# Information-Theoretic Tradeoffs in Distributed Hypothesis Testing

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Swiss Seminar on Coding and Crypto 24 November, 2021

## **Example: Distributed Control-System for Smart Cars**

- Smart cars measuring speed, distance, road conditions
- Fixed road-side sensors measuring same parameters
- Intact car system: measurements highly correlated
- Erroneous car system: measurements independent

## Task of Distributed Control-System

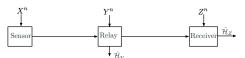
Decide on joint distribution underlying the observations

### Outline of the Talk

Simple single-sensor system

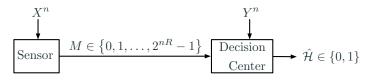


- · Infinitely many communication bits
- Single communication bit
- nR communication bits
- Multihop System



- Restriction only on expected communication rate
  - Simple single-sensor system
  - Multi-hop system

# **Distributed Hypothesis Testing**

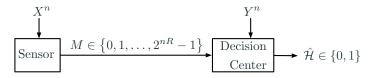


- "Normal situation"  $\mathcal{H} = 0$ :  $(X^n, Y^n) \sim \text{ i.i.d. } P_{XY}$
- "Hazardous event"  $\mathcal{H} = 1$ :  $(X^n, Y^n) \sim \text{ i.i.d. } Q_{XY}$
- Constraints on type-I and type-II error probabilities:

$$\begin{split} \overline{\lim}_{n \to \infty} \alpha_n &= \overline{\lim}_{n \to \infty} \mathbb{P} \big[ \hat{\mathcal{H}} = 1 | \mathcal{H} = 0 \big] \leq \epsilon \\ - \overline{\lim}_{n \to \infty} \frac{1}{n} \log \beta_n &= - \overline{\lim}_{n \to \infty} \frac{1}{n} \log \mathbb{P} \big[ \hat{\mathcal{H}} = 0 | \mathcal{H} = 1 \big] \geq \theta \end{split}$$

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# **Distributed Hypothesis Testing**



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## Rate-Exponent Tradeoff $\theta_{\epsilon}^*(R)$

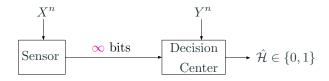
Given R > 0, largest exponent  $\theta$  that is  $\epsilon$ -achievable.

• Ahlswede & Csiszár'86:  $\theta_{\epsilon}^*(R)$  does not depend on  $\epsilon \in [0, 1/2]$ 

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# Centralized Hypothesis Testing — The Ideal Case

# **Centralized Hypothesis Testing**



- Rate R is so large that sensor can send all X<sup>n</sup> to decision center
- Optimal decision: raise alarm unless statistics tp  $(X^n, Y^n) \approx P_{XY}$
- Intuition: If system perfectly fits  $\mathcal{H}=0$ , decide on  $\hat{\mathcal{H}}=0$ , otherwise decide on  $\hat{\mathcal{H}}=1$

## **Example of Doubly-Symmetric Binary Sources**

$$P_{XY}(x,y) = \begin{cases} 1/3 & x = y \\ 1/6 & x \neq y \end{cases}$$
 and  $Q_{XY}(x,y) = \frac{1}{2} \cdot \frac{1}{2}$ 

Ex. 1: if observations

$$X^n = (0,1,1,0,1,0,1,0,1,1,0,1,0,0,1,1,0,0)$$
  
 $Y^n = (0,0,1,0,1,0,1,0,1,0,1,1,1,0,1,0,1,0)$ 

 $\operatorname{tp}\left(X^{n},Y^{n}
ight)=P_{XY}
ightarrow\operatorname{decision}$  center decides on  $\hat{\mathcal{H}}=0$ 

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Ex. 1: if observations

$$X^{n} = (0,1,1,0,1,0,1,0,1,1,0,1,0,0,1,1,0,0)$$

$$Y^{n} = (0,0,1,0,1,0,1,0,1,0,1,1,1,0,1,0,1,0)$$

$$\text{tp}(X^{n},Y^{n}) = P_{XY} \rightarrow \text{decision center decides on } \hat{\mathcal{H}} = 0$$

Ex. 2: if observations

$$\begin{array}{rcl} X^n & = & (0,1,{\color{red}0},1,1,1,{\color{red}0},1,0,1,1,0,1,0,0,1,1,0,{\color{red}0}) \\ Y^n & = & (1,0,{\color{red}0},1,0,1,{\color{red}0},0,0,1,0,1,1,1,1,1,0,1,{\color{red}0}) \\ \text{tp}\,(X^n,Y^n) \text{ far from } P_{XY} \rightarrow \text{decision center decides on } \hat{\mathcal{H}} = 1 \end{array}$$

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# **Analysis of Proposed Centralized Scheme**

• Type-I error probability (by the weak law of large numbers)

$$lpha_n = 1 - \mathbb{P} \big[ \hat{\mathcal{H}} = 0 | \mathcal{H} = 0 \big] = 1 - P_{XY}^{\otimes n} \Big( \operatorname{tp} (X^n, Y^n) \approx P_{XY} \Big)$$
  
 $\to 0 \quad \text{as} \quad n \to \infty$ 

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 $\to 0 \quad \text{as} \quad n \to \infty$ 

Probability of miss detection:

$$\beta_n = Q_{XY}^{\otimes n} \Big( \operatorname{tp}(X^n, Y^n) \approx P_{XY} \Big)$$

$$= \dots$$

• sequences  $\mathbf{x} = (x_1, \dots, x_n)$  and  $\mathbf{y} = (y_1, \dots, y_n)$  of type  $\pi_{\mathbf{x}\mathbf{y}}$ 

$$Q_{XY}^{\otimes n}(\mathbf{x},\mathbf{y}) = \prod_{i=1}^{n} Q_{XY}(x_i,y_i) = \prod_{(x,y) \in \mathcal{X} \times \mathcal{Y}} (Q_{XY}(x,y))^{n \cdot \pi_{\mathbf{x}\mathbf{y}}(x,y)}$$

$$Q_{XY}^{\otimes n}\Big(\operatorname{tp}\left(X^{n},\,Y^{n}
ight)=\pi\Big)pprox\expig(-n\cdot D(\pi||Q_{XY})ig)$$

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$$= \exp\left(n \cdot \sum_{(x, y) \in \mathcal{X} \times \mathcal{Y}} \pi_{\mathbf{x}\mathbf{y}}(x, y) \cdot \log Q_{XY}(x, y)\right)$$

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$$\begin{aligned} Q_{XY}^{\otimes n}(\mathbf{x}, \mathbf{y}) &= \prod_{i=1}^{n} Q_{XY}(x_{i}, y_{i}) = \prod_{(x, y) \in \mathcal{X} \times \mathcal{Y}} \left( Q_{XY}(x, y) \right)^{n \cdot \pi_{\mathbf{x}\mathbf{y}}(x, y)} \\ &= \exp \left( n \cdot \sum_{(x, y) \in \mathcal{X} \times \mathcal{Y}} \pi_{\mathbf{x}\mathbf{y}}(x, y) \cdot \log Q_{XY}(x, y) \right) \\ &= \exp \left( -n \cdot \sum_{(x, y) \in \mathcal{X} \times \mathcal{Y}} \pi_{\mathbf{x}\mathbf{y}}(x, y) \left( \log \frac{\pi_{\mathbf{x}\mathbf{y}}(x, y)}{Q_{XY}(x, y)} + \log \frac{1}{\pi_{\mathbf{x}\mathbf{y}}(x, y)} \right) \right) \end{aligned}$$

$$Q_{XY}^{\otimes n} \Big( \operatorname{tp} \left( X^n, Y^n 
ight) = \pi \Big) pprox \exp ig( - n \cdot D(\pi || Q_{XY}) ig)$$

• sequences  $\mathbf{x} = (x_1, \dots, x_n)$  and  $\mathbf{y} = (y_1, \dots, y_n)$  of type  $\pi_{\mathbf{x}\mathbf{y}}$ 

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$$Q_{XY}^{\otimes n}\Big(\operatorname{tp}\left(X^{n},\,Y^{n}
ight)=\pi\Big)pprox \expig(-n\cdot D(\pi||Q_{XY})ig)$$

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$$Q_{XY}^{\otimes n}(\mathbf{x}, \mathbf{y}) = \prod_{i=1}^{n} Q_{XY}(x_i, y_i) = \prod_{(x,y) \in \mathcal{X} \times \mathcal{Y}} (Q_{XY}(x,y))^{n \cdot \pi_{xy}(x,y)}$$

$$= \exp\left(n \cdot \sum_{(x,y) \in \mathcal{X} \times \mathcal{Y}} \pi_{xy}(x,y) \cdot \log Q_{XY}(x,y)\right)$$

$$= \exp\left(-n \cdot \sum_{(x,y) \in \mathcal{X} \times \mathcal{Y}} \pi_{xy}(x,y) \left(\log \frac{\pi_{xy}(x,y)}{Q_{XY}(x,y)} + \log \frac{1}{\pi_{xy}(x,y)}\right)\right)$$

$$= \exp(-n \cdot (D(\pi_{xy}||Q_{XY}) + H(\pi_{xy})))$$

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$$= \exp(-n \cdot (D(\pi_{\mathbf{x}\mathbf{y}}||Q_{XY}) + H(\pi_{\mathbf{x}\mathbf{y}})))$$

### Theorem (Sanov's theorem)

$$Q_{XY}^{\otimes n} \Big( \operatorname{tp} \left( X^n, \, Y^n 
ight) = \pi \Big) pprox \exp ig( - n \cdot D(\pi || \, Q_{XY}) ig)$$

• There are  $\approx 2^{nH(\pi)}$  sequences of a given type  $\pi$ 

## **Use Sanov's Theorem to Finalize the Analysis**

Type-I error probability (by the weak law of large numbers)

$$\alpha_n = 1 - P_{XY}^{\otimes n} \Big( \operatorname{tp}(X^n, Y^n) \approx P_{XY} \Big) \to 0 \quad \text{as} \quad n \to \infty$$

Type-II error probability

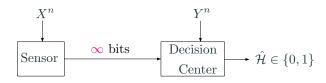
$$\beta_{n} = Q_{XY}^{\otimes n} \Big( \operatorname{tp}(X^{n}, Y^{n}) \approx P_{XY} \Big)$$

$$= \sum_{\pi \approx P_{XY}} Q_{XY}^{\otimes n} \Big( \operatorname{tp}(X^{n}, Y^{n}) = \pi \Big)$$

$$\approx \exp \Big( -n \cdot D(P_{XY} || Q_{XY}) \Big)$$

• N.B. The number of types  $\pi$  is polynomial in n

# **Optimal Exponent for Centralized Hypothesis Testing**



• Rate R is so large that sensor can send all X<sup>n</sup> to decision center

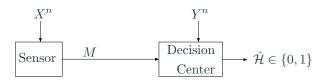
## Theorem (Stein's exponent)

Largest achievable error exponent is:

$$\theta^{\star}(R = \infty) = D(P_{XY} || Q_{XY}) := \sum_{x,y} P_{XY}(x,y) \log \frac{P_{XY}(x,y)}{Q_{XY}(x,y)}$$

# **Distributed Hypothesis Testing with Zero Communication Rate** R = 0

## <u>Distributed</u> Hypothesis Testing with R = 0 (Han'87)



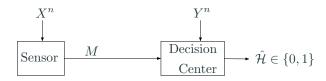
 Send message M, where number of bits representing M is sublinear in n

### Theorem (Han'87)

Largest achievable error exponent is:

$$\theta^{\star}(R=0) = \min_{\substack{\pi_{XY}:\\\pi_{X}=P_{X}\\\pi_{Y}=P_{Y}}} D(\pi_{XY}||Q_{XY})$$

# Scheme for R = 0 (Han'87)



- *M* ∈ {0,1} (1 bit) suffices
- Sensor sends its own local decision: If  $\operatorname{tp}(X^n) \approx P_X \to \operatorname{send} M = 0$ , otherwise send M = 1
- · Decision center:
  - 1. Local decision: L = 0 if  $tp(Y^n) \approx P_Y$  and L = 1 else
  - 2. Final decision:  $\hat{\mathcal{H}} = 0$  if M = L = 0, and  $\hat{\mathcal{H}} = 1$  else.

# **Example of Doubly-Symmetric Binary Sources**

$$P_{XY}(x,y) = \begin{cases} 1/6 & x = y \\ 1/3 & x \neq y \end{cases}$$
 and  $Q_{XY}(x,y) = \frac{1}{2} \cdot \frac{1}{2}$ 

Ex.: 
$$X^n = (0, 1, 0, 1, 1, 1, 0, 1, 0, 1, 1, 0, 1, 0, 0, 1, 1, 0, 0)$$
  
 $Y^n = (1, 0, 0, 1, 0, 1, 0, 0, 0, 1, 0, 1, 1, 1, 1, 1, 0, 1, 0)$ 

 $\operatorname{tp}(X^n) = P_X$  and  $\operatorname{tp}(Y^n) = P_Y \to \operatorname{decision}$  center decides  $\hat{\mathcal{H}} = 0$ 

For centralized setting  $(R = \infty)$ , decision center decides  $\hat{\mathcal{H}} = 1$ !

## **Analysis of 1-bit Communication Scheme**

Type-I error probability (by the weak law of large numbers)

$$lpha_n = 1 - P_{XY}^{\otimes n} \Big( \operatorname{tp}(X^n) pprox P_X \quad \text{and} \quad \operatorname{tp}(Y^n) pprox P_Y \Big) o 0 \quad \text{as} \quad n o \infty$$

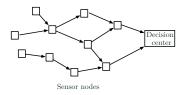
Type-II error probability:

$$\beta_{n} = Q_{XY}^{\otimes n} \Big( \operatorname{tp}(X^{n}) \approx P_{X} \text{ and } \operatorname{tp}(Y^{n}) \approx P_{Y} \Big)$$

$$= \sum_{\substack{\pi_{xy} : \pi_{x} \approx P_{X} \\ \pi_{y} \approx P_{Y}}} Q_{XY}^{\otimes n} \Big( \operatorname{tp}(X^{n}, Y^{n}) = \pi_{xy} \Big)$$

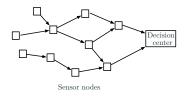
$$\approx \exp \Big( -n \cdot \Big( \min_{\substack{\pi_{xy} : \pi_{x} = P_{X} \\ \pi_{y} = P_{Y}}} D(\pi_{xy} || Q_{XY}) \Big) \Big)$$

# **Many-Sensors and One Detector**



- "Normal situation"  $\mathcal{H}=0$ :  $(X_1^n,\ldots,X_K^n,Y^n)\sim \text{ i.i.d. } P_{X_1\cdots X_KY}$
- "Hazardous event"  $\mathcal{H}=1$ :  $(X_1^n,\dots,X_K^n,Y^n)\sim \text{ i.i.d. }Q_{X_1\cdots X_KY}$

# **Many-Sensors and One Detector**



- Each sensor k produces local decision
- Unanimous decision forwarding: sensor sends  $M_k = 0$  only if all incoming messages 0 and local decision  $\mathcal{H} = 0$
- Decision center: raises alarm <u>unless</u> all incoming messages 0 and local decision  $\mathcal{H}=0$
- Optimal exponent (independent of  $\epsilon$ )

$$\theta_{\epsilon}^* = \min_{\substack{\pi_{X_1 \cdots X_K Y}: \\ \pi_{X_k} = P_{X_k}, \ \forall k \\ \pi_Y = P_Y}} D(\pi_{X_1 \cdots X_K Y} || Q_{X_1 \cdots X_K Y})$$

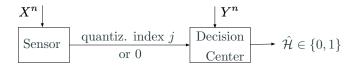
## When is a single-bit transmission optimal for R = 0?

- · Single-bit transmission (unanimous decision forward) optimal if
  - Two hypotheses  $\mathcal{H}=0$  and  $\mathcal{H}=1$
  - All decision centers interested in same exponent
  - Interactive communication allowed [Katz-Piantanida-Debbah-2016]

- · Sending a single bit is not sufficient if
  - · Exponents under both decisions need to be maximized
  - Different decision centers interested in different exponents
  - More than K > 2 hypotheses

# **Distributed Hypothesis Testing under Positive Rates** R > 0

# <u>Distributed</u> Hypothesis Testing with R > 0 (Han'87)



- Random codebook  $C_S = \{S^n(1), \dots, S^n(2^{nR} 1)\}$
- Quantization: Send M = j if  $\operatorname{tp}(S^n(j), X^n) \approx P_{SX}$ ; else M = 0
- Alarm  $\hat{\mathcal{H}} = 1$  unless  $M \ge 1$  and  $\operatorname{tp}(S^n(M), Y^n) \approx P_{SY}$
- If  $R \ge I(S; X)$  quantization succeeds and  $\alpha_n \to 0$  as  $n \to \infty$
- Achievable type-II error exponent

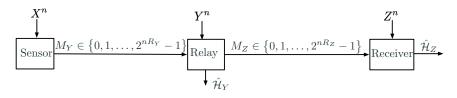
$$\theta_{\epsilon}^*(R) \geq \max_{\substack{P_{S|X}:\\R \geq I(S;X)}} \quad \min_{\substack{\pi_{SXY}:\\\pi_{SY} = P_{SY}\\\pi_{SY} = P_{SY}}} D(\pi_{SXY}||P_{S|X}Q_{XY})$$

- $\mathcal{H} = 0$ :  $(X^n, Y^n) \sim \text{i.i.d. } P_{XY}$
- $\mathcal{H} = 1$ :  $(X^n, Y^n) \sim \text{i.i.d. } P_X P_Y$

## Optimal Rate-Exponent Tradeoff (does not depend on $\epsilon$ )

$$\theta_{\epsilon}^{*}(R) = \max_{\substack{P_{S|X}: \\ R \geq I(S;X)}} I(S; Y) =: \eta_{XY}(R)$$

# **Testing Against Independence over Two Hops**



- Two decision centers (relay and receiver)
- Markov chain  $X \to Y \to Z$  under both hypotheses

• 
$$\mathcal{H} = 0 : (X^n, Y^n, Z^n) \text{ i.i.d. } \sim P_X \cdot P_{Y|X} \cdot P_{Z|Y}$$

• 
$$\mathcal{H} = 1 : (X^n, Y^n, Z^n) \text{ i.i.d. } \sim P_X \cdot P_Y \cdot P_Z$$

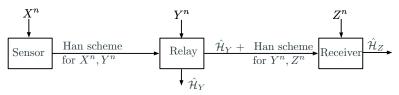
# **Definition of Exponents Region**

- Error Constraints at the Relay
  - $\alpha_{Y,n} \triangleq \mathbb{P}[\hat{\mathcal{H}}_Y = 1 | \mathcal{H} = 0] \leq \epsilon_Y$
  - $\beta_{Y,n} \triangleq \mathbb{P}[\hat{\mathcal{H}}_Y = 0 | \mathcal{H} = 1] \leq 2^{-n\theta_Y}$
- · Error Constraints at the Receiver
  - $\alpha_{Z,n} \triangleq \mathbb{P}[\hat{\mathcal{H}}_Z = 1 | \mathcal{H} = 0] \leq \epsilon_Z$
  - $\beta_{Z,n} \triangleq \mathbb{P}[\hat{\mathcal{H}}_Z = 0 | \mathcal{H} = 1] \leq 2^{-n\theta_Z}$

#### **Definition**

 $\mathcal{E}^*_{\epsilon_Y,\epsilon_Z}(R_Y,R_Z)$  is the closure of the set of all  $(\epsilon_Y,\epsilon_Z)$ -achievable exponent pairs  $(\theta_Y,\theta_Z)$ 

# **Optimal Scheme and Exponents Region**



- Independent Han-scheme over each link
- Unanimous Decision-Forwarding
- $\hat{\mathcal{H}}_Z=0$  only if  $\hat{\mathcal{H}}_Y=0$  and Han scheme for  $(Y^n,Z^n)$  indicates  $\mathcal{H}=0$

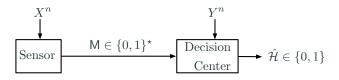
## Theorem (Salehkalaibar, W', Wang 2017, Cao, Zhu, Tan'2019)

$$\mathcal{E}_{\epsilon_Y,\epsilon_Z}(R_Y,R_Z)$$
 is the set of all pairs  $(\theta_Y,\theta_Z)$  satisfying

$$\theta_{Y} \leq \eta_{XY}(R_{Y}), \qquad \quad \theta_{Z} \leq \eta_{XY}(R_{Y}) + \eta_{YZ}(R_{Z}).$$

# **Testing under Expected Rate Constraints**

# Point-to-Point Testing under **Expected Rate** Constraints



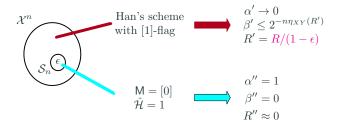
- Message M is a variable-length bit-string
- Rate constraint  $\mathbb{E}[\operatorname{len}(M)] \leq nR$

#### Theorem (Salehkalaibar and Wigger'2020)

The largest  $\epsilon$ -achievable exponent is

$$\theta_{\text{VL},\epsilon}^*(R) = \eta_{XY}(R/(1-\epsilon))$$
 (depends on  $\epsilon$ )

# **Optimal Point-to-Point Variable-Length Scheme**

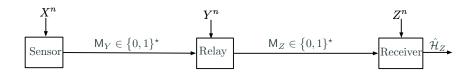


- Set  $S_n$  has probability  $\epsilon$
- Average performances

$$\alpha_n \le \epsilon$$
 
$$\beta_n \le (1-\epsilon)2^{-n\cdot\eta_{XY}(R/(1-\epsilon))}$$
 and total rate is  $(1-\epsilon)R'=R$ 

$$\rightarrow$$
 Achievability of  $\theta = \eta_{XY}(R/(1-\epsilon))$ 

### Two-Hop Testing under **Expected Rate** Constraints

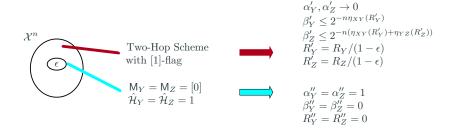


- Messages M<sub>Y</sub> and M<sub>Z</sub> are variable-length bit-strings
- Rate constraints  $\mathbb{E}[\operatorname{len}(M_Y)] \leq nR_Y$  and  $\mathbb{E}[\operatorname{len}(M_Z)] \leq nR_Z$

#### **Definition**

Exponents region  $\mathcal{E}_{VL,\epsilon_Y,\epsilon_Z}(R_Y,R_Z)$  is the set of all  $(\epsilon_Y,\epsilon_Z)$ -achievable exponent pairs  $(\theta_Y,\theta_Z)$  under expected rate constraints

# Achievability with Expected Rate Constraints for $\epsilon_Y = \epsilon_Z \triangleq \epsilon$

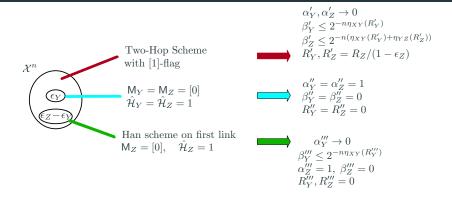


#### Theorem (Hamad, Sarkiss, W'2021)

 $\mathcal{E}_{\epsilon,\epsilon}(R_Y,R_Z)$  is the set of all pairs  $(\theta_Y,\theta_Z)$  satisfying

$$\theta_Y \leq \eta_{XY}(R_Y/(1-\epsilon)), \qquad \quad \theta_Z \leq \eta_{XY}(R_Y/(1-\epsilon)) + \eta_{YZ}(R_Z/(1-\epsilon)).$$

### Scheme for Expected Rate Constraints with $\epsilon_Y < \epsilon_Z$



- Average type-I error prob.  $\alpha_Y \to \epsilon_Y$  and  $\alpha_Z \to \epsilon_Z$  as  $n \to \infty$
- Average type-II error prob. at Relay:  $\beta_Y \leq (1 \epsilon_Z) 2^{-n\eta_{XY}(R_Y')} + (\epsilon_Z \epsilon_Y) 2^{-n\eta_{XY}(R_Y'')})$
- Average rate of first link  $(1 \epsilon_Z)R_Y' + (\epsilon_Z \epsilon_Y)R_Y'''$

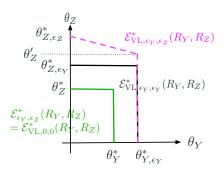
# Optimal Exponents Region for $\epsilon_Y < \epsilon_Z$

#### Theorem (Hamad, W', Sarkiss '2021)

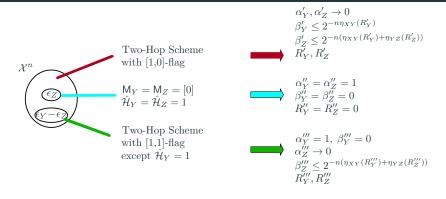
 $\mathcal{E}^*_{VL,\epsilon_Y,\epsilon_Z}(R_Y,R_Z)$  set of  $(\theta_Y,\theta_Z)$  pairs satisfying

$$\begin{aligned} \theta_Y &\leq \min\{\eta_{XY}(R_Y'), \eta_{XY}(R_Y''')\} \\ \theta_Z &\leq \eta_{XY}(R_Y') + \eta_{YZ}(R_Z/(1 - \epsilon_Z)), \end{aligned}$$

for some  $R_{Y'}, R_Y''' > 0$  satisfying  $R_Y \ge (1 - \epsilon_Z)R_Y' + (\epsilon_Z - \epsilon_Y)R_Y'''$ .



### Scheme for Expected Rate Constraints with $\epsilon_Y > \epsilon_Z$



- Average type-I error prob.  $\alpha_Y \to \epsilon_Y$  and  $\alpha_Z \to \epsilon_Z$  as  $n \to \infty$
- Average type-II error prob. at Receiver:  $\beta_Y \leq (1 \epsilon_Y) 2^{-n(\eta_{XY}(R'_Y) + \eta_{YZ}(R'_Z))} + (\epsilon_Y \epsilon_Z) 2^{-n(\eta_{XY}(R'''_Y) + \eta_{YZ}(R'''_Z))}$
- Average rate of first link  $(1 \epsilon_Y)R'_Y + (\epsilon_Y \epsilon_Z)R'''_Y$

# Optimal Exponents Region for $\epsilon_Y > \epsilon_Z$

#### Theorem (Hamad, W', Sarkiss '2021)

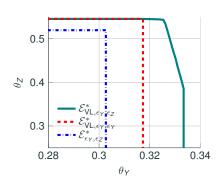
$$\mathcal{E}^*_{\mathsf{VL},\epsilon_Y,\epsilon_Z}(R_Y,R_Z)$$
 set of  $(\theta_Y,\theta_Z)$  pairs satisfying

$$\theta_{Y} \leq \eta_{XY}(R'_{Y})$$
  
$$\theta_{Z} \leq \min\{\eta_{XY}(R'_{Y}) + \eta_{YZ}(R'_{Z}), \eta_{XY}(R'''_{Y}) + \eta_{YZ}(R'''_{Z})\}$$

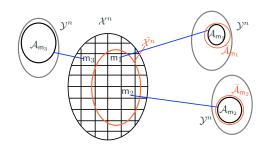
for some 
$$R_{Y'}$$
,  $R_{Y'}^{"}$ ,  $R_{Z}^{"}$ ,  $R_{Z}^{""} > 0$  satisfying  $R_{Y} \geq (1 - \epsilon_{Z})R_{Y}^{\prime} + (\epsilon_{Z} - \epsilon_{Y})R_{Y}^{""}$  and  $R_{Z} \geq (1 - \epsilon_{Z})R_{Z}^{\prime} + (\epsilon_{Z} - \epsilon_{Y})R_{Z}^{""}$ 

### Two-Hop Setup with $\epsilon_Y > \epsilon_Z$ : Numerical Results

- Binary example
  - X ~ Bernoulli(0.5)
  - $Y = X \oplus Bernoulli(0.9)$
  - *Z* = *Y* ⊕ Bernoulli(0.8)
- Type-I error probabilities  $\epsilon_Y = 0.1$  and  $\epsilon_Z = 0.05$
- Rates  $R_Y = R_Z = 0.5$



### Idea of Converse Proof for Single-Sensor System



Original acceptance region:  $A = \bigcup_{m} \{m\} \times A_m$ 

New problem:

 $\tilde{X}^n$  lives on  $\tilde{\mathcal{X}^n}$ 

New acceptance region:

$$\bar{\mathcal{A}} = \bigcup_{m} \{m\} \times \bar{\mathcal{A}}_{m}$$

- Restrict to subset  $\tilde{\mathcal{X}^n} \to \text{change of measure}$
- ullet Slightly-enlarge the acceptance region o blowing-up lemma

New problem has 
$$\tilde{\alpha}_n \approx 0$$
 and  $\mathbb{E}[\operatorname{len}(\tilde{\mathbb{M}})] \leq \frac{\mathbb{E}[\operatorname{len}(\mathbb{M})]}{1-\epsilon}$  and  $-\frac{1}{n}\log\tilde{\beta}_n \approx -\frac{1}{n}\log\beta_n$ 

### **Converse Proof: Bound on Rate**

- Pick set  $\mathcal{D}_n := \left\{ x^n \in \mathcal{T}_{\epsilon}^{(n)}(P_X) \colon \mathbb{P}[\hat{\mathcal{H}} = 0 | \mathcal{H} = 0, X^n = x^n] \ge \eta \right\}$  for small parameters  $\eta, \epsilon > 0$  that will tend to 0
- Since  $\mathbb{P}[\hat{\mathcal{H}} = 0 | \mathcal{H} = 0, X^n = x^n] \ge 1 \epsilon \rightarrow \mathbb{P}[\mathcal{D}_n] \ge \frac{1 \epsilon \eta}{1 \eta}$
- Change of measure:

$$(\tilde{X}^n, \tilde{Y}^n, \tilde{\mathsf{M}}) \sim \prod_{i=1}^n P_X(x_i) \frac{\mathbb{1}\{x^n \in \mathcal{D}_n\}}{\mathbb{P}[\mathcal{D}_n]} \prod_{i=1}^n P_{Y|X}(y_i|x_i) \mathbb{1}\{m = \mathsf{enc}(x^n)\}$$

•  $R \ge \frac{1}{n} \mathbb{E}[\operatorname{len}(M)] \ge \frac{1}{n} \mathbb{P}[\mathcal{D}_n] \cdot \mathbb{E}[\operatorname{len}(\tilde{M})]$ 

$$\begin{split} &\mathbb{E}[\operatorname{len}(\tilde{\mathsf{M}})] \geq H(\tilde{\mathsf{M}}|\operatorname{len}(\tilde{\mathsf{M}})) = (H(\tilde{\mathsf{M}}) - H(\operatorname{len}(\tilde{\mathsf{M}}))) \approx H(\tilde{\mathsf{M}}) \\ &\geq I(\tilde{\mathsf{M}}; \tilde{X}^n) = \sum_{i=1}^n I(\tilde{\mathsf{M}}; \tilde{X}_i | \tilde{X}^{i-1}) \\ &= \sum_{i=1}^n I(\tilde{\mathsf{M}}, \tilde{X}^{i-1}; \tilde{X}_i) + \log(\mathbb{P}[\mathcal{D}_n]) = \sum_{i=1}^n I(U_i; \tilde{X}_i) + \log(\mathbb{P}[\mathcal{D}_n]) \end{split}$$

### **Converse Proof: Marton's Blowing Up Lemma**

### Lemma (Marton's blowing up lemma)

Let  $S_1, S_2, \ldots$  be i.i.d.  $\sim P_S$  and  $\{\epsilon_n\} \downarrow 0$ . There exist sequences  $\{\ell_n\}$  and  $\{\zeta_n\}$  both  $\downarrow 0$  s.t. for any set  $\mathcal{B}_n$ : If  $P_S^{\otimes n}(\mathcal{B}_n) \geq \exp(-n\epsilon_n)$ , then  $P_S^{\otimes n}(\bar{\mathcal{B}}_n^{(\ell_n)}) \geq 1 - \zeta_n$ .

- Blow-up acceptance regions
   Ā<sub>m</sub> := {y<sup>n</sup>: ∃ỹ<sup>n</sup> s.t. d<sub>H</sub>(y<sup>n</sup>, ỹ<sup>n</sup>) ≤ ℓ<sub>n</sub>, ỹ<sup>n</sup> ∈ A<sub>m</sub>}
- By the blowing-up lemma and the definition of  $\mathcal{D}_n$

$$\mathbb{P}[(\tilde{\mathsf{M}}, \tilde{\mathsf{Y}}^n) \in \bar{\mathcal{A}} | \mathcal{H} = 0] \geq 1 - \zeta_n$$

Since the blowup is very small

$$-\frac{1}{n}\log \mathbb{P}[(\tilde{M}, \tilde{Y}^n) \in \bar{\mathcal{A}}\big|\mathcal{H} = \mathbf{1}] \approx -\frac{1}{n}\log \mathbb{P}[(\tilde{M}, \tilde{Y}^n) \in \mathcal{A}\big|\mathcal{H} = \mathbf{1}]$$
$$\approx -\frac{1}{n}\log \mathbb{P}[(M, Y^n) \in \mathcal{A}\big|\mathcal{H} = \mathbf{1}]$$

# **Converse Proof: Bound on Exponent**

$$\begin{split} &-\frac{1}{n}\log\mathbb{P}[(\tilde{\mathbf{M}},\tilde{Y}^n)\in\bar{\mathcal{A}}\big|\mathcal{H}=\mathbf{1}]\approx\frac{1}{1-\zeta_n}D\big(P_{(\tilde{\mathbf{M}},\tilde{Y}^n)|\mathcal{H}=\mathbf{0}}\|P_{(\tilde{\mathbf{M}},\tilde{Y}^n)|\mathcal{H}=\mathbf{1}}\big)\\ &\approx &\frac{1}{1-\zeta_n}I(\tilde{\mathbf{M}};\,\tilde{Y}^n)=\frac{1}{1-\zeta_n}\sum_{i=1}^nI(\tilde{\mathbf{M}};\,\tilde{Y}_i|\,\tilde{Y}^{i-1})\\ &\approx &\frac{1}{1-\zeta_n}\sum_{i=1}^nI(\tilde{\mathbf{M}},\,\tilde{Y}^{i-1};\,\tilde{Y}_i)\\ &\geq &\frac{1}{1-\zeta_n}\sum_{i=1}^nI(\tilde{\mathbf{M}},\,|\tilde{X}^{i-1};\,\tilde{Y}_i)=\frac{1}{1-\zeta_n}\sum_{i=1}^nI(U_i;\,\tilde{Y}_i) \end{split}$$

- We related the exponent to the rate via the same variables *U<sub>i</sub>*
- Remaining steps by introducing a uniform random variable T for the time index and taking  $n \to \infty$ ,  $\epsilon, \eta \downarrow 0$

### **Summary**

- Distributed hypothesis testing under zero-rate constraints →
   Local type-based decisions and unanimous decision-forwarding
- Distributed hypothesis testing under maximum rate-constraints
   → Quantization and unanimous-decision forwarding
- Distributed hypothesis testing under expected rate-constraints:
  - Combine different degenerate versions of optimal fixed-length schemes
  - Rate-boost on all rates
  - Tradeoff between different decisions