# Unsupervised learning: Anomaly detection Part II: Functional data 

Pavlo Mozharovskyi<br>LTCI, Telecom Paris, Institut Polytechnique de Paris

## Parcours Data Science BPCE

Paris, the 13th of June 2023

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Functional data framework

- Let $\boldsymbol{F}=\left\{\boldsymbol{F}(t) \in \mathbb{R}^{d}, t \in[0,1]\right\}$ be a random variable that takes its values in a (multivariate) functional space.


## Functional data framework

- Let $\boldsymbol{F}=\left\{\boldsymbol{F}(t) \in \mathbb{R}^{d}, t \in[0,1]\right\}$ be a random variable that takes its values in a (multivariate) functional space.
- In practice, we only have access to the realization of $\boldsymbol{F}$ at a finite number of arguments/times, $\boldsymbol{f}=\left\{\boldsymbol{f}\left(t_{1}\right), \ldots, \boldsymbol{f}\left(t_{p}\right)\right\}$ such that $0 \leq t_{1}<\cdots<t_{p} \leq 1$.


## Functional data framework

- Let $\boldsymbol{F}=\left\{\boldsymbol{F}(t) \in \mathbb{R}^{d}, t \in[0,1]\right\}$ be a random variable that takes its values in a (multivariate) functional space.
- In practice, we only have access to the realization of $\boldsymbol{F}$ at a finite number of arguments/times, $\boldsymbol{f}=\left\{\boldsymbol{f}\left(t_{1}\right), \ldots, \boldsymbol{f}\left(t_{p}\right)\right\}$ such that $0 \leq t_{1}<\cdots<t_{p} \leq 1$.
- The first step: reconstruct functional object from partial observations (time-series) with interpolation or basis decomposition.



## Taxonomy of functional anomalies (Hubert et al., 2015)

A non-complete taxomony of functional abnormalities:

Shape anomalies


Shift anomalies


Isolated anomalies


## Taxonomy of functional anomalies (Airbus data)

A non-complete taxomony of functional abnormalities:
Magnitude (=location, shift) anomalies


Shape anomalies



Isolated anomalies


## Contents

## Anomaly detection in functional framework

Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## FIF in the context of FAD contributions



## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Functional Isolation Forest

- $X_{1}, \ldots, X_{n}$ are random variables in Hilbert space $\mathcal{H}$ and $\mathcal{D} \subset \mathcal{H}$.
- This ensemble learning algorithm builds a collection of binary tree based on a recursive and randomized tree-structured partitioning procedure.


## Functional Isolation Forest

- $X_{1}, \ldots, X_{n}$ are random variables in Hilbert space $\mathcal{H}$ and $\mathcal{D} \subset \mathcal{H}$.
- This ensemble learning algorithm builds a collection of binary tree based on a recursive and randomized tree-structured partitioning procedure.



## Functional Isolation Forest

- $X_{1}, \ldots, X_{n}$ are random variables in Hilbert space $\mathcal{H}$ and $\mathcal{D} \subset \mathcal{H}$.
- This ensemble learning algorithm builds a collection of binary tree based on a recursive and randomized tree-structured partitioning procedure.


$$
\left\{\left\langle X_{i}, \mathbf{d}\right\rangle_{\mathcal{H}}, \quad i \leq n\right\}
$$

Step 2:

$$
\left\{\left\langle X_{i}, \mathbf{d}\right\rangle_{\mathcal{H}}, i \leq n\right\}
$$


$\left\{x:\langle x, \mathbf{d}\rangle_{H} \leq y\right\} \quad\left\{x:\langle x, \mathbf{d}\rangle_{H}>y\right\}$

- The trick: an anomaly should be isolated faster than normal data.


## Functional Isolation Forest <br> Illustration: Isolation tree

Isolation tree, split 25


## Children node construction in a functional isolation tree

 If a node $(j, k)$ is non terminal, it is split in three steps as follows:1. Choose a Split function $\mathbf{d}$ according to the probability distribution $\boldsymbol{\nu}$ on $\mathcal{D}$.

## Children node construction in a functional isolation tree

 If a node $(j, k)$ is non terminal, it is split in three steps as follows:1. Choose a Split function $\mathbf{d}$ according to the probability distribution $\boldsymbol{\nu}$ on $\mathcal{D}$.
2. Choose randomly and uniformly a Split value $\gamma$ in the interval

$$
\left[\min _{\mathbf{x} \in \mathcal{S}_{j, k}}\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}, \max _{\mathbf{x} \in \mathcal{S}_{j, k}}\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}\right],
$$

## Children node construction in a functional isolation tree

 If a node $(j, k)$ is non terminal, it is split in three steps as follows:1. Choose a Split function $\mathbf{d}$ according to the probability distribution $\boldsymbol{\nu}$ on $\mathcal{D}$.
2. Choose randomly and uniformly a Split value $\gamma$ in the interval

$$
\left[\min _{\mathbf{x} \in \mathcal{S}_{j, k}}\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}, \max _{\mathbf{x} \in \mathcal{S}_{j, k}}\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}\right],
$$

3. Form the children subsets

$$
\begin{aligned}
\mathcal{C}_{j+1,2 k} & =\mathcal{C}_{j, k} \cap\left\{\mathbf{x} \in \mathcal{H}:\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}} \leq \gamma\right\}, \\
\mathcal{C}_{j+1,2 k+1} & =\mathcal{C}_{j, k} \cap\left\{\mathbf{x} \in \mathcal{H}:\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}>\gamma\right\} .
\end{aligned}
$$

as well as the children training datasets

$$
\mathcal{S}_{j+1,2 k}=\mathcal{S}_{j, k} \cap \mathcal{C}_{j+1,2 k} \text { and } \mathcal{S}_{j+1,2 k+1}=\mathcal{S}_{j, k} \cap \mathcal{C}_{j+1,2 k+1} .
$$

## Children node construction in a functional isolation tree

 If a node $(j, k)$ is non terminal, it is split in three steps as follows:1. Choose a Split function $\mathbf{d}$ according to the probability distribution $\nu$ on $\mathcal{D}$.
2. Choose randomly and uniformly a Split value $\gamma$ in the interval

$$
\left[\min _{\mathbf{x} \in \mathcal{S}_{j, k}}\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}, \max _{\mathbf{x} \in \mathcal{S}_{j, k}}\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}\right],
$$

3. Form the children subsets

$$
\begin{aligned}
\mathcal{C}_{j+1,2 k} & =\mathcal{C}_{j, k} \cap\left\{\mathbf{x} \in \mathcal{H}:\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}} \leq \gamma\right\}, \\
\mathcal{C}_{j+1,2 k+1} & =\mathcal{C}_{j, k} \cap\left\{\mathbf{x} \in \mathcal{H}:\langle\mathbf{x}, \mathbf{d}\rangle_{\mathcal{H}}>\gamma\right\} .
\end{aligned}
$$

as well as the children training datasets

$$
\mathcal{S}_{j+1,2 k}=\mathcal{S}_{j, k} \cap \mathcal{C}_{j+1,2 k} \text { and } \mathcal{S}_{j+1,2 k+1}=\mathcal{S}_{j, k} \cap \mathcal{C}_{j+1,2 k+1} .
$$

Stop when only one observation is in each node; isolation.

## Anomaly score prediction

- One may then define the piecewise constant function $h_{\tau}: \mathcal{H} \rightarrow \mathbb{N}$ by: $\forall \boldsymbol{x} \in \mathcal{H}$,
$h_{\tau}(\boldsymbol{x})=j$ if and only if $x \in \mathcal{C}_{j, k}$ and $\mathcal{C}_{j, k}$ is associated to a terminal



## Anomaly score prediction

Anomaly score calculation for observation $\boldsymbol{x}$ :

1. For each isolation tree $i \in\{1, \ldots, N\}$, locate $\boldsymbol{x}$ in a terminal node and calculate the depth of this node $h_{i}(\boldsymbol{x})$.
2. Attribute the anomaly score:

$$
s_{n}(\boldsymbol{x})=2^{-\frac{1}{N \cdot c(n)} \sum_{i=1}^{N} h_{i}(\boldsymbol{x})},
$$

with $c(n)=2 H(n-1)-\frac{2(n-1)}{n}$ where $H(k)$ is the harmonic number and can be estimated by $\ln (k)+0.5772156649$.

## Anomaly score prediction

Anomaly score calculation for observation $\boldsymbol{x}$ :

1. For each isolation tree $i \in\{1, \ldots, N\}$, locate $\boldsymbol{x}$ in a terminal node and calculate the depth of this node $h_{i}(\boldsymbol{x})$.
2. Attribute the anomaly score:

$$
s_{n}(\boldsymbol{x})=2^{-\frac{1}{N \cdot c(n)} \sum_{i=1}^{N} h_{i}(\boldsymbol{x})}
$$

with $c(n)=2 H(n-1)-\frac{2(n-1)}{n}$ where $H(k)$ is the harmonic number and can be estimated by $\ln (k)+0.5772156649$.

Score behavior:

- when $\frac{1}{N} \sum_{i=1}^{N} h_{i}(x) \rightarrow c(n), s_{n}(\boldsymbol{x}) \rightarrow 0.5$,
- when $\frac{1}{N} \sum_{i=1}^{N} h_{i}(\boldsymbol{x}) \rightarrow 0, s_{n}(\boldsymbol{x}) \rightarrow 1$,
- when $\frac{1}{N} \sum_{i=1}^{N} h_{i}(\boldsymbol{x}) \rightarrow n-1, s_{n}(\boldsymbol{x}) \rightarrow 0$.


## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Parameters of FIF

- Classical parameters of Isolation Forest :
- number of trees,
- size of the subsample,
- height limit.
- New parameters due to the functional setup :

1. The dictionary $\mathcal{D}$.
2. The probability measure $\boldsymbol{\nu}$.
3. The scalar product $\langle., .\rangle_{\mathcal{H}}$.

## The role of the scalar product

- Compromise between both location and shape :

$$
\langle\mathbf{f}, \mathbf{g}\rangle:=\alpha \times \frac{\langle\mathbf{f}, \mathbf{g}\rangle_{L_{2}}}{\|\mathbf{f}\|\|\mathbf{g}\|}+(1-\alpha) \times \frac{\left\langle\mathbf{f}^{\prime}, \mathbf{g}^{\prime}\right\rangle_{L_{2}}}{\left\|\mathbf{f}^{\prime}\right\|\left\|\mathbf{g}^{\prime}\right\|}, \quad \alpha \in[0,1]
$$

## Example on a toy dataset :

- 90 curves defined by $\mathbf{x}(t)=30(1-t)^{q} t^{q}$ with $q$ equispaced in $[1,1.4]$,
- 10 abnormal curves defined by $\mathbf{x}(t)=30(1-t)^{1.2} t^{1.2}$ noised by $\varepsilon \sim \mathcal{N}\left(0,0.3^{2}\right)$ on the interval [ $\left.0.2,0.8\right]$.

$$
\alpha=1
$$



$$
\alpha=0
$$



## Ability to detect a variety of anomalies

- Sobolev inner product: $\langle.,.\rangle W_{1,2}$.
- Gaussian wavelets dictionary

$$
\mathbf{d}_{\theta, \sigma}(t)=\frac{2}{\sqrt{3 \sigma} \pi^{1 / 4}}\left(1-\left(\frac{t-\theta}{\sigma}\right)^{2}\right) \exp \left(\frac{-(t-\theta)^{2}}{2 \sigma^{2}}\right) .
$$

- Uniform measure $\boldsymbol{\nu}$.





## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking

## Extension of FIF: Connection to data depth

Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Performance on real datasets (1)

- FIF with 4 setups (Dictionary+scalar product):
- Dyadic indicator (DI) $+L_{2}$
- Cosine (Cos) $+L_{2}$
- Cosine (Cos)+Sobolev
- Dataset itself (Self) $+L_{2}$


## Competitors:

- Isolation Forest,Local Outlier Factor, One-class SVM after dimension reduction by FPCA.
- $f H D_{R P}$ : Random projection method with functional Halspace depth.
- fSDO : Functional Stahel-Donoho Outlyingness.


## Performance on real datasets (2)

| Methods: | $\mathrm{DI}_{L_{2}}$ | Cos $_{\text {obb }}$ | Cos $_{L_{2}}$ | Self $_{L_{2}}$ | IF | LOF | OCSVM | fHD ${ }_{R P}$ | fSDO |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Chinatown | 0.93 | 0.82 | 0.74 | 0.77 | 0.69 | 0.68 | 0.70 | 0.76 | 0.98 |
| Coffee | 0.76 | 0.87 | 0.73 | 0.77 | 0.60 | 0.51 | 0.59 | 0.74 | 0.67 |
| ECGFiveDays | 0.78 | 0.75 | 0.81 | 0.56 | 0.81 | 0.89 | 0.90 | 0.60 | 0.81 |
| ECG200 | 0.86 | 0.88 | 0.88 | 0.87 | 0.80 | 0.80 | 0.79 | 0.85 | 0.86 |
| Handoutlines | 0.73 | 0.76 | 0.73 | 0.72 | 0.68 | 0.61 | 0.71 | 0.73 | 0.76 |
| SonyRobotAI1 | 0.89 | 0.80 | 0.85 | 0.83 | 0.79 | 0.69 | 0.74 | 0.83 | 0.94 |
| SonyRobotAI2 | 0.77 | 0.75 | 0.79 | 0.92 | 0.86 | 0.78 | 0.80 | 0.86 | 0.81 |
| StarLightCurves | 0.82 | 0.81 | 0.76 | 0.86 | 0.76 | 0.72 | 0.77 | 0.77 | 0.85 |
| TwoLeadECG | 0.71 | 0.61 | 0.61 | 0.56 | 0.71 | 0.63 | 0.71 | 0.65 | 0.69 |
| Yoga | 0.62 | 0.54 | 0.60 | 0.58 | 0.57 | 0.52 | 0.59 | 0.55 | 0.55 |
| EOGHorizontal | 0.72 | 0.76 | 0.81 | 0.74 | 0.70 | 0.69 | 0.74 | 0.73 | 0.75 |
| CinECGTorso | 0.70 | 0.92 | 0.86 | 0.43 | 0.51 | 0.46 | 0.41 | 0.64 | 0.80 |
| ECG5000 | 0.93 | 0.98 | 0.98 | 0.95 | 0.96 | 0.93 | 0.95 | 0.91 | 0.93 |

Table: AUC of different anomaly detection methods calculated on the test set. Bold numbers correspond to the best result.

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Extension to multivariate functional data

FIF can be easily extended to the multivariate functional data, i.e. when the quantity of interest lies in $\mathbb{R}^{d}$ for each moment of time:

$$
\begin{aligned}
x: & {[0,1] }
\end{aligned} \longrightarrow \mathbb{R}^{d}, ~\left(\left(x^{1}(t), \ldots, x^{d}(t)\right)\right.
$$

- Coordinate-wise sum of the d corresponding scalar products:

$$
\langle\mathbf{f}, \mathbf{g}\rangle_{L_{2}^{\otimes d}}:=\sum_{i=1}^{d}\left\langle f^{(i)}, g^{(i)}\right\rangle_{L_{2}}
$$

- Dictionaries: Composed by univariate function on each axis, multivariate wavelets, multivariate Brownian motion ...


## Example with MNIST dataset

We extract the digits' contours and obtain bivariate functional curves from MNIST dataset. Each digit is transformed into a curve in $\left(L_{2}([0,1]) \times L_{2}([0,1])\right)$ using length parametrization on $[0,1]$.




2
2


## Connection to data depth and supervised classification

- One may define a functional depth by

$$
D_{\text {FIF }}(x ; \mathcal{S})=1-s_{n}(x ; \mathcal{S})
$$

Assume that we have a training classification dataset of $q$ classes $\mathcal{S}=\mathcal{S}^{1} \cup \ldots \cup \mathcal{S}^{q}$.

- Low dimensional representation based on depth-based map can be defined by

$$
\mathbf{x} \mapsto \phi(\mathbf{x})=\left(D_{\text {FIF }}\left(\mathbf{x} ; \mathcal{S}^{1}\right), \ldots, D_{F I F}\left(\mathbf{x} ; \mathcal{S}^{q}\right)\right) \in[0,1]^{q} .
$$

- One may define a DD-plot classifier by using a classifier on the low dimension representation of the functional dataset.


## Example of depth map on MNIST dataset

$\mathcal{S}$ is constructed by taking 100 digits from class 1,100 from class 5 and 100 from class 7 .


Figure: Depth space embedding of the three digits ( 1,5 and 7 ) of the MNIST dataset.

## Some remarks on FIF

- New anomaly detection algorithm for functional data:
- Great flexibility but dictionaries (and scalar product) are tricky to choose in an unsupervised setting.
- Low complexity and memory requierement.
- Lack of theoretical garanties!

Staerman, G., Mozharovskyi, P., Clémençon, S., and
D'Alché-Buc, F. Functional Isolation Forest. ACML 2019.

All codes are available at:
https://github.com/guillaumestaermanML/FIF.

## Contents

```
Anomaly detection in functional framework
Functional isolation forest
    The method
    FIF parameters
    Real data benchmarking
    Extension of FIF: Connection to data depth
```

Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Detection of (multivariate) functional anomalies



## Detection of (multivariate) functional anomalies



- Functional depth of $\boldsymbol{f}$ w.r.t. $\mathcal{F}=\left\{\boldsymbol{f}_{i}\right\}_{i=1}^{n}$ :

$$
D(\boldsymbol{f} \mid \mathcal{F})=\int_{t_{\min }}^{t_{\max }} D^{1}(\boldsymbol{f}(t) \mid \mathcal{F}(t)) d t
$$

## Detection of (multivariate) functional anomalies



- Functional depth of $\boldsymbol{f}$ w.r.t. $\mathcal{F}=\left\{\boldsymbol{f}_{i}\right\}_{i=1}^{n}$ :

$$
D(\boldsymbol{f} \mid \mathcal{F})=\int_{t_{\min }}^{t_{\max }} D^{1}\left(\boldsymbol{f}(t) \mid\left\{\boldsymbol{f}_{1}(t), \ldots, \boldsymbol{f}_{n}(t)\right\}\right) \boldsymbol{d} t
$$

where $D^{d}(\cdot \mid \cdot)$ is a multivariate data depth, as defined above.

## Detection of (multivariate) functional anomalies



- Functional depth of $\boldsymbol{f}$ w.r.t. $\mathcal{F}=\left\{\boldsymbol{f}_{i}\right\}_{i=1}^{n}$ :

$$
D(\boldsymbol{f} \mid \mathcal{F})=\int_{t_{\min }}^{t_{\max }} D^{1}\left(\boldsymbol{f}(t) \mid\left\{\boldsymbol{f}_{1}(t), \ldots, \boldsymbol{f}_{n}(t)\right\}\right) \boldsymbol{d} t
$$

where $D^{d}(\cdot \mid \cdot)$ is a multivariate data depth, as defined above.

## Detection of (multivariate) functional anomalies



- Functional depth of $\boldsymbol{f}$ w.r.t. $\mathcal{F}=\left\{\boldsymbol{f}_{i}\right\}_{i=1}^{n}$ :

$$
D(\boldsymbol{f} \mid \mathcal{F})=\int_{t_{\min }}^{t_{\max }} D^{1}\left(\boldsymbol{f}(t) \mid\left\{\boldsymbol{f}_{1}(t), \ldots, \boldsymbol{f}_{n}(t)\right\}\right) \boldsymbol{d} t
$$

where $D^{d}(\cdot \mid \cdot)$ is a multivariate data depth, as defined above.

## Detection of (multivariate) functional anomalies



- Functional depth of $\boldsymbol{f}$ w.r.t. $\mathcal{F}=\left\{\boldsymbol{f}_{i}\right\}_{i=1}^{n}$ :

$$
D(\boldsymbol{f} \mid \mathcal{F})=\int_{t_{\min }}^{t_{\max }} D^{1}\left(\boldsymbol{f}(t) \mid\left\{\boldsymbol{f}_{1}(t), \ldots, \boldsymbol{f}_{n}(t)\right\}\right) d t
$$

where $D^{d}(\cdot \mid \cdot)$ is a multivariate data depth, as defined above.

- Label $\boldsymbol{f}$ as anomaly if $D(\boldsymbol{f} \mid \mathcal{F})<\min (D)$.


## Integrated depth for functional data



Let $\boldsymbol{F}$ be a stochastic process with continuous paths defined on $[0,1]$, and $\boldsymbol{f}$ its realization.

## Integrated depth for functional data



Let $\boldsymbol{F}$ be a stochastic process with continuous paths defined on $[0,1]$, and $\boldsymbol{f}$ its realization. Then:

$$
D(\boldsymbol{f} \mid \boldsymbol{F})=\int_{0}^{1} D(\boldsymbol{f}(t) \mid \boldsymbol{F}(t)) d t
$$

see Fraiman, Muniz, 2001; also López-Pintado, Romo, 2011.

## Integrated depth for functional data



Let $\boldsymbol{F}$ be a stochastic process with continuous paths defined on $[0,1]$, and $\boldsymbol{f}$ its realization. Then:

$$
D(\boldsymbol{f} \mid \boldsymbol{F})=\int_{0}^{1} \min \left\{F_{\boldsymbol{F}(t)}(\boldsymbol{f}(t)), 1-F_{\boldsymbol{F}(t)}\left(\boldsymbol{f}(t)^{-}\right)\right\} d t
$$

see Fraiman, Muniz, 2001; also López-Pintado, Romo, 2011.

## Multivariate functional halfspace depth

Let $\boldsymbol{F}$ be a $d$-real-valued stochastic process with continuous paths defined on $[0,1]$, and $\boldsymbol{f}$ its realization. Then:

$$
\begin{gathered}
M F D(\boldsymbol{f} \mid \boldsymbol{F})=\int_{0}^{1} D(\boldsymbol{f}(t) \mid \boldsymbol{F}(t)) \cdot w(t) d t \\
w(t)=w_{\alpha}(t, \boldsymbol{F}(t))=\frac{\operatorname{vol}\left\{D_{\alpha}(\boldsymbol{F}(t))\right\}}{\int_{0}^{1} \operatorname{vol}\left\{D_{\alpha}(\boldsymbol{F}(u))\right\} d u}
\end{gathered}
$$

see Claeskens, Hubert, Slaets, Vakili, 2014.

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

Functional depth: Motivation 1


Functional depth: Motivation 1


Functional depth: Motivation 1


## Functional depth: Motivation 1



Regard the following different parametrizations of a curve:
Parametrization A:

$$
\begin{aligned}
& x_{1}(t)=-(\cos (t)+1) \mathbb{1}\left\{t<\frac{3 \pi}{2}\right\}-(\cos (3 t-3 \pi)+1) \mathbb{1}\left\{t \geq \frac{3 \pi}{2}\right\}+1 \\
& x_{2}(t)=(\sin (t)+1) \mathbb{1}\left\{t<\frac{3 \pi}{2}\right\}-(\sin (3 t-3 \pi)+1) \mathbb{1}\left\{t \geq \frac{3 \pi}{2}\right\}
\end{aligned}
$$

Parametrization B:

$$
\begin{aligned}
& x_{1}(t)=-(\cos (3 t)+1) \mathbb{1}\left\{t<\frac{\pi}{2}\right\}-(\cos (t+\pi)+1) \mathbb{1}\left\{t \geq \frac{\pi}{2}\right\}+1 \\
& x_{2}(t)=(\sin (3 t)+1) \mathbb{1}\left\{t<\frac{\pi}{2}\right\}-(\sin (t+\pi)+1) \mathbb{1}\left\{t \geq \frac{\pi}{2}\right\}
\end{aligned}
$$

Functional depth: Motivation 1


Regard the following different parametrizations of a curve:



Functional depth: Motivation 1

Parametrization A


## Parametrization $B$



Parametrization:



Functional depth: Motivation 1

Parametrization A


## Parametrization B



Parametrization:



Functional depth: Motivation 1

Parametrization A


## Parametrization $B$



## Functional depth: Motivation 1

Parametrization A

## Parametrization B


${ }^{\times 2}$ The depth-induced orders differ! ${ }^{\times 2}$


## Functional depth: Motivation 2

## Functional halfspace depth for the FDA-data



Parametrization by length


Depth-induced ranking for parametrizations by time and by length:

| Time | 2 | 3 | 13 | 12 | 4 | 8 | 1 | 17 | 11 | 9 | 7 | 19 | 15 | 20 | 18 | 16 | 14 | 5 | 6 | 10 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Length | 6 | 3 | 16 | 14 | 5 | 7 | 13 | 11 | 1 | 17 | 2 | 19 | 8 | 20 | 12 | 18 | 15 | 4 | 9 | 10 |

## Functional depth: Motivation 3

Simulated hurricane tracks: curve boxplot

MFH depth - par. time

mSB depth - par. time


MFH depth - par. length

mSB depth - par. length


## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## The space of curves

- Let $\left(\mathbb{R}^{d},|\cdot|_{2}\right)$ be the Euclidean space.


## The space of curves

- Let $\left(\mathbb{R}^{d},|\cdot|_{2}\right)$ be the Euclidean space.
- A parametrized curve $\beta:[0,1] \rightarrow \mathbb{R}^{d}$ is a continuous map. A reparametrization $\gamma:[0,1] \rightarrow[0,1]$ is increasing continuous function: $\gamma(0)=0$ and $\gamma(1)=1$.


## The space of curves

- Let $\left(\mathbb{R}^{d},|\cdot|_{2}\right)$ be the Euclidean space.
- A parametrized curve $\beta:[0,1] \rightarrow \mathbb{R}^{d}$ is a continuous map. A reparametrization $\gamma:[0,1] \rightarrow[0,1]$ is increasing continuous function: $\gamma(0)=0$ and $\gamma(1)=1$.
- Two parametrized curves $\beta_{1}, \beta_{2}$ are equivalent if and only if there exist two reparametrizations $\gamma_{1}, \gamma_{2}: \beta_{1} \circ \gamma_{1}=\beta_{2} \circ \gamma_{2}$.


## The space of curves

- Let $\left(\mathbb{R}^{d},|\cdot|_{2}\right)$ be the Euclidean space.
- A parametrized curve $\beta:[0,1] \rightarrow \mathbb{R}^{d}$ is a continuous map. A reparametrization $\gamma:[0,1] \rightarrow[0,1]$ is increasing continuous function: $\gamma(0)=0$ and $\gamma(1)=1$.
- Two parametrized curves $\beta_{1}, \beta_{2}$ are equivalent if and only if there exist two reparametrizations $\gamma_{1}, \gamma_{2}: \beta_{1} \circ \gamma_{1}=\beta_{2} \circ \gamma_{2}$.
- An unparametrized curve, noted $\mathcal{C}:=\mathcal{C}_{\beta}$, is defined as the equivalence class of $\beta$ up to the above equivalence relation. The space of unparametrized curves is then defined as

$$
\mathfrak{B}=\left\{\mathcal{C}_{\beta}: \beta \in \mathcal{C}\left([0,1], \mathbb{R}^{d}\right)\right\} .
$$

## The space of curves

- Let $\left(\mathbb{R}^{d},|\cdot|_{2}\right)$ be the Euclidean space.
- A parametrized curve $\beta:[0,1] \rightarrow \mathbb{R}^{d}$ is a continuous map. A reparametrization $\gamma:[0,1] \rightarrow[0,1]$ is increasing continuous function: $\gamma(0)=0$ and $\gamma(1)=1$.
- Two parametrized curves $\beta_{1}, \beta_{2}$ are equivalent if and only if there exist two reparametrizations $\gamma_{1}, \gamma_{2}: \beta_{1} \circ \gamma_{1}=\beta_{2} \circ \gamma_{2}$.
- An unparametrized curve, noted $\mathcal{C}:=\mathcal{C}_{\beta}$, is defined as the equivalence class of $\beta$ up to the above equivalence relation. The space of unparametrized curves is then defined as

$$
\mathfrak{B}=\left\{\mathcal{C}_{\beta}: \beta \in \mathcal{C}\left([0,1], \mathbb{R}^{d}\right)\right\}
$$

- We endow $\mathfrak{B}$ with the Fréchet metric:

$$
d_{\mathfrak{B}}\left(\mathcal{C}_{1}, \mathcal{C}_{2}\right)=\inf \left\{\left\|\beta_{1}-\beta_{2}\right\|_{\infty}, \beta_{1} \in \mathcal{C}_{1}, \beta_{2} \in \mathcal{C}_{2}\right\}, \quad \mathcal{C}_{1}, \mathcal{C}_{2} \in \mathfrak{B}
$$

Associated distribution and the sampling scheme

- n


## Associated distribution and the sampling scheme

$\rightarrow$ n

- Let $\mathcal{C}$ be an unparameterized curve. The length of $\mathcal{C}$ :

$$
L(\mathcal{C})=\sup _{\tau}\left\{\sum_{i=1}^{N}\left|\beta\left(\tau_{i}\right)-\beta\left(\tau_{i-1}\right)\right|_{2}: \tau \text { is a partition of }[0,1]\right\}
$$

for all $\beta \in \mathcal{C}$.

## Associated distribution and the sampling scheme

$\rightarrow \mathrm{n}$

- Let $\mathcal{C}$ be an unparameterized curve. The length of $\mathcal{C}$ :
$L(\mathcal{C})=\sup _{\tau}\left\{\sum_{i=1}^{N}\left|\beta\left(\tau_{i}\right)-\beta\left(\tau_{i-1}\right)\right|_{2}: \tau\right.$ is a partition of $\left.[0,1]\right\}$,
for all $\beta \in \mathcal{C}$.
- An unparametrized curve $\mathcal{C}$ is called rectifiable if $L(\mathcal{C})$ is finite. The length $L: \mathfrak{B} \rightarrow \mathbb{R}+\cup\{\infty\}$ is measurable:
$\mathcal{P}=\left\{P\right.$ prob. measure on $\left.\left(\mathfrak{B}, d_{\mathfrak{B}}\right): P(\{\mathcal{C} \in \mathfrak{B} ; 0<L(\mathcal{C})<\infty\})=1\right\}$.


## Associated distribution and the sampling scheme

$\rightarrow \mathrm{n}$

- Let $\mathcal{C}$ be an unparameterized curve. The length of $\mathcal{C}$ :
$L(\mathcal{C})=\sup _{\tau}\left\{\sum_{i=1}^{N}\left|\beta\left(\tau_{i}\right)-\beta\left(\tau_{i-1}\right)\right|_{2}: \tau\right.$ is a partition of $\left.[0,1]\right\}$,
for all $\beta \in \mathcal{C}$.
- An unparametrized curve $\mathcal{C}$ is called rectifiable if $L(\mathcal{C})$ is finite. The length $L: \mathfrak{B} \rightarrow \mathbb{R}+\cup\{\infty\}$ is measurable:
$\mathcal{P}=\left\{P\right.$ prob. measure on $\left.\left(\mathfrak{B}, d_{\mathfrak{B}}\right): P(\{\mathcal{C} \in \mathfrak{B} ; 0<L(\mathcal{C})<\infty\})=1\right\}$.
- Let $\mathcal{X}$ be a random element of $\mathfrak{B}$ stemming from distribution $P \in \mathcal{P}$.


## Associated distribution and the sampling scheme

$\rightarrow \mathrm{n}$

- Let $\mathcal{C}$ be an unparameterized curve. The length of $\mathcal{C}$ :
$L(\mathcal{C})=\sup _{\tau}\left\{\sum_{i=1}^{N}\left|\beta\left(\tau_{i}\right)-\beta\left(\tau_{i-1}\right)\right|_{2}: \tau\right.$ is a partition of $\left.[0,1]\right\}$,
for all $\beta \in \mathcal{C}$.
- An unparametrized curve $\mathcal{C}$ is called rectifiable if $L(\mathcal{C})$ is finite. The length $L: \mathfrak{B} \rightarrow \mathbb{R}+\cup\{\infty\}$ is measurable:
$\mathcal{P}=\left\{P\right.$ prob. measure on $\left.\left(\mathfrak{B}, d_{\mathfrak{B}}\right): P(\{\mathcal{C} \in \mathfrak{B} ; 0<L(\mathcal{C})<\infty\})=1\right\}$.
- Let $\mathcal{X}$ be a random element of $\mathfrak{B}$ stemming from distribution $P \in \mathcal{P}$.
- We derive the probability distribution $Q_{P}$ on $\mathbb{R}^{d}$ as follows: if $X \sim Q_{P}$, then distribution of $X \mid \mathcal{X}=\mathcal{C}$ is the (uniform on $\mathcal{C})$ probability distribution $\mu_{\mathcal{C}}$ :

$$
\mu_{\mathcal{C}}(A)=\int_{\mathcal{C}} \mathbb{1}_{A}(x) d x
$$

## Associated distribution and the sampling scheme

The statistical model:

$$
\mathcal{X}_{1}, \ldots, \mathcal{X}_{n} \text { i.i.d. from } P .
$$

For Monte-Carlo estimation, we can consider the following sampling scheme:

$$
\left\{\begin{array}{l}
\mathcal{X}_{1}, \ldots, \mathcal{X}_{n} \text { i.i.d. from } P \\
\text { for all } i=1, \ldots, n \\
\quad X_{i, 1}, \ldots, X_{i, m} \text { i.i.d. from } \mu_{\mathcal{X}_{i}} .
\end{array}\right.
$$

## Data depth for an unparametrized curve

Definition
The Tukey curve depth of $\mathcal{C} \in \mathfrak{B}$ w.r.t. $Q_{P}$ is defined as:

$$
D\left(\mathcal{C} \mid Q_{P}\right)=\int_{\mathcal{C}} D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right) d \mu_{\mathcal{C}}(\boldsymbol{x})
$$

## Data depth for an unparametrized curve

## Definition

The Tukey curve depth of $\mathcal{C} \in \mathfrak{B}$ w.r.t. $Q_{P}$ is defined as:

$$
D\left(\mathcal{C} \mid Q_{P}\right)=\int_{\mathcal{C}} D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right) d \mu_{\mathcal{C}}(\boldsymbol{x})
$$

where the depth $D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right)$ of an arbitrary point $\boldsymbol{x} \in \mathcal{C}$ w.r.t. the distribution $Q_{P}$ is defined as:

$$
D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right)=\inf \left\{\frac{Q_{P}(H)}{\mu_{\mathcal{C}}(H)}: H \text { closed half-space } \subset \mathbb{R}^{d}, \boldsymbol{x} \in \partial H\right\}
$$

where convention $\frac{0}{0}=0$ is adopted.

## Data depth for an unparametrized curve

## Definition

The Tukey curve depth of $\mathcal{C} \in \mathfrak{B}$ w.r.t. $Q_{P}$ is defined as:

$$
D\left(\mathcal{C} \mid Q_{P}\right)=\int_{\mathcal{C}} D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right) d \mu_{\mathcal{C}}(\boldsymbol{x})
$$

where the depth $D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right)$ of an arbitrary point $\boldsymbol{x} \in \mathcal{C}$ w.r.t. the distribution $Q_{P}$ is defined as:

$$
D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right)=\inf \left\{\frac{Q_{P}(H)}{\mu_{\mathcal{C}}(H)}: H \text { closed half-space } \subset \mathbb{R}^{d}, \boldsymbol{x} \in \partial H\right\}
$$

where convention $\frac{0}{0}=0$ is adopted.
Definition
The sample Tukey curve depth of $\mathcal{C} \in \mathfrak{B}$ w.r.t. $\mathcal{X}_{1}, \ldots, \mathcal{X}_{n}$ is:

$$
D\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)=\int_{\mathcal{C}} D\left(\boldsymbol{x} \mid Q_{n}, \mu_{\mathcal{C}}\right) d \mu_{\mathcal{C}}(\boldsymbol{x})
$$

## Data depth for an unparametrized curve

## Definition

The Tukey curve depth of $\mathcal{C} \in \mathfrak{B}$ w.r.t. $Q_{P}$ is defined as:

$$
D\left(\mathcal{C} \mid Q_{P}\right)=\int_{\mathcal{C}} D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right) d \mu_{\mathcal{C}}(\boldsymbol{x})
$$

where the depth $D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right)$ of an arbitrary point $\boldsymbol{x} \in \mathcal{C}$ w.r.t. the distribution $Q_{P}$ is defined as:

$$
D\left(\boldsymbol{x} \mid Q_{P}, \mu_{\mathcal{C}}\right)=\inf \left\{\frac{Q_{P}(H)}{\mu_{\mathcal{C}}(H)}: H \text { closed half-space } \subset \mathbb{R}^{d}, \boldsymbol{x} \in \partial H\right\}
$$

where convention $\frac{0}{0}=0$ is adopted.

## Definition

The sample Tukey curve depth of $\mathcal{C} \in \mathfrak{B}$ w.r.t. $\mathcal{X}_{1}, \ldots, \mathcal{X}_{n}$ is:

$$
D\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)=\int_{\mathcal{C}} D\left(\boldsymbol{x} \mid Q_{n}, \mu_{\mathcal{C}}\right) d \mu_{\mathcal{C}}(\boldsymbol{x})
$$

where $Q_{n}=\left(\mu_{\mathcal{X}_{1}}+\cdots+\mu_{\mathcal{X}_{n}}\right) / n$.

Data depth for an unparametrized curve: intuition


Data depth for an unparametrized curve: intuition


Traditional reasoning:
$\widehat{Q}_{P}\left(H_{u_{1}}^{x_{1}}\right)=\frac{25}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{u_{1}}^{x_{1}}\right)=\frac{4}{8}$
$\widehat{Q}_{P}\left(H_{-u_{1}}^{\times_{1}}\right)=\frac{15}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{-u_{1}}^{\times_{1}}\right)=\frac{4}{8}$


Curve-based reasoning: $\widehat{Q}_{P}\left(H_{u_{2}}^{x_{2}}\right)=\frac{25}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{u_{2}}^{x_{2}}\right)=\frac{6}{8}$
$\widehat{Q}_{P}\left(H_{-u_{2}}^{x_{2}}\right)=\frac{15}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{-u_{2}}^{x_{2}}\right)=\frac{2}{8}$

Data depth for an unparametrized curve: intuition


Traditional reasoning:
$\widehat{Q}_{P}\left(H_{u_{1}}^{x_{1}}\right)=\frac{25}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{u_{1}}^{\chi_{1}}\right)=\frac{4}{8}$
$\hat{Q}_{P}\left(H_{-u_{1}}^{x_{1}}\right)=\frac{15}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{-u_{1}}^{x_{1}}\right)=\frac{4}{8}$


Curve-based reasoning: $\widehat{Q}_{P}\left(H_{u_{2}}^{x_{2}}\right)=\frac{25}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{u_{2}}^{x_{2}}\right)=\frac{6}{8}$ $\widehat{Q}_{P}\left(H_{-u_{2}}^{x_{2}}\right)=\frac{15}{40}, \widehat{\mu}_{\mathcal{C}}\left(H_{-u_{2}}^{x_{2}}\right)=\frac{2}{8}$

## Data depth for an unparametrized curve: empirical version

- Let a chosen curve consist of two (independently drawn on $\mathcal{C}$ ) parts $\mathbb{Y}_{1, m}=\left(Y_{1,1}, \ldots, Y_{1, m}\right)$ and $\mathbb{Y}_{2, m}=\left(Y_{2,1}, \ldots, Y_{2, m}\right)$ with empirical distribution

$$
\widehat{\mu}_{m}=\frac{1}{m} \sum_{i=1}^{m} \delta_{Y_{1, i}}
$$

where $\delta_{\boldsymbol{x}}$ is the Dirac measure in $\boldsymbol{x} \in \mathbb{R}^{d}$.

## Data depth for an unparametrized curve: empirical version

- Let a chosen curve consist of two (independently drawn on $\mathcal{C}$ ) parts $\mathbb{Y}_{1, m}=\left(Y_{1,1}, \ldots, Y_{1, m}\right)$ and $\mathbb{Y}_{2, m}=\left(Y_{2,1}, \ldots, Y_{2, m}\right)$ with empirical distribution

$$
\widehat{\mu}_{m}=\frac{1}{m} \sum_{i=1}^{m} \delta_{Y_{1, i}}
$$

where $\delta_{\boldsymbol{x}}$ is the Dirac measure in $\boldsymbol{x} \in \mathbb{R}^{d}$.

- Let $\widehat{Q}_{n, m}$ be the empirical distribution (observed sample) $\mathbb{X}_{n, m}=\left\{X_{i, j}, i=1, \ldots, n, j=1, \ldots, m\right\}$

$$
\widehat{Q}_{n, m}=\frac{1}{n m} \sum_{i=1}^{n} \sum_{j=1}^{m} \delta X_{i, j}
$$

## Data depth for an unparametrized curve: empirical version

- Let a chosen curve consist of two (independently drawn on $\mathcal{C}$ ) parts $\mathbb{Y}_{1, m}=\left(Y_{1,1}, \ldots, Y_{1, m}\right)$ and $\mathbb{Y}_{2, m}=\left(Y_{2,1}, \ldots, Y_{2, m}\right)$ with empirical distribution

$$
\widehat{\mu}_{m}=\frac{1}{m} \sum_{i=1}^{m} \delta_{Y_{1, i}}
$$

where $\delta_{\boldsymbol{x}}$ is the Dirac measure in $\boldsymbol{x} \in \mathbb{R}^{d}$.

- Let $\widehat{Q}_{n, m}$ be the empirical distribution (observed sample) $\mathbb{X}_{n, m}=\left\{X_{i, j}, i=1, \ldots, n, j=1, \ldots, m\right\}$

$$
\widehat{Q}_{n, m}=\frac{1}{n m} \sum_{i=1}^{n} \sum_{j=1}^{m} \delta X_{i, j}
$$

- To compute the sample Tukey curve depth, a Monte Carlo approximation is used.


## Data depth for an unparametrized curve: empirical version

- Let $H$ be a closed halfspace in $\mathbb{R}^{d}$ and $\mathcal{H}_{\Delta}^{n, m}$ denote a collection of such halfspaces such that for all $H \in \mathcal{H}_{\Delta}^{n, m}$ either $\widehat{Q}_{n, m}(H)=0$ or $\widehat{\mu}_{m}(H)>\Delta$, almost surely, for $\Delta \in\left(0, \frac{1}{2}\right)$.


## Data depth for an unparametrized curve: empirical version

- Let $H$ be a closed halfspace in $\mathbb{R}^{d}$ and $\mathcal{H}_{\Delta}^{n, m}$ denote a collection of such halfspaces such that for all $H \in \mathcal{H}_{\Delta}^{n, m}$ either $\widehat{Q}_{n, m}(H)=0$ or $\widehat{\mu}_{m}(H)>\Delta$, almost surely, for $\Delta \in\left(0, \frac{1}{2}\right)$.

Definition
The Monte Carlo approximation of the Tukey curve depth of $\mathcal{C}$ w.r.t. $\mathcal{X}_{1}, \ldots, \mathcal{X}_{n}$ is defined as:

$$
\widehat{D}_{n, m, \Delta}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)=\frac{1}{m} \sum_{i=1}^{m} \widehat{D}\left(Y_{2, i} \mid \widehat{Q}_{n, m}, \widehat{\mu}_{m}, \mathcal{H}_{\Delta}^{n, m}\right)
$$

## Data depth for an unparametrized curve: empirical version

- Let $H$ be a closed halfspace in $\mathbb{R}^{d}$ and $\mathcal{H}_{\Delta}^{n, m}$ denote a collection of such halfspaces such that for all $H \in \mathcal{H}_{\Delta}^{n, m}$ either $\widehat{Q}_{n, m}(H)=0$ or $\widehat{\mu}_{m}(H)>\Delta$, almost surely, for $\Delta \in\left(0, \frac{1}{2}\right)$.


## Definition

The Monte Carlo approximation of the Tukey curve depth of $\mathcal{C}$ w.r.t. $\mathcal{X}_{1}, \ldots, \mathcal{X}_{n}$ is defined as:

$$
\widehat{D}_{n, m, \Delta}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)=\frac{1}{m} \sum_{i=1}^{m} \widehat{D}\left(Y_{2, i} \mid \widehat{Q}_{n, m}, \widehat{\mu}_{m}, \mathcal{H}_{\Delta}^{n, m}\right)
$$

with the depth of an arbitrary point $\boldsymbol{x} \in \mathbb{R}^{d}$ w.r.t. $\widehat{Q}_{n, m}$ being:

$$
\widehat{D}\left(x \mid \widehat{Q}_{n, m}, \widehat{\mu}_{m}, \mathcal{H}_{\Delta}^{n, m}\right)=\inf \left\{\frac{\widehat{Q}_{n, m}(H)}{\widehat{\mu}_{m}(H)}: H \in \mathcal{H}_{\Delta}^{n, m}, \boldsymbol{x} \in \partial H\right\}
$$

and $\frac{0}{0}=0$ as before.

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Calculation of the Tukey curve depth



## Data depth for an unparametrized curve: consistency

Theorem
Let $\mathcal{C} \in \mathfrak{B}$ be a rectifiable curve, and let $P$ be a probability measure in the space of curves such that $P \in \mathcal{P}$. Let $\left(\Delta_{m}\right)$ be a decreasing sequence of positive numbers such that $\left(\Delta_{m}\right)$ and $\left(\sqrt{\frac{\log (m)}{m}} / \Delta_{m}^{2}\right)$ converges to zero when $m \rightarrow \infty$.

## Data depth for an unparametrized curve: consistency

Theorem
Let $\mathcal{C} \in \mathfrak{B}$ be a rectifiable curve, and let $P$ be a probability measure in the space of curves such that $P \in \mathcal{P}$. Let $\left(\Delta_{m}\right)$ be a decreasing sequence of positive numbers such that $\left(\Delta_{m}\right)$ and $\left(\sqrt{\frac{\log (m)}{m}} / \Delta_{m}^{2}\right)$ converges to zero when $m \rightarrow \infty$.

Then:

- the Monte Carlo approximation $\widehat{D}_{n, m, \Delta_{m}}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ converges in probability to $D\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ as $m \rightarrow \infty$;


## Data depth for an unparametrized curve: consistency

## Theorem

Let $\mathcal{C} \in \mathfrak{B}$ be a rectifiable curve, and let $P$ be a probability measure in the space of curves such that $P \in \mathcal{P}$. Let $\left(\Delta_{m}\right)$ be a decreasing sequence of positive numbers such that $\left(\Delta_{m}\right)$ and $\left(\sqrt{\frac{\log (m)}{m}} / \Delta_{m}^{2}\right)$ converges to zero when $m \rightarrow \infty$.

Then:

- the Monte Carlo approximation $\widehat{D}_{n, m, \Delta_{m}}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ converges in probability to $D\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ as $m \rightarrow \infty$;
- the Monte Carlo approximation $\widehat{D}_{n, m, \Delta_{m}}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ converges in probability to $D(\mathcal{C} \mid P)$ as $m, n \rightarrow \infty$;


## Data depth for an unparametrized curve: consistency

## Theorem

Let $\mathcal{C} \in \mathfrak{B}$ be a rectifiable curve, and let $P$ be a probability measure in the space of curves such that $P \in \mathcal{P}$. Let $\left(\Delta_{m}\right)$ be a decreasing sequence of positive numbers such that $\left(\Delta_{m}\right)$ and $\left(\sqrt{\frac{\log (m)}{m}} / \Delta_{m}^{2}\right)$ converges to zero when $m \rightarrow \infty$.

Then:

- the Monte Carlo approximation $\widehat{D}_{n, m, \Delta_{m}}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ converges in probability to $D\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ as $m \rightarrow \infty$;
- the Monte Carlo approximation $\widehat{D}_{n, m, \Delta_{m}}\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ converges in probability to $D(\mathcal{C} \mid P)$ as $m, n \rightarrow \infty$;
- the sample Tukey curve depth $D\left(\mathcal{C} \mid \mathcal{X}_{1}, \ldots, \mathcal{X}_{n}\right)$ converges in probability to $D(\mathcal{C} \mid P)$ as $n \rightarrow \infty$.


## Data depth for an unparametrized curve: properties

Restrict to $\mathfrak{B}_{\ell}$, the subset of unparametrized curves of positive length bounded by $\ell>0$. Then the Tukey curve depth satisfies the following properties:

- Nonnegativity and boundedness by one:

$$
D\left(\mathcal{C} \mid Q_{P}\right) \in[0,1] .
$$

## Data depth for an unparametrized curve: properties

Restrict to $\mathfrak{B}_{\ell}$, the subset of unparametrized curves of positive length bounded by $\ell>0$. Then the Tukey curve depth satisfies the following properties:

- Nonnegativity and boundedness by one:

$$
D\left(\mathcal{C} \mid Q_{P}\right) \in[0,1] .
$$

- Similarity invariance: Let $f: \mathbb{R}^{d} \rightarrow \mathbb{R}^{d} f$ be a similarity, i.e. there exists an orthogonal matrix $A$, a factor $r>0$ and a vector $\boldsymbol{b} \in \mathbb{R}^{d}$ such that for all $\boldsymbol{x} \in \mathbb{R}^{d}, f(\boldsymbol{x})=r A \boldsymbol{x}+\boldsymbol{b}$. In particular for all $\boldsymbol{x}$ and $\boldsymbol{y}$ in $\mathbb{R}^{d},|f(\boldsymbol{x})-f(\boldsymbol{y})|_{2}=r|\boldsymbol{x}-\boldsymbol{y}|_{2}$. Denote by $P_{f}$ the distribution of the image through $f$ of a stochastic process having a distribution $P$. Then

$$
D\left(f \circ \mathcal{C} \mid Q_{P_{f}}\right)=D\left(\mathcal{C} \mid Q_{P}\right) .
$$

## Data depth for an unparametrized curve: properties

Restrict to $\mathfrak{B}_{\ell}$, the subset of unparametrized curves of positive length bounded by $\ell>0$. Then the Tukey curve depth satisfies the following properties:

- Nonnegativity and boundedness by one:

$$
D\left(\mathcal{C} \mid Q_{P}\right) \in[0,1] .
$$

- Similarity invariance: Let $f: \mathbb{R}^{d} \rightarrow \mathbb{R}^{d} f$ be a similarity, i.e. there exists an orthogonal matrix $A$, a factor $r>0$ and a vector $\boldsymbol{b} \in \mathbb{R}^{d}$ such that for all $\boldsymbol{x} \in \mathbb{R}^{d}, f(\boldsymbol{x})=r A \boldsymbol{x}+\boldsymbol{b}$. In particular for all $\boldsymbol{x}$ and $\boldsymbol{y}$ in $\mathbb{R}^{d},|f(\boldsymbol{x})-f(\boldsymbol{y})|_{2}=r|\boldsymbol{x}-\boldsymbol{y}|_{2}$. Denote by $P_{f}$ the distribution of the image through $f$ of a stochastic process having a distribution $P$. Then

$$
D\left(f \circ \mathcal{C} \mid Q_{P_{f}}\right)=D\left(\mathcal{C} \mid Q_{P}\right) .
$$

- Vanishing at infinity:

$$
\lim _{d_{\mathbb{E}}(\mathcal{C}, \mathbf{0}) \rightarrow \infty, \mathcal{C} \in \mathfrak{B}_{\ell}} D\left(\mathcal{C}, Q_{P}\right)=\inf _{\mathcal{C} \in \mathfrak{B}_{\ell}} D\left(\mathcal{C}, Q_{P}\right)=0
$$

## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties

## Illustrations

Brain imaging
Practical session

Binary supervised classification: MNIST ("0" vs " 1 ")
Some examples:


Given: training sample $\mathcal{S}_{0}=\left\{\mathcal{C}_{1}, \ldots, \mathcal{C}_{m}\right\}$ stemming from $P_{0}$ and $\mathcal{S}_{1}=\left\{\mathcal{C}_{m+1}, \ldots, \mathcal{C}_{m+n}\right\}$ stemming from $P_{1}$ in $\mathfrak{B}$.

Find: classifier $g: \mathfrak{B} \rightarrow\{0,1\}$ best separating $P_{0}$ and $P_{1}$.

Binary supervised classification: MNIST ("0" vs " 1 ")
Consider DD-plot (Li, Cuesta-Albertos, Liu '12):

$$
\boldsymbol{Z}=\left\{\boldsymbol{z}_{i}: \boldsymbol{z}_{i}=\left(D\left(\mathcal{C}_{i} \mid Q_{P_{0}}\right), D\left(\mathcal{C}_{i} \mid Q_{P_{1}}\right)\right), i=1, \ldots, m+n\right\} .
$$

Binary supervised classification: MNIST ("0" vs " 1 ")
Consider DD-plot (Li, Cuesta-Albertos, Liu '12):

$$
\boldsymbol{Z}=\left\{\boldsymbol{z}_{i}: \boldsymbol{z}_{i}=\left(D\left(\mathcal{C}_{i} \mid Q_{P_{0}}\right), D\left(\mathcal{C}_{i} \mid Q_{P_{1}}\right)\right), i=1, \ldots, m+n\right\} .
$$



Binary supervised classification: MNIST ("0" vs " 1 ")
Consider DD-plot (Li, Cuesta-Albertos, Liu '12):

$$
\boldsymbol{Z}=\left\{\boldsymbol{z}_{i}: \boldsymbol{z}_{i}=\left(D\left(\mathcal{C}_{i} \mid Q_{P_{0}}\right), D\left(\mathcal{C}_{i} \mid Q_{P_{1}}\right)\right), i=1, \ldots, m+n\right\} .
$$



Binary supervised classification: MNIST ("0" vs " 1 ")
Consider DD-plot (Li, Cuesta-Albertos, Liu '12):

$$
\boldsymbol{Z}=\left\{\boldsymbol{z}_{i}: \boldsymbol{z}_{i}=\left(D\left(\mathcal{C}_{i} \mid Q_{P_{0}}\right), D\left(\mathcal{C}_{i} \mid Q_{P_{1}}\right)\right), i=1, \ldots, m+n\right\} .
$$



Unsupervised classification: MNIST ("0", "1", and "7")
Some examples:


Task: Find reasonable grouping with data depth (Jörnsten '04).

Unsupervised classification: MNIST ("0", " 1 ", and " 7 ")
Depth-based clustering (Jörnsten '04):

## Unsupervised classification: MNIST ("0", "1", and "7")

Depth-based clustering (Jörnsten '04):

- Let $\left\{\mathcal{C}_{1}, \ldots, \mathcal{C}_{\sum_{j} n_{j}}\right\}$ be the observed sample and let $I_{j}$, $j=1, \ldots, J$ denote the corresponding partitioning into $J$ clusters (indices of observations belonging to each cluster $j$ ) with $\cup_{j} I_{j}=\left\{1, \ldots, \sum_{j} n_{j}\right\}$ and $I_{j_{1}} \cap I_{j_{2}}=\emptyset$ for all $j_{1} \neq j_{2}$.


## Unsupervised classification: MNIST ("0", " 1 ", and " 7 ")

Depth-based clustering (Jörnsten '04):

- Let $\left\{\mathcal{C}_{1}, \ldots, \mathcal{C}_{\sum_{j} n_{j}}\right\}$ be the observed sample and let $I_{j}$, $j=1, \ldots, J$ denote the corresponding partitioning into $J$ clusters (indices of observations belonging to each cluster $j$ ) with $\cup_{j} l_{j}=\left\{1, \ldots, \sum_{j} n_{j}\right\}$ and $I_{j_{1}} \cap I_{j_{2}}=\emptyset$ for all $j_{1} \neq j_{2}$.
- Define the silhouette width of an observation $i$ belonging to cluster $j$ as

$$
\operatorname{Sil}_{i}^{j}=\frac{\bar{d}_{i}^{-j}-\bar{d}_{i}^{j}}{\max \left\{\bar{d}_{i}^{-j}, \bar{d}_{i}^{j}\right\}},
$$

where $\bar{d}_{i}^{j}=\frac{1}{\# I_{j}-1} \sum_{i^{\prime} \in I_{j}, i^{\prime} \neq i} d_{\mathfrak{B}}\left(\mathcal{C}_{i}, \mathcal{C}_{i^{\prime}}\right)$ and $\bar{d}_{i}^{-j} \in \operatorname{argmin}_{j^{\prime} \neq j} \frac{1}{\# I_{j^{\prime}}} \sum_{i^{\prime} \in I_{j^{\prime}}} d_{\mathfrak{B}}\left(\mathcal{C}_{i}, \mathcal{C}_{i^{\prime}}\right)$ are average distances to the observations in its own cluster and in the closest among foreign clusters respectively.

## Unsupervised classification: MNIST ("0", " 1 ", and " 7 ")

Depth-based clustering (Jörnsten '04):

- Let $\left\{\mathcal{C}_{1}, \ldots, \mathcal{C}_{\sum_{j} n_{j}}\right\}$ be the observed sample and let $I_{j}$, $j=1, \ldots, J$ denote the corresponding partitioning into $J$ clusters (indices of observations belonging to each cluster $j$ ) with $\cup_{j} l_{j}=\left\{1, \ldots, \sum_{j} n_{j}\right\}$ and $l_{j_{1}} \cap l_{j_{2}}=\emptyset$ for all $j_{1} \neq j_{2}$.
- Define the silhouette width of an observation $i$ belonging to cluster $j$ as

$$
\operatorname{Sil}_{i}^{j}=\frac{\bar{d}_{i}^{-j}-\bar{d}_{i}^{j}}{\max \left\{\bar{d}_{i}^{-j}, \bar{d}_{i}^{j}\right\}},
$$

where $\bar{d}_{i}^{j}=\frac{1}{\# I_{j}-1} \sum_{i^{\prime} \in I_{j}, i^{\prime} \neq i} d_{\mathfrak{B}}\left(\mathcal{C}_{i}, \mathcal{C}_{i^{\prime}}\right)$ and
$\bar{d}_{i}^{-j} \in \operatorname{argmin}_{j^{\prime} \neq j} \frac{1}{\# l_{j^{\prime}}} \sum_{i^{\prime} \in I_{j^{\prime}}} d_{\mathfrak{B}}\left(\mathcal{C}_{i}, \mathcal{C}_{i^{\prime}}\right)$ are average distances
to the observations in its own cluster and in the closest among foreign clusters respectively.

- The relative depth is defined as

$$
\operatorname{Re}_{i}^{j}=D\left(\mathcal{C}_{i} \mid\left\{\mathcal{C}_{i^{\prime}}\right\}_{i^{\prime} \in I_{j}}\right)-\max _{j^{\prime} \neq j} D\left(\mathcal{C}_{i} \mid\left\{\mathcal{C}_{i^{\prime}}\right\}_{i^{\prime} \in I_{j^{\prime}}}\right)
$$

Unsupervised classification: MNIST ("0", " 1 ", and " 7 ")
Clustering criterion:

$$
C\left(\left\{I_{j}\right\}_{1}^{J}\right)=\frac{1}{\sum_{j} n_{j}} \sum_{j=1}^{J} \sum_{i \in I_{j}} c_{i}\left(\left\{I_{j}\right\}_{1}^{J}\right),
$$

with the observation-wise clustering criterion:

$$
c_{i}\left(\left\{I_{j}\right\}_{1}^{J}\right)=(1-\lambda) \operatorname{Sil}_{i}^{j}+\lambda \operatorname{Re} D_{i}^{j} .
$$



## Comparison with functional depth: Example 1

Simulated S letters: depth-induced ranking


## Comparison with functional depth: Example 2

Simulated hurricane tracks: curve boxplot

$$
\text { MFHD - time } \quad \text { MFHD - length }
$$


mSBD - time


mSBD - length


Curve depth


Comparison with functional depth: Anomaly detection 1
Data set 1 with introduced anomalies:


Ordered depth values:


Comparison with functional depth: Anomaly detection 2
Data set 2 with introduced anomalies:


Ordered depth values:


## Contents

Anomaly detection in functional framework
Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging
Practical session

## Application: brain imaging - OATS data

- The Older Australian Twins Study (OATS) includes diffusion tensor magnetic resonance images (DTI) of 34 twin pairs: 11 dizygotic (DZ) and 23 monozygotic (MZ).


## Application: brain imaging - OATS data

- The Older Australian Twins Study (OATS) includes diffusion tensor magnetic resonance images (DTI) of 34 twin pairs: 11 dizygotic (DZ) and 23 monozygotic (MZ).
- For each individual, $\mathbf{1 0 0 0}$ fiber tracts connecting the motor cortex with the brain stem extracted for each hemisphere.


## Application: brain imaging - OATS data

- The Older Australian Twins Study (OATS) includes diffusion tensor magnetic resonance images (DTI) of 34 twin pairs: 11 dizygotic (DZ) and 23 monozygotic (MZ).
- For each individual, $\mathbf{1 0 0 0}$ fiber tracts connecting the motor cortex with the brain stem extracted for each hemisphere.

Questions to answer:

- Information compression for better understanding of brain functioning.


## Application: brain imaging - OATS data

- The Older Australian Twins Study (OATS) includes diffusion tensor magnetic resonance images (DTI) of 34 twin pairs: 11 dizygotic (DZ) and 23 monozygotic (MZ).
- For each individual, $\mathbf{1 0 0 0}$ fiber tracts connecting the motor cortex with the brain stem extracted for each hemisphere.

Questions to answer:

- Information compression for better understanding of brain functioning.
- Outlier detection for indication of wrongly tracked fibers.


## Application: brain imaging - OATS data

- The Older Australian Twins Study (OATS) includes diffusion tensor magnetic resonance images (DTI) of 34 twin pairs: 11 dizygotic (DZ) and 23 monozygotic (MZ).
- For each individual, $\mathbf{1 0 0 0}$ fiber tracts connecting the motor cortex with the brain stem extracted for each hemisphere.

Questions to answer:

- Information compression for better understanding of brain functioning.
- Outlier detection for indication of wrongly tracked fibers.
- Curve registration for aligning data from different individuals before further analysis.


## Application: brain imaging - OATS data

- The Older Australian Twins Study (OATS) includes diffusion tensor magnetic resonance images (DTI) of 34 twin pairs: 11 dizygotic (DZ) and 23 monozygotic (MZ).
- For each individual, $\mathbf{1 0 0 0}$ fiber tracts connecting the motor cortex with the brain stem extracted for each hemisphere.

Questions to answer:

- Information compression for better understanding of brain functioning.
- Outlier detection for indication of wrongly tracked fibers.
- Curve registration for aligning data from different individuals before further analysis.
- Studying genetic dependency (DZ vs. MZ) for identifying disease causes.


## Application: brain imaging



## Application: brain imaging - depth-based ordering



## Application: brain imaging - information compression



Application: brain imaging, right stem - outlier detection


## Application: brain imaging, right stem - registration



Subject 110


Subject 131


- The red and the dark blue curves are respectively the deepest curves before registration of the respective subject and subject 235, the subject whose deepest curve is the deepest of all.
- We bring the red curve as close as possible (in terms of the distance) to the black curve. The transformed curve (after registration) is the light blue curve.
- Distances from each curve to the deepest one (dark blue) before (red) and after (light blue) registration are 10.271 and 3.245 (for subject 104), 4.539 and 3.395 (for subject 110), 3.329 and 2.084 (for subject 131), respectively.

Application: brain imaging, right stem - twins comparison





106 vs. 206 (MZ)
131 vs. 231 (MZ)



## Contents

## Anomaly detection in functional framework

Functional isolation forest
The method
FIF parameters
Real data benchmarking
Extension of FIF: Connection to data depth
Data depth: the integrated approach
Depth for curve data
Motivation
Methodology
Computation and properties
Illustrations
Brain imaging

## Practical session

## Thank you for attention! (and a short list of literature)

- Chandola, V., Banerjee, A., and Kumar, V. (2009). Anomaly detection: A survey. ACM Computing Surveys (CSUR), 41(3):15, 1-58.
- Breunig, M.M., Kriegel, H.-P., Ng, R.T., and Sander, J. (2000). LOF: Identifying density-based local outliers. In: Proceedings of the 2000 ACM SIGMOD International Conference on Management of Data, 29, 93-104.
- Schölkopf, B., Platt, J.C., Shawe-Taylor, J., Smola, A., and Williamson, R. (2001). Estimating the support of a high-dimensional distribution. Neural Computation, 13(7), 1443-1471.
- Liu, F.T., Ting, K.M., and Zhou, Z. (2008). Isolation forest. In: Proceedings of the 2008 Eighth IEEE International Conference on Data Mining, 413-422.
- Mosler, K. (2013). Depth statistics. In: Robustness and Complex Data Structures: Festschrift in Honour of Ursula Gather, 17-34.
- Hubert, M., Rousseeuw, P.J., and Segaert, P. (2015). Multivariate functional outlier detection. Statistical Methods \& Applications, 24(2), 177-202.


## Practical session (part II)

Notebooks:

- anomdet_simulation1.Rmd,
- anomdet_hurricanes.Rmd,
- anomdet_cars.ipynb,
- anomdet_airbus.ipynb.

Data sets:

- carsanom.csv: Data set on anomaly detection for cars.
- airbus_data.csv: Data set from Airbus.
$\rightarrow$ hurdat2-1851-2019-052520.txt: Historical hurricane data.
Supplementary scripts:
- depth routines.py: Routines for data depth calculation.
- FIF.py: Implementation of the functional isolation forest.
- depth_routines.R: Routines for curves' parametrization.


## Literature (mentioned in the tutorial) (1)

- Boser, B.E., Guyon, I., and Vapnik, V.N. (1992). A training algorithm for optimal margin classifiers. In: Proceedings of the Fifth Annual Workshop of Computational Learning Theory, Pittsburgh, ACM, 5, 144-152.
- Breunig, M.M., Kriegel, H.-P., Ng, R.T., and Sander, J. (2000). LOF: Identifying density-based local outliers. In: Proceedings of the 2000 ACM SIGMOD International Conference on Management of Data, 29, 93-104.
- Chandola, V., Banerjee, A., and Kumar, V. (2009). Anomaly detection: A survey. ACM Computing Surveys (CSUR), 41(3):15, 1-58.
- Chaudhuri P. (1996). On a geometric notion of quantiles for multivariate data. Journal of the American Statistical Association, 91, 862-872.
- Claeskens, G., Hubert, M., Slaets, L., and Vakili, K. (2014). Multivariate functional halfspace depth. Journal of the American Statistical Association, 109(505), 411-423.
- Cortes, C. and Vapnik, V. (1995). Support-vector networks. Machine Learning, 20, 273-297.
- Donoho D. (1982). Breakdown Properties of Multivariate Location Estimators. Ph.D. thesis, Harvard University.
- Donoho D.L., Gasko M. (1992). Breakdown properties of location estimates based on halfspace depth and projected outlyingness. The Annals of Statistics, 20, 1803-1827.


## Literature (mentioned in the tutorial) (2)

- Fraiman, R. and Muniz, G. (2001). Trimmed means for functional data. TEST, 10, 419-440.
- Hariri, S., Carrasco Kind, M., and Brunner, R.J. (2018). Extended isolation forest. arXiv:1811.02141.
- Hubert, M., Rousseeuw, P.J., and Segaert, P. (2015). Multivariate functional outlier detection. Statistical Methods \& Applications, 24(2), 177-202.
- Koltchinskii V. (1997). M-estimation, convexity and quantiles. The Annals of Statistics, 25, 435-477.
- Koshevoy G., Mosler K. (1997). Zonoid trimming for multivariate distributions. The Annals of Statistics, 25, 1998-2017.
- Liu R.Y. (1990). On a notion of data depth based on random simplices. The Annals of Statistics, 18, 405-414.
- Liu, Z. and Modarres, R. (2011). Lens data depth and median. Journal of Nonparametric Statistics, 23, 1063-1074.
- Liu, F.T., Ting, K.M., and Zhou, Z. (2008). Isolation forest. In: Proceedings of the 2008 Eighth IEEE International Conference on Data Mining, 413-422.


## Literature (mentioned in the tutorial) (3)

- López-Pintado, S. and Romo, J. (2009). On the concept of depth for functional data. Journal of the American Statistical Association, 104(486), 718-734.
- Mahalanobis P.C. (1936). On the generalized distance in statistics. Proceedings of the National Institute of Sciences of India, 12, 49-55.
- Markou, M. and Singh, S. (2003). Novelty detection: a review - Part 1: Statistical approaches. Signal Processing, 83(12), 2481-2497.
- Markou, M. and Singh, S. (2003). Novelty detection: a review - Part 2: Neural network based approaches. Signal Processing, 83(12), 2499-2521.
- Miljković, D. (2010). Review of novelty detection methods. The 33rd International Convention MIPRO, Opatija, 593-598.
- Mosler, K. (2013). Depth statistics. In: Robustness and Complex Data Structures: Festschrift in Honour of Ursula Gather, 17-34.
- Oja, H. (1983). Descriptive statistics for multivariate distributions. Statistics and Probability Letters, 1, 327-332.
- Pimentel, M.A.F., Clifton, D.A., Clifton, L., and Tarassenko, L. (2014). A review of novelty detection. Signal Processing, 99, 215-249.
- Schölkopf, B., Platt, J.C., Shawe-Taylor, J., Smola, A., and Williamson, R. (2001). Estimating the support of a high-dimensional distribution. Neural Computation, 13(7), 1443-1471.


## Literature (mentioned in the tutorial) (4)

- Serfling, R. (2002). A depth function and a scale curve based on spatial quantiles. In: Statistical Data Analysis Based on the $L_{1}$-Norm and Related Methodsm Birkhäser, Basel, 25-38.
- Stahel W. (1981). Robust Estimation: Infinitesimal Optimality and Covariance Matrix Estimators (In German). Ph.D. thesis, Swiss Federal Institute of Technology in Zurich.
- Tukey J.W. (1975). Mathematics and the picturing of data. In: Proceedings of the International Congress of Mathematicians, volume 2, Canadian Mathematical Congress, 523-531.
- Vapnik, V. and Chervonenkis, A. (1974). Theory of Pattern Recognition (in Russian). Nauka, Moscow.
- Vapnik, V. and Lerner, A. (1963). Pattern recognition using generalized portraits. Avtomatika i Telemekhanika, 24, 774-780.
- Vardi Y., Zhang C. (2000). The multivariate $L_{1}$-median and associated data depth. Proceedings of the National Academy of Sciences of the United States of America, 97, 1423-1426.
- Zuo Y., Serfling R. (2000). General notions of statistical depth function. The Annals of Statistics, 28, 461-482.

