



Evaluation of Side-Channel Attacks Using α -Information

CryptArchi Workshop 2022

<u>Yi Liu</u>¹, Wei Cheng^{2,1}, Sylvain Guilley^{2,1}, and Olivier Rioul¹ ¹Télécom Paris, IP Paris; ²Secure-IC May 31st, 2022

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Motivation

 $\alpha\text{-Information Theory}$

Evaluation of Side-Channel Attacks

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Side-Channel Attacks:



Information-theoretic tools have been frequently used in the side-channel analysis.



Motivation

Chérisey et al. (CHES'19) establish some universal inequalities between **the probability of success** of a side-channel attack and **the minimum number of queries** to reach a given success rate.

This result is:

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- valid for any leakage model, any kind of attack.
- with the best possible knowledge on the attacker's side.

They mainly use two information measures:

- Mutual Information: $I(K; Y) = \mathbb{E}_Y \sum_k p(k|y) \log \frac{p(k|y)}{p(k)}$ (k: the secret, y: the leakage.)
- **Bianry Divergence**: $d(p||q) = p \log \frac{p}{q} + (1-p) \log \frac{1-p}{1-q}$ (where p, q are two probability distributions.)

Side-Channel Model



- *K*: a secret key; normally we assume $K \sim U(M)$;
- **T** = $(T_1, T_2, ..., T_q)$: a *q*-element vector, each element T_i is a plain or cypher text;
- $\boldsymbol{X} = (X_1, X_2, \dots, X_q)$: the output of the side-channel; the leakage function $\boldsymbol{X} = f(K, \boldsymbol{T})$ is deterministic;
- $\mathbf{Y} = (Y_1, Y_2, \dots, Y_q)$: traces measured by the attacker;
- \hat{K} : the guess of the secret key;

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■ \mathbb{P}_{s} : the probability of success, defined as $\mathbb{P}_{s} = \mathbb{P}(K = \hat{K})$; it's different from **success rate**.

Main Theorem

One has the following upper bound on the probability of success \mathbb{P}_s :

$$I(\boldsymbol{X}, \boldsymbol{Y} | \boldsymbol{T}) \geq d(\mathbb{P}_{s} \| \frac{1}{M})$$

As shown in the following figure, $d(\mathbb{P}_s \| \frac{1}{M})$ is increasing in \mathbb{P}_s . Hence this inequality gives an upper bound on \mathbb{P}_s when $\mathbb{P}_s \ge 1/M$.



• One can evaluate \mathbb{P}_s by calculating the numerical value of $I(\mathbf{X}, \mathbf{Y} | \mathbf{T})$.



Bounds on Probability of Success



In this figure(Chérisey et al., CHES'19):

- the purple line: the numerical value of I(X, Y|T);
- the blue and green lines: some explicit formula bounds.
 - ML distinguisher is optimal if leakage modle is known.

Properties

To prove the main theorem, the following properties of mutual information turn out to be essential:

- Conditional Consistency (CC): If T is independent of (X, Y) then I(X; Y|T) = I(X; Y)
- Uniform Expansion Property (UEP): If $K \sim U(M)$ and it is independent of T, then $I(K; Y|T) = H(K) - H(K|YT) = \log M - H(K|YT)$
- **Data Processing Inequality (DPI):** If W X Y Z forms a conditional Markov chain given *T*, then I(X; Y|T) > I(W; Z|T)
- Fano's Inequality: $H(K|\hat{K}) \leq H_2(\mathbb{P}_e) + \mathbb{P}_e \log_2(M-1)$

$$\iff I(K;\hat{K}) \geq d(\mathbb{P}_{s} \| \frac{1}{M})$$



Our aims:

- make it more flexible by introducing a Rényi parameter α ;
- extend the work of Chérisey et al. to α-information quantities depending on a parameter α;
- obtain tighter bounds by changing the value of α .

 \implies need an appropriate definition of $I_{\alpha}(X; Y|T)$ with four properties :

CC + UEP + DPI + Fano's inequality

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Let p, q be two probability distributions. We use the following notations:

$$egin{aligned} \| oldsymbol{p} \|_lpha &= ig(\sum oldsymbol{p}^lphaig)^{1/lpha} \ \langle oldsymbol{p} \| oldsymbol{q}
angle_lpha &= ig(\sum oldsymbol{p}^lphaoldsymbol{q}^{1-lpha}ig)^{1/lpha} \end{aligned}$$

• α -entropy (Rényi, 1961): $H_{\alpha}(P) = \frac{\alpha}{1-\alpha} \log \|p\|_{\alpha}$

• α -divergence (Rényi, 1961): $D_{\alpha}(P \| Q) = \frac{1}{\alpha - 1} \log \langle p \| q \rangle_{\alpha}^{\alpha}$



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Their conditional versions:

Conditional *α***-entropy** (Arimoto, 1975):

$$H_{\alpha}(X|Y) = \frac{\alpha}{1-\alpha} \log \mathbb{E}_{Y} \| p_{X|Y} \|_{\alpha}$$

Conditional α-divergence (Verdú, 2015):

$$D_{lpha}(P_{Y|X} \| Q_{Y|X} | P_X) = rac{1}{lpha - 1} \log \mathbb{E}_X \langle p_{Y|X} \| q_{Y|X}
angle_{lpha}^{lpha}$$

Among several definitions of α -information, Sibson's proposal seems to be the most appropriate one (Verdú, 2015).

• α -information (Sibson, 1969):

$$I_{\alpha}(X;Y) = \frac{\alpha}{\alpha - 1} \log \mathbb{E}_{Y} \langle p_{X|Y} \| p_{X} \rangle_{\alpha}$$

Conditional α-information (Liu et al., ITW'21): our proposal of conditional α-information is

$$I_{\alpha}(X;Y|Z) = \frac{\alpha}{\alpha-1} \log \mathbb{E}_{Z} \mathbb{E}_{Y|Z} \langle p_{X|YZ} \| p_{X|Z} \rangle_{\alpha}$$



Conditional α -Information

Among many proposed definitions in the literature, our definition is the only one that satisfies the following properties:

Conditional Consistency (CC): If *T* is independent of (*X*, *Y*) then

 $I_{\alpha}(X;Y|T) = I_{\alpha}(X;Y)$

Uniform Expansion Property (UEP): If K ~ U(M) is uniformly distributed independent of T, then

 $I_{\alpha}(K;Y|T) = H_{\alpha}(K) - H_{\alpha}(K|YT) = \log M - H_{\alpha}(K|YT)$

Data Processing Inequality (DPI): If W – X – Y – Z forms a conditional Markov chain given T, then

 $I_{\alpha}(X;Y|T) \geq I_{\alpha}(W;Z|T)$

Rioul's Generalized Fano Inequality (GFI) (Rioul, GSI'21):

 $I_{\alpha}(X;Y) \geq d_{\alpha}(\mathbb{P}_{s}(X|Y)||\mathbb{P}_{s}(X))$





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Theorem (Upper Bound on \mathbb{P}_{S})

One has the following upper bound on the probability of success \mathbb{P}_s :

$$I_{lpha}(oldsymbol{X},oldsymbol{Y}|oldsymbol{T})\geq d_{lpha}(\mathbb{P}_{s}\|rac{1}{M})$$

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Proof.

Because $K - \mathbf{X} - \mathbf{Y}$ is a Markov chain given \mathbf{T} ,

$$I_{\alpha}(\boldsymbol{X}, \boldsymbol{Y}|\boldsymbol{T}) \geq I_{\alpha}(K, \boldsymbol{Y}|\boldsymbol{T})$$
 (DPI)

Because **X** is a deterministic function of K and **T**, $\mathbf{X} - K - \mathbf{Y}$ is a Markov chains given **T**, which gives us

 $I_{\alpha}(\boldsymbol{X}, \boldsymbol{Y} | \boldsymbol{T}) \leq I_{\alpha}(\boldsymbol{K}, \boldsymbol{Y} | \boldsymbol{T})$ (DPI)

So
$$I_{\alpha}(\mathbf{X}, \mathbf{Y}|\mathbf{T}) = I_{\alpha}(K, \mathbf{Y}|\mathbf{T})$$
. Then one has
 $I_{\alpha}(\mathbf{X}, \mathbf{Y}|\mathbf{T}) = I_{\alpha}(K, \mathbf{Y}|\mathbf{T}) \ge I_{\alpha}(K; \hat{K}|\mathbf{T})$ (DPI)
 $= \log M - H_{\alpha}(K|\hat{K}, \mathbf{T})$ (UEP)
 $\ge \log M - H_{\alpha}(K|\hat{K})$ (*)
 $= I_{\alpha}(K, \hat{K})$ (UEP)
 $\ge d_{\alpha}(\mathbb{P}_{s}(K|\mathbf{Y})||\mathbb{P}_{s}(K))$ (GFI)

 $\ast:$ condition reduce $\alpha\text{-entropy.}$

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Theorem (Upper Bound on \mathbb{P}_{S})

$$I_{lpha}(\boldsymbol{X}, \boldsymbol{Y} | \boldsymbol{T}) \geq d_{lpha}(\mathbb{P}_{s} \| \frac{1}{M})$$





Theorem (Upper Bound on \mathbb{P}_{S})

$$I_{\alpha}(\boldsymbol{X}, \boldsymbol{Y} | \boldsymbol{T}) \geq d_{\alpha}(\mathbb{P}_{s} \| \frac{1}{M})$$

We consider an **implementation of the AES**, with the **Hamming weight leakage model** and **additive white Gaussian noise (AWGN) channel**.

$$\mathbf{Y} = \mathbf{X} + \mathbf{N} = w_H(S(\mathbf{T} \oplus K)) + \mathbf{N} \qquad (i = 1, 2, \dots, q)$$
(1)

- w_H : the Hamming weight
- S: an S-box permutation
- **•** $\mathbf{N} = (N_1, \ldots, N_q)$, N_i are i.i.d $\sim \mathcal{N}(\mathbf{0}, \sigma^2)$

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Numerical Simulations

$$\mathbf{Y} = w_H(S(\mathbf{T} \oplus K)) + \mathbf{N}$$
 $(i = 1, 2, \dots, q)$

Aim: compute

$$I_{\alpha}(\mathbf{X}, \mathbf{Y}|\mathbf{T}) = I_{\alpha}(K, \mathbf{Y}|\mathbf{T}) = \frac{\alpha}{\alpha - 1} \log \mathbb{E}_{\mathbf{T}} \mathbb{E}_{\mathbf{Y}|\mathbf{T}} (\sum_{k} p_{K|\mathbf{T}}^{\alpha} p_{K|\mathbf{T}}^{1-\alpha})^{\frac{1}{\alpha}}$$
(2)

We compute $I_{\alpha}(\mathbf{X}, \mathbf{Y} | \mathbf{T})$ using Monte Carlo simulation.

Monte Carlo methods rely on **repeated random sampling** to obtain numerical results.

By the law of large numbers, **integrals described by the expected value of some random variable** can be approximated by taking the empirical mean (a.k.a. the sample mean) of independent samples of the variable.



Numerical Simulations

 $I_{\alpha}(\mathbf{X},\mathbf{Y}|\mathbf{T})$ can be writen as

$$I_{\alpha}(\mathbf{X}, \mathbf{Y}|\mathbf{T}) = \frac{\alpha}{\alpha - 1} \log \left(\mathbb{E}_{YT} \frac{\left(\sum_{k} p(k|\mathbf{t}) p^{\alpha}(\mathbf{y}|\mathbf{t}, k) \right)^{\frac{1}{\alpha}}}{p(\mathbf{y}|\mathbf{t})} \right).$$
(3)

Therefore, it can be estimated by using Monte-Carlo simulation in the following way:

$$\exp\left(\frac{\alpha-1}{\alpha}I_{\alpha}(\mathbf{X},\mathbf{Y}|\mathbf{T})\right) \approx \lim_{N_{C}\to\infty} \frac{1}{N_{C}} \sum_{j=1}^{N_{C}} \frac{\left(\sum_{k} p(k|\mathbf{t}^{j})p^{\alpha}(\mathbf{y}^{j}|\mathbf{t}^{j},k)\right)^{\frac{1}{\alpha}}}{p(\mathbf{y}^{j}|\mathbf{t}^{j})}$$
$$= \lim_{N_{C}\to\infty} \frac{1}{N_{C}} \sum_{j=1}^{N_{C}} \frac{\left(\sum_{k} p(k)p^{\alpha}(\mathbf{y}^{j}|\mathbf{t}^{j},k)\right)^{\frac{1}{\alpha}}}{\sum_{k} p(k)p(\mathbf{y}^{j}|\mathbf{t}^{j},k)},$$

where $\mathbf{t}^{i} \sim \mathcal{U}(\mathbb{F}_{2^{8}}^{q})$ and $\mathbf{y}^{j} \sim \mathcal{N}(f(\mathbf{t}^{j}, k^{j}), \sigma^{2}\mathbf{I}_{q}) \in \mathbb{R}^{q}$ by choosing $k^{j} \sim \mathcal{U}(\mathbb{F}_{2^{8}})$ and $f(\mathbf{t}^{j}, k^{j}) = w_{H}(S(\mathbf{t}^{j} \oplus k^{j}))$.

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Noise level: $\sigma = 4$.

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Noise level: $\sigma = 8$



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