y: \quad Y_0 = 0 \quad Y_t = X_t + \varepsilon \sum_{i=1}^t W_i \quad t \geq 1$$

with $s \geq 0$, $\varepsilon > 0$, and V_i 's and W_i 's i.i.d. $\mathcal{N}(0, 1)$

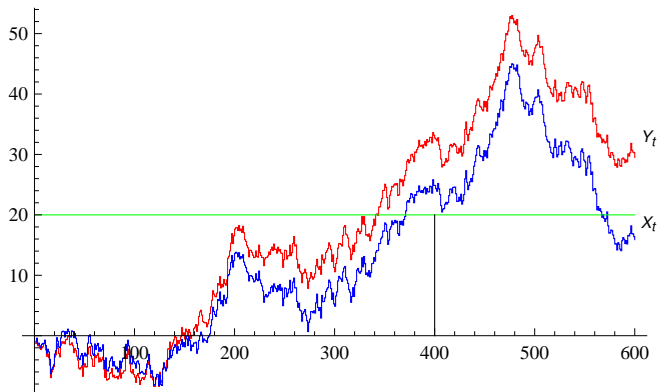


$$\tau = \inf\{t \geq 0 : X_t \geq T\} \quad T > 0$$

Find

$$\inf_{\eta} \mathbb{E}|\eta - \tau|$$

Problem formulation (cont.)



Theorem: asymptotics

For fixed $s > 0$ and $\varepsilon > 0$

$$\inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - \tau| = \inf_{\eta} \mathbb{E}|\eta - \tau| = \sqrt{\frac{2T\varepsilon^2}{\pi s^3(1 + \varepsilon^2)}} [1 + o(1)]$$

as $T \rightarrow \infty$.

⇒ Causality doesn't come at the expense of delay (asym.).

Theorem: asymptotics (general)

Let $1/2 < q < 1$. In the asymptotic regime where $T/s \geq 2$,

$$s \left(\frac{T}{s} \right)^{q-1/2} \rightarrow \infty,$$

and

$$\left(\frac{T}{s} \right)^{1-q} \frac{\varepsilon^2}{1 + \varepsilon^2} \rightarrow \infty,$$

we have

$$\inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - \tau| = \inf_{\eta} \mathbb{E}|\eta - \tau| = \sqrt{\frac{2T\varepsilon^2}{\pi s^3(1 + \varepsilon^2)}} [1 + o(1)].$$

Theorem: upper bound

Fix $\varepsilon > 0$, $s > 0$, $T > 0$, and let

$$\hat{X}_0 = 0 \quad \hat{X}_t = st + \frac{1}{1 + \varepsilon^2}(Y_t - st) \quad t \geq 1.$$

Then, $\eta = \inf\{t \geq 0 : \hat{X}_t \geq T\}$ satisfies

$$\mathbb{E}|\eta - \tau| \leq \sqrt{\frac{2T\varepsilon^2}{\pi(1 + \varepsilon^2)s^3}} + \frac{6}{s} \left(\frac{T}{(2\pi s)^3} \right)^{1/4} + \sqrt{\frac{8(s+2)}{\pi s^3}} + 10 + \frac{20}{s}$$

Theorem: lower bound

Let $\varepsilon > 0$, $s > 0$, and $T/s \geq 2$. Then, for any integer n such that $1 \leq n < T/s$

$$\inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - \tau| \geq \sqrt{\frac{2n\varepsilon^2}{\pi s^2(1+\varepsilon^2)}} \left(1 - Q\left(\frac{T-sn}{\sqrt{n(1+\varepsilon)}}\right) \right) - \sqrt{\frac{2}{\pi s^3}} \left(T - sn + \sqrt{\frac{n}{2\pi}} \right)^{1/2} - 2 - \frac{4}{s}.$$

Theorem: case $s = 0$

If $s = 0$, $\varepsilon > 0$, and $T > 0$, then

$$\mathbb{E}|\eta - \tau|^p = \infty$$

for all $\eta = \eta(Y_0^\infty)$ and $p \geq 1/2$.

Extension to Brownian motion processes

Theorem

$$\begin{aligned}
 X: \quad X_0 &= 0 & X_t &= B_t + st & t > 0 \\
 Y: \quad Y_0 &= 0 & Y_t &= X_t + \varepsilon W_t & t > 0
 \end{aligned}$$

with $s \geq 0$, $\varepsilon > 0$, and $\{B_t\}$ and $\{W_t\}$ independent standard Brownian motions. Then,

$$\inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - \tau| = \inf_{\eta} \mathbb{E}|\eta - \tau| = \sqrt{\frac{2T\varepsilon^2}{\pi s^3(1 + \varepsilon^2)}} [1 + o(1)] .$$

Bounds

Theorem

- for $\varepsilon > 0$, $s > 0$, and $T > 0$

$$\inf_{\eta} \mathbb{E} |\eta - \tau| \leq \sqrt{\frac{2T\varepsilon^2}{\pi(1 + \varepsilon^2)s^3}} + \frac{6}{s} \left(\frac{T}{(2\pi s)^3} \right)^{1/4}$$

- for $\varepsilon > 0$, $s > 0$, $T/s \geq 2$ and any integer n such that $1 \leq n < T/s$

$$\inf_{\eta(Y_0^\infty)} \mathbb{E} |\eta - \tau| \geq \sqrt{\frac{2n\varepsilon^2}{\pi s^2(1 + \varepsilon^2)}} \left(1 - Q \left(\frac{T - sn}{\sqrt{n(1 + \varepsilon)}} \right) \right) - \sqrt{\frac{2}{\pi s^3}} \left(T - sn + \sqrt{\frac{n}{2\pi}} \right)^{1/2}.$$

Upper bound: sketch

$$X: \quad X_0 = 0 \quad X_t = \sum_{i=1}^t V_i + st \quad t \geq 1$$

$$Y: \quad Y_0 = 0 \quad Y_t = X_t + \varepsilon \sum_{i=1}^t W_i \quad t \geq 1$$

- (X_t, Y_t) jointly Gaussian
- hence, linear estimator $\hat{X}_t = aY_t + b$ minimizes $\mathbb{E}|\hat{X}_t - X_t|$
 (actually minimizes $\mathbb{E}|\hat{X}_t - X_t|^p$ for all $p \geq 1$),
- natural tracker: $\eta = \inf\{t \geq 1 : \hat{X}_t \geq T\}$

Upper bound: sketch (cont.)

- to evaluate $\mathbb{E}|\eta - \tau|$, first evaluate $\mathbb{E}|\hat{X}_\eta - X_\tau|$
- to evaluate $\mathbb{E}\hat{X}_\eta$ and $\mathbb{E}X_\tau$ use bounds on overshoot
- connect $\mathbb{E}|\hat{X}_\eta - X_\tau|$ to $\mathbb{E}|\eta - \tau|$ using Wald's equation
 ($\mathbb{E} \sum_{t=1}^{\mu} Z_t = \mathbb{E}\mu\mathbb{E}Z$)

Lower bound: sketch

Idea: reduce the problem to the estimation of X_n for $n \approx T/s$.



$$\inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - \tau| \geq \frac{1}{s} \inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - X_n| + \text{small}$$

- (X_n, Y_n) are jointly Gaussian, thus $X_n \stackrel{d}{=} aY_n + bV + c$ with $V \sim \mathcal{N}(0, 1)$ independent of Y_n , hence

$$\frac{1}{s} \inf_{\eta(Y_0^\infty)} \mathbb{E}|\eta - X_n| \approx \sqrt{\frac{2n\varepsilon^2}{\pi s^2(1 + \varepsilon^2)}}$$

Summary

- the TST problem formulation naturally appears in
 - detection
 - prediction
 - communication
 - quality control
 - information econometrics
- Contribution: correlated gaussian random walks
 - non-asymptotic upper and lower bounds on minimum value of $\mathbb{E}|\eta - \tau|$
 - asymptotic characterization on minimum value of $\mathbb{E}|\eta - \tau|$ (extends to $\mathbb{E}|\eta - \tau|^p, p \geq 1$)
- A lot remains to be done:
 - more general delay penalty functions
 - Gaussianity is key for our results, what if non-gaussian?