

# Unprofiled Expectation Maximization Attack

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## A Talk About Side-Channel Analysis



Figure: An Electromagnetic Capture for SCA

# State of the Art

## ■ Supervised

## ■ Profiling Phase

## ■ Parametric Model

Properties	Supervised	Profiling Phase	Parametric Model	Flexible Model
Template [Chari et al.]	✓	✓	✓	✓
P-EM [Lemke et al. ]	—	✓	✓	✓
Proposed U-EM [This talk]	—	—	✓	✓
2O-CPA [Prouff et al.]	—	—	✓	—
MIA/KSA [Whitnall et al.]	—	—	—	—

## Notations and Setup

- An  $n$ -bit secret key  $k$  encrypts  $Q$  plaintext bytes  $\mathbf{t} = (t_1, t_2, \dots, t_Q)$
- The random masks  $\mathbf{M} = (M_1, M_2, \dots, M_Q)$  are i.i.d. uniform in  $\mathbb{F}_2^n$
- The sensitive variable  $\mathbf{X} = (X_1, X_2, \dots, X_Q)$  is such that

$$X_q = S(k \oplus t_q) \oplus M_q \quad (q = 1, 2, \dots, Q) \quad (1)$$

where  $S$  is the substitution box

- The noise is assumed Additive White Gaussian (AWGN)
- A bivariate leakage model  $\mathbf{Y} = (Y_1, \dots, Y_Q)$  where for each  $q \in \{1, 2, \dots, Q\}$  one has  $Y_q = (Y_q^{(1)}, Y_q^{(2)})$  with noise variance  $\sigma^2 = (\sigma_1^2, \sigma_2^2)$

## Leakage Models

### Hamming Weight Leakage Model.

$$\begin{cases} Y_q^{(1)} &= a^{*,(1)} w_H(X_q) + b^{*,(1)} + N_q^{(1)} \\ Y_q^{(2)} &= a^{*,(2)} w_H(M_q) + b^{*,(2)} + N_q^{(2)} \end{cases} \quad (q = 1, 2, \dots, Q) \quad (2)$$

where  $a^{*,(i)} \in \mathbb{R}$  and  $b^{*,(i)}$  are unknown parameters.

### Linear Leakage Model.

$$\begin{cases} Y_q^{(1)} &= \langle a^{*(1)}, X_q \rangle + b^{*(1)} + N_q^{(1)} \\ Y_q^{(2)} &= \langle a^{*(2)}, M_q \rangle + b^{*(2)} + N_q^{(2)} \end{cases} \quad (q = 1, 2, \dots, Q) \quad (3)$$

where  $\langle \cdot, \cdot \rangle$  denotes a bitwise scalar product over the reals and where vector  $a^{*(i)} \in \mathbb{R}^n$  and  $b^{*(i)} \in \mathbb{R}$  are unknown parameters.

## Generic Notation for the Sensitive Variable

To simplify the derivations we introduce the notations.

### ■ Hamming Weight Model

$$\begin{cases} x^{(1)}(a, b, k, t, m) = a \cdot w_H(m) + b \\ x^{(2)}(a, b, k, t, m) = a \cdot w_H(S(k \oplus t) \oplus m) + b \end{cases} \quad (4)$$

### ■ Linear Model

$$\begin{cases} x^{(1)}(a, b, k, t, m) = \langle a, m \rangle + b \\ x^{(2)}(a, b, k, t, m) = \langle a, S(k \oplus t) \oplus m \rangle + b \end{cases} \quad (5)$$

## Theoretical Optimal Distinguisher [Annelie et al.]

The Maximum Likelihood (ML)-based distinguisher is

$$\hat{k}(\mathbf{y}) = \arg \max_{k,a,b} \sum_{q=1}^Q \log \left[ \sum_{m_q \in \mathbb{F}_2^n} \exp\left(-\frac{1}{2} \|y_q - x(a, b, k, t_q, m_q)\|^2\right) \right]. \quad (6)$$

## 20-CPA [Prouff et al.]

For a given key hypothesis  $k$ , we write  $\mathbf{x}(k) = (X(k)_1, \dots, X(k)_Q)$  where

$$x(k)_q = \frac{1}{2^n} \sum_{m \in \mathbb{F}_2^n} w_H(m) w_H(S(k \oplus t_q) \oplus m) \quad (q = 1, 2, \dots, Q). \quad (7)$$

The distinguisher is then

$$\hat{k}_{20\text{-CPA}}(\mathbf{y}) = \arg \max_k |\rho(\mathbf{x}(k), \widetilde{\mathbf{y}}^{(1)} \widetilde{\mathbf{y}}^{(2)})| = \arg \max_k \left| \frac{\text{Cov}(\mathbf{x}(k), \widetilde{\mathbf{y}}^{(1)} \widetilde{\mathbf{y}}^{(2)})}{\sigma_{\mathbf{x}(k)} \sigma_{\mathbf{y}^{(1)} \mathbf{y}^{(2)}}} \right| \quad (8)$$

where  $\rho$  is the empirical Pearson correlation coefficient and  $\widetilde{y}$  denotes the centered version of the vector  $y$ .



## Expectation Maximization for SCA

$$a_{p+1}, b_{p+1} \leftarrow \underbrace{\arg \max_{a,b}}_{\text{M-Step}} \underbrace{\mathbb{E}_{\mathbf{M} \sim \mathcal{U}(\mathbb{F}_2^{\eta})} [\log(\mathbb{P}(\mathbf{Y} = \mathbf{y}, \mathbf{M} | k, a, b))]}_{\text{E-Step}} \quad (9)$$

## Bayes Posterior

For fixed  $k$  and  $q$ , let

$$\beta_q^{(p)}(m) = \mathbb{P}(M_q = m) \exp\left(-\frac{1}{2}\|y_q - x(a_p, b_p, k, t_q, m)\|^2\right).$$

Then the Bayes posterior of the mask  $m$  being used for the  $q$ -th traces given  $a_p, b_p$  is

$$\alpha_q^{(p)}(m) (= \mathbb{P}(M_q = m | y, a_p, b_p)) = \frac{\beta_q^{(p)}(m)}{\sum_{m'} \beta_q^{(p)}(m')}.$$

## Explicit E-Step

The E-Step at the  $p$ -th iteration is derived as follows.

$$\begin{aligned}\mathbb{E}[\log(\mathbb{P}(\mathbf{Y} = \mathbf{y}, \mathbf{M} | a, b))] &= \mathbb{E} \left[ \sum_q \log(\mathbb{P}(Y_q = y_q, M_q | a, b)) \right] \\ &= \sum_q \sum_{m_q} \mathbb{P}(M_q = m_q | Y_q = y_q, a_p, b_p) \log(\mathbb{P}(Y_q = y_q, M_q = m_q | a, b)) \\ &= \sum_{q,m} \alpha_q^{(p)}(m) \log(\mathbb{P}(Y_q = y_q | M_q = m, a, b)) + cst\end{aligned}\quad (10)$$

where the constant  $cst$  is independent of  $a$  and  $b$ . Thus the E-Step of the EM reduces to

$$(\mathbf{a}_{p+1}, \mathbf{b}_{p+1}) \leftarrow \arg \min_{(a,b)} \sum_q \sum_{m_q} \alpha_q^{(p)}(m_q) \|y_q - x(a, b, k, t_q, m_q)\|^2. \quad (11)$$

## M-Step for Hamming Weight Model

Then the empirical covariance and variance are

$$\widehat{\text{Cov}}_{\mathbf{xy}}^{(i)} = \frac{1}{Q} \sum_{q,m} \alpha_q(m) y_q^{(i)} x_{k,t_q,m}^{(i)}$$

$$\widehat{\text{Var}}_{\mathbf{x}}^{(i)} = \frac{1}{Q} \sum_{q,m} \alpha_q(m) y_q^{(i)2}$$

The M-Step is given by the following update rule ( $i = 1, 2$ ):

$$a^{(i)} = \frac{\widehat{\text{Cov}}_{\mathbf{xy}}^{(i)}}{\widehat{\text{Var}}_{\mathbf{x}}^{(i)}} \quad \text{and} \quad b^{(i)} = -a^{(i)} \bar{\mathbf{x}}^{(i)} \quad (12)$$

## M-Step for Linear Model

The empirical autocorrelation matrix and intercorrelation vector are

$$\widehat{R}_{\mathbf{xx}}^{(i)} = \sum_{q,m} \alpha_q(m) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)}) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)})^\perp \in \mathbb{R}^{n \times n}.$$

$$\widehat{R}_{\mathbf{xy}}^{(i)} = \sum_{q,m} \alpha_q(m) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)}) y_q^{(i)\perp} \in \mathbb{R}^n$$

The M-Step is given by the following update rule ( $i = 1, 2$ ):

$$\mathbf{a}^{(i)} = (\widehat{R}_{\mathbf{xx}}^{(i)})^{-1} \widehat{R}_{\mathbf{xy}}^{(i)} \quad \text{and} \quad b^{(i)} = -\langle \mathbf{a}^{(i)}, \bar{\mathbf{x}}^{(i)} \rangle \quad (13)$$

## Pseudo-Code of U-EM

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### Algorithm 2: Pseudo-code: U-EM-LIN.

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**Data:** The traces  $\mathbf{y} = (y_1, \dots, y_Q)$  and the noise standard deviation  $\sigma = (\sigma^{(1)}, \sigma^{(2)})$

**Input:** Convergence threshold  $\epsilon$

**Output:** Estimated key  $\hat{k}$

```
1  $\bar{y} \leftarrow \frac{1}{Q} \sum_{q=1}^Q y_q$  ; // Precompute the mean of the traces
2  $y^{(i)} \leftarrow \frac{y^{(i)} - \bar{y}^{(i)}}{\sigma^{(i)}}$  for  $i = 1, 2$  ; // Center and normalise the traces by  $\sigma$ 
3 forall key hypothesis  $k \in \mathbb{F}_2^n$  do
   | /* Initializations of the parameters  $a$  and  $b$  */
   |  $a, b \leftarrow ((1, \dots, 1), (1, \dots, 1)), (0, 0)$  ; // Arbitrary, but could be chosen
   | while True do
```

\*/

```

6  forall  $q, m$  compute do
7     $x_{k, t_q, m} \leftarrow (S(k \oplus t_q) \oplus m, m)$ 
8  forall  $q$  do
9     $c_q \leftarrow \max_m -\frac{1}{2} \|y_q - \langle a, x_{k, t_q, m} \rangle - b\|^2$ 
10 forall  $q, m$  do
11    $\beta_q(m) \leftarrow p(m) \exp(c_q - \frac{1}{2} \|y_q - \langle a, x_{k, t_q, m} \rangle - b\|^2)$ 
12 forall  $q$  compute do
13   forall  $m$  compute do
14      $\alpha_q(m) \leftarrow \frac{\beta_q(m)}{\sum \beta_q(m')}$ 
15      $\tilde{x}_q \leftarrow \sum_m \alpha_q(m) x_{k, t_q, m}$ 
16  $\bar{x} \leftarrow \frac{1}{Q} \sum_{q=1}^Q \tilde{x}_q$ 

```

```

17   /* M-Step */
18   for  $i = 1, 2$  do
19      $\widehat{R_{xx}}^{(i)} \leftarrow \sum_{q,m} \alpha_q(m) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)}) (x_{k,t_q,m}^{(i)} - \bar{\mathbf{x}}^{(i)})^\perp \in \mathbb{R}^{n \times n}$ 
20      $\widehat{R_{xy}}^{(i)} \leftarrow \sum_q \alpha_q(m) (\tilde{x}_q^{(i)} - \bar{\mathbf{x}}^{(i)}) y_q^{(i)\perp} \in \mathbb{R}^n$ 
21      $a'^{(i)} \leftarrow \widehat{R_{xx}}^{(i)-1} \widehat{R_{xy}}^{(i)}$ ,  $b'^{(i)} \leftarrow -\langle a'^{(i)}, \bar{\mathbf{x}}^{(i)} \rangle$ 
22     if  $(\|a - a'\|^2 + \|b - b'\|^2) < \epsilon$  then // Exit condition
23       Break ;
24      $a, b \leftarrow a', b'$ 

```

$$\text{LogLikelihood}(k) \leftarrow \sum_{q=1}^Q \log(\sum \beta_q) + c_q$$

**Result:**  $\hat{k} = \arg \max_k \text{LogLikelihood}(k)$





## Profiled EM [ Lemke et al.]

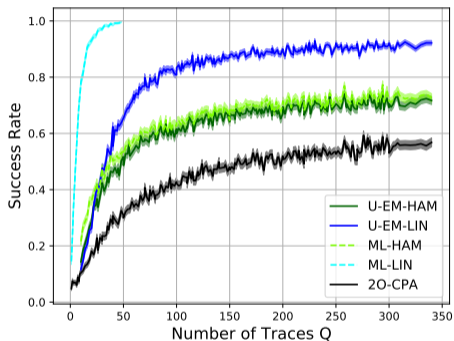
- Use EM to profile a template for each key Hypothesis.
- Require to have the necessary traces for profiling for each key hypothesis.
- Perform Maximum Likelihood with the derived templates.

## Numerical Evaluation Framework and Epistemic Noise

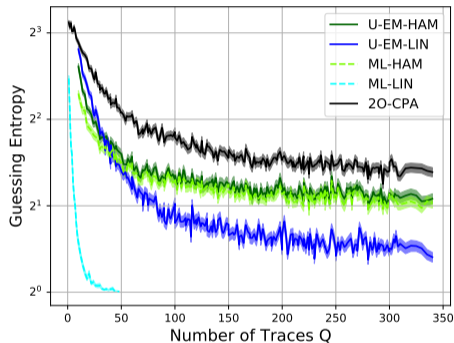
- 4 bits PRESENT substitution box.
- Noise with standard deviation  $\sigma$
- *Epistemic Noise* with standard deviation  $\sigma_a$
- $10^3$  independant repetitions. All attacks are performed on the same traces.
- Real Data: DPA Contest V4.2 with 8 bits Sub-Byte look up table

The epistemic noise is AWGN on the leakage coefficients and represents the physical peculiarities of a leaking device.

## Numerical Evaluations



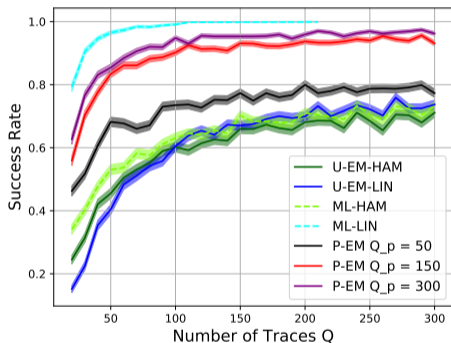
(a) Success rate.



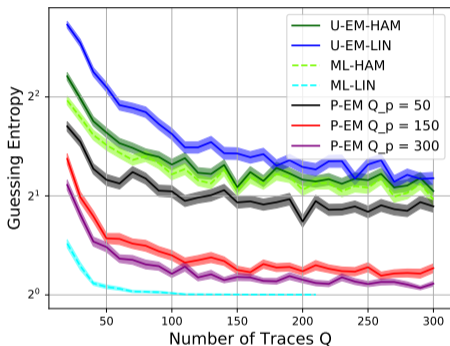
(b) Guessing entropy.

Figure: Attack metrics evaluated with  $\sigma = 0.3$  and  $\sigma_a = 0.8$ .

## Comparison with P-EM



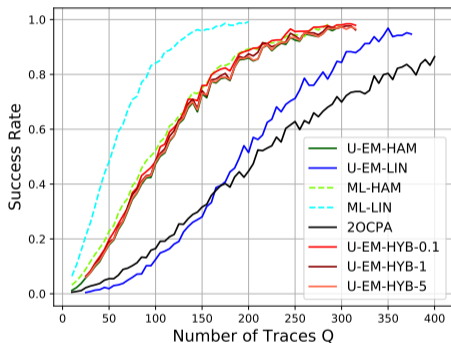
(a) Success rate on the DPA Contest.



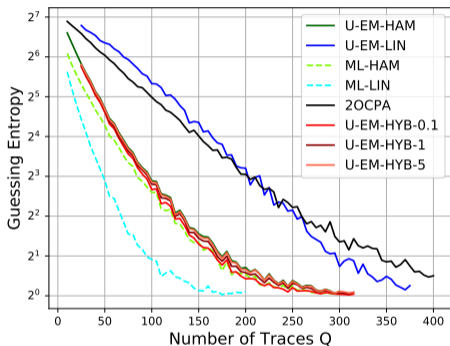
(b) Guessing entropy on the DPA Contest.

Figure: Attack metrics evaluated with  $\sigma = 0.4$  and  $\sigma_a = 0.4$ .

## Results on the Real Data from DPA Contest



(a) Success rate on the DPA Contest.



(b) Guessing entropy on the DPA Contest.

Figure: Attack metrics on the traces of the DPA Contest.



## Conclusion and Perspectives

- U-EM is interesting when profiling is not feasible
- U-EM is computationally efficient
- U-EM can be easily adapted to higher order attacks and is flexible
- When profiling is possible template or P-EM are better than our U-EM as it suffers from "overfitting"

## References

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# Thank You For Your Attention

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