



# Audio source separation



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M2 MVA Audio signal analysis, indexing and transformation

> Art of estimating "source" signals, assumed independent, from the observation of one or

# Part I

## Introduction



#### Introduction

Source separation

Application examples:

several "mixtures" of these sources

Remix, transformations, re-spatialization

Denoising (cocktail party, suppression of vuvuzela, karaoke)

Separation of the instruments in polyphonic music

2/48	Une école de l'IMT	Audio source separation
	Typology of the mix	ture models

- Definition of the problem
  - Observations: *M* mixtures  $x_m(t)$ , concatenated in a vector  $\mathbf{x}(t)$
  - Unknowns: K sources  $s_k(t)$ , concatenated in a vector  $\mathbf{s}(t)$
  - General mixture model: function  $\mathscr{A}$  which transforms  $\mathbf{s}(t)$  into  $\mathbf{x}(t)$
- ▶ Stationarity: *A* is translation invariant
- ► Linearity: 𝒴 is a linear map
- Memory:
  - Convolutive mixtures
  - Instantaneous mixtures:  $\mathbf{x}(t) = \mathbf{A}\mathbf{s}(t)$ 
    - ▶  $\mathscr{A}$  is defined by the "mixture matrix" **A** (of dimension  $M \times K$ )
- Inversibility:
  - Determined mixtures: M = K
  - Over-determined mixtures: M > K
  - Under-determined mixtures: M < K

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(a) Convolutive mixture

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(b) Binaural mixture

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- Notation:  $\phi[\mathbf{x}]$  denotes a function of  $p(\mathbf{x})$
- Mean vector:  $\mu_x = \mathbb{E}[\mathbf{x}]$
- Covariance matrix:  $\mathbf{\Sigma}_{xx} = \mathbb{E}[(\mathbf{x} \boldsymbol{\mu}_x)(\mathbf{x} \boldsymbol{\mu}_x)^T]$
- Characteristic function:  $\phi_{\mathbf{x}}(\mathbf{f}) = \mathbb{E}[e^{-2i\pi\mathbf{f}^T\mathbf{x}}] = \int_{\mathbb{R}} p(\mathbf{x})e^{-2i\pi\mathbf{f}^T\mathbf{x}}d\mathbf{x}$
- Probability distribution:  $p(\mathbf{x}) = \int_{\mathbb{D}} \phi_{\mathbf{x}}(\mathbf{f}) e^{+2i\pi \mathbf{f}^T \mathbf{x}} d\mathbf{f}$
- Cumulants:
  - Definition:  $\ln(\phi_x(\mathbf{f})) = \sum_{n=1}^{+\infty} \frac{(-2i\pi)^n}{n!} \sum_{k_1=1}^K \sum_{k_n=1}^K \kappa_{k_1\dots k_n}^n [\mathbf{x}] f_{k_1} \dots f_{k_n}$

  - $\kappa^n[\mathbf{x}]$  is an *n*-th order tensor  $\kappa^1[\mathbf{x}]$  is the mean vector,  $\kappa^2[\mathbf{x}]$  is the covariance matrix
  - If  $p(\mathbf{x})$  is symmetric  $(p(-\mathbf{x}) = p(\mathbf{x}))$ ,  $\kappa^n[\mathbf{x}] = 0$  for any odd value n
  - the ratio  $\kappa_{k,k,k,k}^4[\mathbf{x}]/(\kappa_{k,k}^2[\mathbf{x}])^2$  is called "kurtosis"

- The Gaussian distribution is the one such that all cumulants of order n > 2 are zero
- Characteristic function

$$\phi_{\mathsf{x}}(\mathbf{f}) = \exp(-2i\pi\mathbf{f}^{\mathsf{T}}\boldsymbol{\mu}_{\mathsf{x}} - 2\pi^{2}\mathbf{f}^{\mathsf{T}}\boldsymbol{\Sigma}_{\mathsf{xx}}\mathbf{f})$$

• Probability density function (defined if  $\Sigma_{xx}$  is invertible)

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{K}{2}} \det(\mathbf{\Sigma}_{xx})^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_{x})^{T} \mathbf{\Sigma}_{xx}^{-1}(\mathbf{x} - \boldsymbol{\mu}_{x})\right)$$

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	WSS vector processe	25		Information theory		_

- Definition: the cumulants of orders 1 et 2 are translation-invariant
- Covariance matrices of 2 centered WSS processes  $\mathbf{x}(t)$  and  $\mathbf{y}(t)$ :
  - Definition:  $\mathbf{R}_{xy}(\tau) = \mathbb{E}\left[\mathbf{x}(t+\tau)\mathbf{y}(t)^{T}\right]$
  - Property:  $\mathbf{R}_{xx}(0) = \boldsymbol{\Sigma}_{xx}$  is Hermitian and positive semi-definite.
- PSD matrices of a WSS process x(t):
  - Definition:  $\mathbf{S}_{xx}(v) = \sum_{\tau \in \mathbb{Z}_{+}} \mathbf{R}_{xx}(\tau) e^{-2i\pi v \tau}$
  - Property:  $\forall v, \mathbf{S}_{xx}(v)$  is Hermitian and positive semi-definite

- Shannon entropy
  - Definition:  $\mathbb{H}[\mathbf{x}] = -\mathbb{E}[\ln(p(\mathbf{x}))]$
  - H[x] is not necessarily non-negative for a continuous r.v.
- Kullback-Leibler divergence
  - $D_{KL}(p||q) = \int p(\mathbf{x}) \ln\left(\frac{p(\mathbf{x})}{q(\mathbf{x})}\right) d\mathbf{x}$

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- Property:  $D_{KL}(p||q) \ge 0$ ,  $D_{KL}(p||q) = 0$  if and only if p = q
- Mutual information
  - ► Definition:  $\mathbb{I}[\mathbf{x}] = \mathbb{E}\left[\ln\left(\frac{p(\mathbf{x})}{p(x_1)...p(x_K)}\right)\right] = D_{KL}(p(\mathbf{x})||p(x_1)...p(x_K))$
  - Property:  $\mathbb{I}[\mathbf{x}] = 0$  if and only if  $x_1 \dots x_K$  are mutually independent
  - Relationship with entropy:  $\mathbb{I}[\mathbf{x}] = \sum_{k=1}^{K} \mathbb{H}[x_k] \mathbb{H}[\mathbf{x}]$



11/48

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- Sources are assumed IID:  $p({s_k(t)}_{k,t}) = \prod_{k=1}^{K} \prod_{t=1}^{T} p_k(s_k(t))$
- Problem: estimate **A** and sources  $\mathbf{s}(t)$  given  $\mathbf{x}(t)$
- Definition: non-mixing matrix
  - a matrix **C** of dimension  $K \times K$  is non-mixing if and only if it has a unique non-zero entry in each row and each column
- If s̃(t) = Cs(t) and à = AC<sup>-1</sup>, then x(t) = Ãs̃(t) is another admissible decomposition of the observations
  - Sources can be recovered up to a permutation and a multiplicative factor





• Let  $\mathbf{y}(t) = \mathbf{B}\mathbf{x}(t)$ , where  $\mathbf{B} \in \mathbb{R}^{K \times M}$  is referred to as the "separation matrix"

Part III

Linear instantaneous mixtures

#### • Linear separation is feasible if **A** has rank *K*:

- We get  $\mathbf{y}(t) = \mathbf{s}(t)$  by defining:
  - $\mathbf{B} = \mathbf{A}^{-1}$  in the determined case (M = K)
  - $\mathbf{B} = \mathbf{A}^{\dagger}$  in the over-determined case (M > K)
- the pseudo-inverse  $\mathbf{A}^{\dagger} = (\mathbf{A}^{T}\mathbf{A})^{-1}\mathbf{A}^{T}$  is such that  $\mathbf{A}^{\dagger}\mathbf{A} = \mathbf{I}_{K}$
- In the under-determined case (M < K), separation is not feasible



Independent component analysis



15/48

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#### In practice matrix A is unknown:

- We look for a matrix **B** that makes the  $y_k$  independent (ICA)
- We then get equation  $\mathbf{y}(t) = \mathbf{Cs}(t)$ , where  $\mathbf{C} = \mathbf{BA}$
- The problem is solved if matrix C is non-mixing



- Theorem (identifiability)
  - Let  $s_k$  be K IID sources, among which at most one is Gaussian, and  $\mathbf{y}(t) = \mathbf{Cs}(t)$  with **C** invertible ((over)-determined case). If signals  $y_k(t)$  are independent, then **C** is non-mixing.

- ▶ We now suppose that the sources are centered:  $\mathbb{E}[\mathbf{s}(t)] = \mathbf{0}$  and that the mixture is (over-)determined
- Canonical problem: we can assume without loss of generality that s(t) is spatially white  $(\mathbf{\Sigma}_{ss} = \mathbb{E}[\mathbf{s}(t)\mathbf{s}(t)^T] = \mathbf{I}_K)$
- Then  $\Sigma_{xx} = A\Sigma_{ss}A^T = AA^T$ : A is a matrix square root of  $\Sigma_{xx}$
- We first aim to whiten (decorrelate) the mixture:
  - $\Sigma_{xx}$  is diagonalizable in an orthonormal basis:  $\Sigma_{xx} = Q\Lambda^2 Q^T$  where  $\Lambda = \text{diag}(\lambda_1 \dots \lambda_M)$ with  $\lambda_1 \ge \lambda_K > \lambda_{K+1} = \lambda_M = 0$  (the rank of  $\Sigma_{xx}$  is equal to K) • Let  $\mathbf{S} = \mathbf{Q}_{(:.1:K)} \mathbf{\Lambda}_{(1:K,1:K)} \in \mathbb{R}^{M \times K}$

  - **S** is a matrix square root of  $\Sigma_{xx}$ :  $\Sigma_{xx} = SS^T$
  - Let  $\mathbf{W} = \mathbf{S}^{\dagger}$  and  $\mathbf{z}(t) = \mathbf{W}\mathbf{x}(t)$
  - Then z(t) is white  $(\mathbb{E}[z(t)] = 0$  and  $\Sigma_{zz} = W\Sigma_{xx}W^T = I)$

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	Whitening			Higher order statistic	CS.	

- We conclude without loss of generality that  $U \triangleq WA$  is a rotation matrix  $(\mathbf{U}\mathbf{U}^T = \mathbf{I}).$
- Then  $\mathbf{y}(t) = \mathbf{U}^T \mathbf{z}(t) = \mathbf{U}^T \mathbf{W} \mathbf{x}(t) = (\mathbf{W} \mathbf{A})^{-1} (\mathbf{W} \mathbf{A}) \mathbf{s}(t) = \mathbf{s}(t)$ .
- We can thus assume  $\mathbf{B} = \mathbf{U}^T \mathbf{W}$  where  $\mathbf{U}$  is a rotation matrix.



- One can estimate  $\Sigma_{xx}$  from the observations and get W
- ▶ The whiteness property (second order cumulants) determines W and leaves U unknown.
- $\blacktriangleright$  If sources are Gaussian, the  $z_k$  are independent and **U** cannot be determined.
- ▶ In order to determine rotation U, we need to exploit the non-Gaussianity of sources and characterize the independence property by using cumulants of order greater than 2.



## **Contrast functions**



- Definition:  $\phi$  is a "contrast function" if and only if  $\phi[\mathbf{Cs}(t)] \ge \phi[\mathbf{s}(t)] \forall \mathbf{C}$  and if  $\phi[\mathbf{Cs}(t)] = \phi[\mathbf{s}(t)] \Leftrightarrow \mathbf{C}$  is non-mixing.
- Separation is performed by minimizing  $\phi[\mathbf{y}(t) = \mathbf{Cs}(t)]$  with respect to **U** (or **B**)
- "Canonical" contrast function:  $\phi_{IM}[\mathbf{y}(t)] = \mathbb{I}[\mathbf{y}(t)]$
- Orthogonal contrasts: to be minimized under the constraint  $\mathbb{E}[\mathbf{y}(t)\mathbf{y}(t)^T] = \mathbf{I}$ . For instance,  $\phi_{IM}^{\circ}[\mathbf{y}(t)] = \sum_{k=1}^{K} \mathbb{H}(y_k(t))$
- ► Order 4 approximation of  $\phi_{IM}^{\circ}$ :  $\phi_{ICA}^{\circ}[\mathbf{y}(t)] = \sum_{ijkl \neq iiii} (\kappa_{ijkl}^{4}[\mathbf{y}(t)])^{2}$
- Descent algorithms for minimizing  $\phi$  with respect to **B** or **U**:
  - Gradient algorithm applied to matrix B
  - Parameterization of U with Givens rotations and coordinate descent

- 1. Estimation of the covariance matrix  $\Sigma_{xx}$
- 2. Diagonalization of  $\Sigma_{xx}$ :  $\Sigma_{xx} = Q\Lambda^2 Q^T$  where  $\Lambda = \text{diag}(\lambda_1 \dots \lambda_M)$  with  $\lambda_1 \geq \ldots \geq \lambda_M \geq 0$
- 3. Computation of  $\mathbf{S} = \mathbf{Q}_{(:,1;K)} \mathbf{\Lambda}_{(1;K,1;K)}$
- 4. Computation of the whitening matrix  $\mathbf{W} = \mathbf{S}^{\dagger}$
- 5. Data whitening:  $\mathbf{z}(t) = \mathbf{W}\mathbf{x}(t)$
- 6. Estimation of **U** by minimizing the contrast function  $\phi^{\circ}$
- 7. Estimation of source signals via  $\mathbf{y}(t) = \mathbf{U}^T \mathbf{z}(t)$

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				Temporal coherence	of sources	

Part V	• Model: $\mathbb{E}(\mathbf{s}(t)) = 0$ , $\mathbf{R}_{ss}(\tau) = \mathbb{E}(\mathbf{s}(t+\tau)\mathbf{s}(t)^T) = \operatorname{diag}(r_{s_k}(\tau))$
	• Canonical problem: we assume that $\Sigma_{ss} = \mathbf{R}_{ss}(0) = \mathbf{I}$
	We first aim to spatially whiten the mixture:
Second order methods	• Let <b>S</b> be a matrix square root of $\Sigma_{xx}$

- Let  $\mathbf{W} = \mathbf{S}^{\dagger}$  and  $\mathbf{z}(t) = \mathbf{W}\mathbf{x}(t)$
- Since  $\Sigma_{xx} = AA^T$ ,  $U \triangleq WA$  is a rotation matrix
- However,  $\forall \tau \in \mathbb{Z}$ ,  $\mathbf{R}_{zz}(\tau) = \mathbf{U}\mathbf{R}_{ss}(\tau)\mathbf{U}^T$
- The joint diagonalization of matrices  $\mathbf{R}_{zz}(\tau)$  for various values of  $\tau$  permits us to identify rotation  $\mathbf{U}$



21



### Joint diagonalization



- Unicity theorem :
  - ► Let a set of matrices  $\mathbf{R}_{zz}(\tau)$  of dimension  $K \times K$  and of the form  $\mathbf{R}_{zz}(\tau) = \mathbf{U}\mathbf{R}_{ss}(\tau)\mathbf{U}^{T}$ with  $\mathbf{U}$  unitary and  $\mathbf{R}_{ss}(\tau) = \operatorname{diag}(r_{s_{k}}(\tau))$ . Then  $\mathbf{U}$  is unique (up to a non-mixing matrix) if and only if  $\forall 1 \le k \ne l \le K$ , there is  $\tau$  such that  $r_{s_{k}}(\tau) \ne r_{s_{l}}(\tau)$
- Joint diagonalization methods: minimize the criterion

 $J(\mathbf{U}) = \sum_{\tau} \|\mathbf{U}^{\mathsf{T}} \mathbf{R}_{zz}(\tau) \mathbf{U} - \operatorname{diag}(\mathbf{U}^{\mathsf{T}} \mathbf{R}_{zz}(\tau) \mathbf{U})\|_{F}^{2}$ 

 $\blacktriangleright$  Parameterization of  $\boldsymbol{U}$  with Givens rotations and coordinate descent

- Second Order Blind Identification (SOBI)
  - 1. Estimation and diagonalization of  $\Sigma_{xx}$ :  $\Sigma_{xx} = Q \Lambda^2 Q^T$  where  $\Lambda = \text{diag}(\lambda_1 \dots \lambda_M)$  with  $\lambda_1 \ge \dots \ge \lambda_M \ge 0$
  - 2. Computation of  $\mathbf{S} = \mathbf{Q}_{(:,1:\mathcal{K})} \mathbf{\Lambda}_{(1:\mathcal{K},1:\mathcal{K})}$
  - 3. Computation of the whitening matrix  $\mathbf{W} = \mathbf{S}^{\dagger}$
  - 4. Data whitening:  $\mathbf{z}(t) = \mathbf{W}\mathbf{x}(t)$
  - 5. Estimation of covariance matrices  $\mathbf{R}_{zz}(\tau)$  for various delays  $\tau$
  - 6. Approximate joint diagonalization of matrices  $\mathbf{R}_{zz}(\tau)$  in a common basis U
  - 7. Estimation of source signals via  $\mathbf{y}(t) = \mathbf{U}^T \mathbf{z}(t)$

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	Non-stationarity of s	ources		Conclusion of the fir	rst part	

- Model:  $\mathbb{E}(\mathbf{s}(t)) = \mathbf{0}$ ,  $\mathbf{\Sigma}_{ss}(t) \triangleq \mathbb{E}(\mathbf{s}(t)\mathbf{s}(t)^{\mathsf{T}}) = \operatorname{diag}(\sigma_k^2(t))$
- Then  $\forall t \in \mathbb{Z}$ ,  $\mathbf{\Sigma}_{xx}(t) = \mathbf{A}\mathbf{\Sigma}_{ss}(t)\mathbf{A}^T$
- Joint diagonalization methods: minimize the criterion

$$J(\mathbf{B}) = \sum_{t} \|\mathbf{B}\boldsymbol{\Sigma}_{xx}(t)\mathbf{B}^{T} - \operatorname{diag}(\mathbf{B}\boldsymbol{\Sigma}_{xx}(t)\mathbf{B}^{T})\|_{F}^{2}$$

- Gradient descent algorithm applied to matrix B
- In the over-determined case, B must be constrained to span the principal subspace of all matrices Σ<sub>xx</sub>(t)
- Variant of the SOBI algorithm:

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- 1. Segmentation of source signals and estimation of covariance matrices  $\mathbf{\Sigma}_{\scriptscriptstyle X\!X}(t)$  on windows centered at different times t
- 2. Joint diagonalization of matrices  $\Sigma_{xx}(t)$  in a common basis B
- 3. Estimation of source signals via  $\mathbf{y}(t) = \mathbf{B}\mathbf{x}(t)$

- The use of higher order cumulants is only necessary for the non-Gaussian IID source model
- Second order statistics are sufficient for sources that are:
  - ▶ stationary but not IID ( $\rightarrow$  <u>spectral</u> dynamics)
  - ▶ non stationary (→ temporal dynamics)
- Remember that classical tools (based on second order statistics) are appropriate for blind separation of independent (and possibly Gaussian) sources, on condition that the spectral / temporal source dynamics is taken into account.



# Time-frequency representations

<b>D</b>	Motivations
Part VI	<ul> <li>Spectral and temporal dynamics are highlighted by a time-frequency (TF) representation of signals</li> </ul>
<b>T</b> : C	<ul> <li>TF representations are appropriate to process convolutive and/or under-determined mixtures</li> </ul>
I ime-frequency methods	Use of a filter bank (examples: STFT, MDCT):
	• Decomposition in F sub-bands and decimation of factor $T \leq F$
	• Analysis filters $h_f$ and synthesis filters $g_f$
	• TF representation of sources: $s_k(f,n) = (h_f * s_k)(nT)$

- TF representation of mixtures:  $x_m(f,n) = (h_f * x_m)(nT)$
- Perfect reconstruction:  $s_k(t) = \sum_{f=1}^F \sum_{n \in \mathbb{Z}} g_f(t-nT) s_k(f,n)$
- ▶ Then  $\forall f, n, \mathbf{x}(f, n) = \mathbf{As}(f, n)$  (same linear instantaneous mixture)

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	Non-stationary source	ce model		Separation method		

#### Assumption: independent and centered second order processes

- Model of non-stationary sources:
  - if the time frames n<sub>1</sub> and n<sub>2</sub> are disjoint, then s<sub>k</sub>(., n<sub>1</sub>) and s<sub>k</sub>(., n<sub>2</sub>) are uncorrelated and of distinct variances
- Model of WSS sources:
  - ▶ if sub-bands  $f_1$  and  $f_2$  are disjoint  $(h_{f_1} * \tilde{h}_{f_2} = 0)$ , then  $s_k(f_1,.)$  and  $s_k(f_2,.)$  are WSS, uncorrelated and of distinct variances  $\sigma_k^2(f_1) = (h_{f_1} * \tilde{h}_{f_1} * r_{s_k})(0)$  and
  - $\sigma_k^2(f_2) = (h_{f_2} * \tilde{h}_{f_2} * r_{s_k})(0)$
- Time-frequency source model:
  - ▶ all  $s_k(f, n)$  are uncorrelated for all n and f, of distinct variances  $\sigma_k^2(f, n)$  (⇒ time-frequency dynamics)

- Separation by joint matrix diagonalization:
  - Let  $\Sigma_{ss}(f,n) = \mathbb{E}[\mathbf{s}(f,n)\mathbf{s}(f,n)^T]$  and  $\Sigma_{xx}(f,n) = \mathbb{E}[\mathbf{x}(f,n)\mathbf{x}(f,n)^T]$
  - ► Then  $\Sigma_{xx}(f,n) = \mathbf{A}\Sigma_{ss}(f,n)\mathbf{A}^T$  where  $\Sigma_{ss}(f,n) = \operatorname{diag}(\sigma_k^2(f,n))$
- Variant of the SOBI algorithm:
  - 1. TF analysis of the mixtures:  $x_k(f,n) = (h_f * x_k)(nT)$
  - 2. Estimation of covariance matrices  $\Sigma_{xx}(f, n)$
  - 3. Joint diagonalization of matrices  $\Sigma_{xx}(f, n)$  in a basis B
  - 4. Estimation of the source signals via  $\mathbf{y}(f, n) = \mathbf{B}\mathbf{x}(f, n)$
  - 5. TF synthesis of the sources:  $y_k(t) = \sum_{f=1}^F \sum_{n \in \mathbb{Z}} g_f(t-nT) y_k(f,n)$



31/48

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# **Convolutive linear mixture**



## Convolutive mixtures

- Instantaneous mixture model unsuitable for real acoustic mixtures
- Let  $\mathbf{x}_k(f, n) \in \mathbb{R}^M$  be the image of source  $s_k(f, n)$ 
  - ▶ received multichannel signal if only source  $s_k(f, n)$  was active
- Mixture model:  $\mathbf{x}(f,n) = \sum_{k=1}^{K} \mathbf{x}_k(f,n)$
- Decomposition of the source separation problem
  - **separation**: estimate  $\mathbf{x}_k(f, n)$  from the mixture  $\mathbf{x}(f, n)$
  - deconvolution: estimate  $s_k(f, n)$  from  $x_k(f, n)$
- Mixture model:  $x_m(t) = \sum_{k=1}^{K} (a_{mk} * s_k)(t)$ , i.e.  $\mathbf{x}(t) = \mathbf{A} * \mathbf{s}(t)$
- Theorem (identifiability)
  - Let s<sub>k</sub> be K IID sources, among which at most one is Gaussian, and y(t) = C \* s(t) with C invertible ((over)-determined case). If signals y<sub>k</sub>(t) are independent, then C is non-mixing.

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33/48	Une école de l'IMT	Audio source separation	E IP PARIS 34/48	Une école de l'IMT	Audio source separation	🛞 IP PARIS
	Time-frequency appr	oach		Independent component	nt analysis	

- Mixture model:  $x_m(t) = \sum_{k=1}^{K} (a_{mk} * s_k)(t)$
- Assumptions:
  - the filter bank corresponds to an STFT
  - the IR of  $a_{mk}$  is short compared with the window length
  - $\forall m, k, f, a_{mk}(v)$  varies slowly compared with  $h_f(v)$
  - $\blacktriangleright \Rightarrow (h_f * a_{mk})(t) \approx a_{mk}(f) h_f(t)$
- Approximation of the convolutive mixture model:

$$x_m(f,n) = \sum_{k=1}^K a_{mk}(f) s_k(f,n)$$

- Matrix form:  $\mathbf{x}(f,n) = \mathbf{A}(f)\mathbf{s}(f,n)$ 
  - F instantaneous mixture models in every sub-band
  - $\blacktriangleright$   $\Rightarrow$  we can use any ICA method in every sub-band

- Let  $\mathbf{y}(f,n) = \mathbf{B}(f)\mathbf{x}(f,n)$ , where  $\mathbf{B}(f) \in \mathbb{C}^{K \times M}$
- Linear separation is feasible if A(f) has rank K:
  - We get  $\mathbf{y}(f, n) = \mathbf{s}(f, n)$  by defining:
    - $\mathbf{B}(f) = \mathbf{A}(f)^{-1}$  in the determined case (M = K)
    - $\mathbf{B}(f) = \mathbf{A}(f)^{\dagger}$  in the over-determined case (M > K)
- In the under-determined case (M < K), separation remains impossible
- ▶ In practice matrix **A**(*f*) is unknown:
  - We look for  $\mathbf{B}(f)$  that makes the  $y_k(f, n)$  independent (ICA)
  - We then get  $\mathbf{y}(f,n) = \mathbf{C}(f)\mathbf{s}(f,n)$ , where  $\mathbf{C}(f) = \mathbf{B}(f)\mathbf{A}(f)$
  - C(f) is non-mixing



### Indeterminacies



- Problem: indeterminacies (permutations and multiplicative factors) in matrices  $\mathbf{C}(f)$ 
  - ▶  $\forall k$ , identify indexes  $k_f$  such that  $\forall f$ ,  $y_{k_f}(f,n) = c_{k_f,k} s_k(f,n)$
  - identify the multiplicative factors  $c_{k_f,k}$
- Infinitely many solutions  $\Rightarrow$  need to constrain the model:
  - Assumptions on the mixture
    - continuity of the frequency responses  $a_{mk}(f)$  with respect to f
    - $\blacktriangleright$   $\rightarrow$  beamforming model and anechoic model
  - Assumptions on the sources
    - similarity of the temporal dynamics of  $\sigma_k^2(f, n)$

- Beamforming model:
  - Assumptions: plane waves, far field, no reverberation, linear antenna
  - Model:  $a_{mk}(f) = e^{-2i\pi f \tau_{mk}}$  where  $\tau_{mk} = \frac{d_m}{c} \sin(\theta_k)$
  - Parameters: positions  $d_m$  of the sensors and angles  $\theta_k$  of the sources
- Anechoic model:
  - Assumptions: punctual sources, no reverberation
  - Model:  $a_{mk}(f) = \alpha_{mk} e^{-2i\pi f \tau_{mk}}$  where  $\alpha_{mk} = \frac{1}{\sqrt{4\pi r_{mk}}}$  and  $\tau_{mk} = \frac{r_{mk}}{c}$  Parameters: distances  $r_{mk}$  between the sensors and sources

37/48	Une école de l'IMT	Audio source separation	IP PARIS 38/48	Une école de l'IMT	Audio source separation	E IP PARIS
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Part VIII

# Under-determined mixtures

- ▶ Usual case in audio: monophonic (M = 1) or stereophonic (M = 2) mixtures, with a number of sources K > M
- Convolutive mixture model:  $\mathbf{x}(f,n) = \mathbf{A}(f)\mathbf{s}(f,n)$  with M < K
- Assumption: the mixture model  $\mathbf{A}(f)$  and the source model  $\mathbf{\Sigma}_{ss}(f, n)$  are known
- Even in this case, the exact separation is not feasible, because there is no matrix  $\mathbf{B}(f)$  such that  $\mathbf{B}(f)\mathbf{A}(f) = \mathbf{I}_{K}$







# Separation via non-stationary filtering



- Let  $\mathbf{y}(f,n) = \mathbf{B}(f,n)\mathbf{x}(f,n)$  where  $\mathbf{B}(f,n) \in \mathbb{C}^{K \times M}$  depends on n
- ► Minimum Mean Square Error (MMSE) estimator: we look for B(f,n) which minimizes E[||y(f,n)-s(f,n)||<sub>2</sub><sup>2</sup>]
- ► Solution:  $\mathbf{B}(f,n) = \mathbf{\Sigma}_{sx}(f,n)\mathbf{\Sigma}_{xx}(f,n)^{-1}$ where  $\mathbf{\Sigma}_{xx}(f,n) = \mathbf{A}(f)\mathbf{\Sigma}_{ss}(f,n)\mathbf{A}(f)^{H}$  and  $\mathbf{\Sigma}_{sx}(f,n) = \mathbf{\Sigma}_{ss}(f,n)\mathbf{A}(f)^{H}$  ((.)<sup>H</sup> denotes the Hermitian conjugate)
- Remark:  $\mathbf{x}(f,n) = \mathbf{A}(f)\mathbf{y}(f,n)$  (exact reconstruction)
- Particular case: monophonic mixtures
  - Without loss of generality, we define A(f) = [1, ..., 1]

• We get 
$$y_k(f,n) = \frac{\sigma_k^2(f,n)}{\sum_{l=1}^K \sigma_l^2(f,n)} x(f,n)$$

$$\blacktriangleright$$
  $\Rightarrow$  similar to Wiener filtering

- 1. TF analysis of the mixtures:  $x_k(f,n) = (h_f * x_k)(f,nT)$
- 2. Estimation of  $\mathbf{A}(f)$  and  $\sigma_k^2(f, n)$ 
  - instantaneous mixture model
  - sparse source model
- 3. Computation of  $\mathbf{B}(f,n) = \mathbf{\Sigma}_{ss}(f,n)\mathbf{A}(f)^{H} (\mathbf{A}(f)\mathbf{\Sigma}_{ss}(f,n)\mathbf{A}(f)^{H})^{-1}$
- 4. Estimation of the source signals via  $\mathbf{y}(f,n) = \mathbf{B}(f,n)\mathbf{x}(f,n)$
- 5. TF synthesis of the sources:  $y_k(t) = \sum_{f=1}^F \sum_{n \in \mathbb{Z}} g_f(t nT) y_k(f, n)$

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	Stereophonic mixture	es: temporal sparsity		Sparsity in a transfor	rmed domain	

#### Case of a linear instantaneous stereophonic mixture: $\mathbf{x}(t) = \mathbf{A}\mathbf{s}(t)$



Case of a linear instantaneous stereophonic mixture:  $\mathbf{x}(f,n) = \mathbf{As}(f,n)$ 





1 P PARIS 44/48

04

Une école de l'IMT

43/48

Audio source separation

Une école de l'IMT

# **DUET** method



- Degenerate Unmixing Estimation Technique (DUET)
- Linear instantaneous stereophonic mixture model:  $\mathbf{x}(f,n) = \mathbf{As}(f,n)$ 
  - Without loss of generality, we assume  $\mathbf{A}_{(:,k)} = \begin{bmatrix} \cos(\theta_k) \\ \sin(\theta_k) \end{bmatrix} \forall k$
- **Sparse** source model:
  - ►  $\forall f, n, \exists ! k_{(f,n)}$  such that  $\sigma^2_{k_{(f,n)}}(f,n) > 0$ , and  $\forall l \neq k_{(f,n)}, \sigma^2_l(f,n) = 0$
- If only source k is active at (f, n), then  $\mathbf{x}(f, n) = \mathbf{a}_k s_k(f, n)$

**1**. TF analysis of the mixtures:  $x_k(f,n) = (h_f * x_k)(nT)$ 

2. Estimation of parameters  $\theta_k$  and of the active source  $k_{(f,n)}$ 

- computation of the histogram of the angles of vectors  $\mathbf{x}(f, n)$
- peak detection in order to estimate parameters  $\theta_k$
- determination of the active source at (f, n) by proximity with  $\theta_k$

3. Source separation: for all k,

- estimation of source images via binary masking:  $\mathbf{y}_k(f,n) = \mathbf{x}(f,n) \forall (f,n)$  such that  $k_{(f,n)} = k$  and  $\mathbf{y}_k(f,n) = 0$  for the other time-frequency bins (f,n)
- MMSE estimation of the sources:  $y_k(f,n) = \hat{\mathbf{a}}_k(f)^{\dagger} \mathbf{y}_k(f,n)$
- 4. TF synthesis of the sources:  $y_k(t) = \sum_{f=1}^F \sum_{n \in \mathbb{Z}} g_f(t nT) y_k(f, n)$



