

Control Variates

Gersende FORT, Eric MOULINES

LTCI, CNRS / TELECOM ParisTech

Première partie I

Introduction

Principles

The principle of **control variates** consists in using the estimation error of **known quantities**, to improve the estimation of **unknown quantities**.

Assumptions

- X and Y are positively correlated.
- $\mathbb{E}[X]$ is **known** and $\{X_k, k \geq 0\}$ are i.i.d. with the same distribution than X .
- $\mathbb{E}[Y]$ is **unknown** and $\{Y_k, k \geq 0\}$ are i.i.d. with the same distribution than Y .

Principles

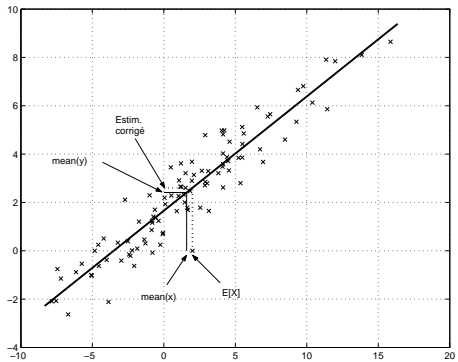
- The plain Monte Carlo estimators of $\mathbb{E}[Y]$ (resp. $\mathbb{E}[X]$) are given by

$$\bar{\mu}_n \stackrel{\text{def}}{=} \frac{1}{n} \sum_{k=1}^n Y_k \qquad \frac{1}{n} \sum_{k=1}^n X_k$$

- The estimation error for $\mathbb{E}[X]$ is

$$\mathcal{E}_n \stackrel{\text{def}}{=} \mathbb{E}[X] - \frac{1}{n} \sum_{k=1}^n X_k$$

- **Idea** : correct $\hat{\mu}_n$ knowing \mathcal{E}_n .



If $\mathcal{E}_n > 0$,

- $n^{-1} \sum_{k=1}^n X_k$ underestimate $\mathbb{E}[X]$.
- Since X and Y are **positively** correlated, it is likely that $\bar{\mu}_n$ also underestimate $\mathbb{E}[Y]$
- **idea** Correct the estimator by **adding a positive** quantity

$$\bar{\mu}_n + b \left(\mathbb{E}[X] - \frac{1}{n} \sum_{k=1}^n X_k \right)$$

Control Variates estimators

- Estimator

$$\hat{\mu}_n(b) \stackrel{\text{def}}{=} \frac{1}{n} \sum_{k=1}^n Y_k - b \left(\frac{1}{n} \sum_{k=1}^n X_k - \mathbb{E}[X] \right)$$

- Questions :

- 1 how to adjust b ?
 - 2 What are the properties of this estimator? (how much do I gain on the variance)
- More general schemes (involving non linear corrections) may be considered as well.

Deuxième partie II

Control Variates : Methodology

Assumptions

- 1 $\{(X_k, Y_k), k \geq 0\}$ are i.i.d.
- 2 $\mathbb{E}[X]$ is known and $\text{Var}(X) < +\infty$.
- 3 $\text{Var}(Y) < +\infty$.

Beware

- $\{X_k, k \geq 0\}$ are i.i.d.
- $\{Y_k, k \geq 0\}$ are i.i.d.
- **But** X_k and Y_k are **dependent** !

Definition

- Plain MC estimator

$$\bar{\mu}_n \stackrel{\text{def}}{=} \frac{1}{n} \sum_{k=1}^n Y_k$$

- Control Variate estimator

$$\hat{\mu}_n(b) \stackrel{\text{def}}{=} \bar{\mu}_n - b \left(\frac{1}{n} \sum_{k=1}^n X_k - \mathbb{E}[X] \right) = \bar{\mu}_n - b (\bar{X}_n - \mathbb{E}[X])$$

where

$$\bar{X}_n \stackrel{\text{def}}{=} \frac{1}{n} \sum_{k=1}^n X_k$$

Some easy properties

The estimator $\hat{\mu}_n(b)$ is

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- unbiased (to show)
- (strongly) consistent (to show)

Variance

- Variance (to show)

$$\text{Var}(\hat{\mu}_n(b)) = \text{Var}(\bar{\mu}_n) + \frac{1}{n} \left(\text{Var}(bX) - 2\text{Cov}(bX, Y) \right)$$

- An easy upper bound

$$\text{Var}(\hat{\mu}_n(b)) \leq \text{Var}(\bar{\mu}_n) \iff \text{Var}(bX) < 2\text{Cov}(bX, Y).$$

- Idea : Choose b to minimize the variance $\text{Var}(\hat{\mu}_n(b))$.

Optimal b

- Optimal b (to show)

$$b_{\star} \stackrel{\text{def}}{=} \frac{\text{Cov}(X, Y)}{\text{Var}(X)}.$$

- Optimal Variance (to show)

$$\text{Var}(\hat{\mu}_n(b_{\star})) = \text{Var}(\bar{\mu}_n) (1 - \rho_{X,Y}^2)$$

where

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}} = \text{Corr}(X, Y)$$

Limiting distribution

- Central Limit Theorem

$$\begin{aligned}\hat{\mu}_n(b) &\stackrel{\text{def}}{=} \bar{\mu}_n - b \left(\frac{1}{n} \sum_{k=1}^n X_k - \mathbb{E}[X] \right) \\ &= \mathbb{E}[Y] + \frac{1}{n} \sum_{k=1}^n (\{Y_k - \mathbb{E}[Y]\} - b \{X_k - \mathbb{E}[X]\})\end{aligned}$$

- Confidence bound

$$\frac{\sqrt{n}}{\sigma(b)} \left(\hat{\mu}_n(b) - \mathbb{E}[Y] \right) \xrightarrow{\mathcal{L}} \mathcal{N}(0, 1)$$

where

$$\sigma^2(b) \stackrel{\text{def}}{=} \text{Var}(Y - bX) = \text{Var}(Y) + b^2\text{Var}(X) - 2b\text{Cov}(X, Y).$$

Confidence Bounds

- The limiting distribution can be used to construct asymptotic confidence intervals with coverage $(1 - \delta)\%$ of $\mathbb{E}[Y]$

$$\left[\hat{\mu}_n(b) - \frac{\sigma(b)}{\sqrt{n}} z_{1-\delta/2}; \hat{\mu}_n(b) + \frac{\sigma(b)}{\sqrt{n}} z_{1-\delta/2} \right]$$

where z_α is the α -th quantile of the standard normal distribution.

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Wrap up

- The variance reduction increases with the correlation.
- The variance reduction vanishes whenever X and Y are independent.
- When $\rho_{X,Y}^2 = 1$, the variance is zero!!
- **Problem** make the dream comes true!

How to compute b_\star

$$b_\star \stackrel{\text{def}}{=} \frac{\text{Cov}(X, Y)}{\text{Var}(X)}.$$

- **Idea** Use the samples $\{(X_k, Y_k), k \leq n\}$ to approximate b_\star by \hat{b}_n .
- **Warning** : check that (asymptotically) the properties of $\hat{\mu}_n(\hat{b}_n)$ are the same than $\hat{\mu}_n(b_\star)$?
- **Alternative** (but costly) Use an independent sample $\{(X'_k, Y'_k), k \leq n_1\}$ to approximate b_\star .

Limiting distribution

- (show)

$$\frac{\sqrt{n}}{\sigma(b_\star)} \left(\hat{\mu}_n(\hat{b}_n) - \mathbb{E}[Y] \right) \xrightarrow{\mathcal{L}} \mathcal{N}(0, 1)$$

where

$$\hat{b}_n \stackrel{\text{def}}{=} \frac{\sum_{k=1}^n (X_k - \mathbb{E}[X])(Y_k - \bar{\mu}_n)}{\sum_{k=1}^n (X_k - \mathbb{E}[X])^2} \qquad b_\star \stackrel{\text{def}}{=} \frac{\text{Cov}(X, Y)}{\text{Var}(X)}$$

- Therefore $\hat{\mu}_n(\hat{b}_n)$ as the same asymptotic distribution $\hat{\mu}_n(b_\star)$.
- The asymptotic distribution is unchanged if $\sigma^2(b_\star)$ by a consistent estimator of this quantity.

Multidimensional extensions

- Several control values $X = (X^{(1)}, \dots, X^{(d)})^T$ can be available
- In such case, the control variate estimate is

$$\hat{\mu}_n(b) \stackrel{\text{def}}{=} \bar{\mu}_n - b^T (\bar{X}_n - \mathbb{E}[X])$$

where

$$\bar{X}_n \stackrel{\text{def}}{=} \frac{1}{n} \sum_{k=1}^n X_k \quad X_k \text{ i.i.d. } \sim X$$

- Along the same lines the **optimal weight** is $b_\star \stackrel{\text{def}}{=} \Sigma_X^{-1} \Sigma_{X,Y}$ where

$$\text{Var}(\hat{\mu}_n(b_\star)) = \text{Var}(\bar{\mu}_n) (1 - R^2) \quad R^2 \stackrel{\text{def}}{=} \frac{\Sigma_{X,Y}^T \Sigma_X^{-1} \Sigma_{X,Y}}{\text{Var}(Y)}$$

where Σ_X is the covariance matrix of X and $\Sigma_{X,Y}$ is the intercorrelation $\text{Cov}(X^{(k)}, Y)$.

Troisième partie III

Miscellaneous applications

Example 1

Example 2 : Call-put parity

Example 3 : Asian Option

Example 1

Objective

$$\mathcal{I} \stackrel{\text{def}}{=} \mathbb{E} \left[(W_1 + W_2)^{5/4} \right]$$

where W_1, W_2 are independent with the same Weibull distribution

$$f(x) = 3/2\sqrt{x} \exp(-x^{3/2}) \mathbb{1}_{\mathbb{R}^+}(x).$$

- 1 If U is uniform on $[0, 1]$, then $W = (-\ln U)^{2/3}$ is distributed as W_1 .
- 2 Estimate \mathcal{I} using a crude MC method.
- 3 Propose a control variate estimator of \mathcal{I} using $X = U_1 U_2$ as a control variate.

Example 1

Objective

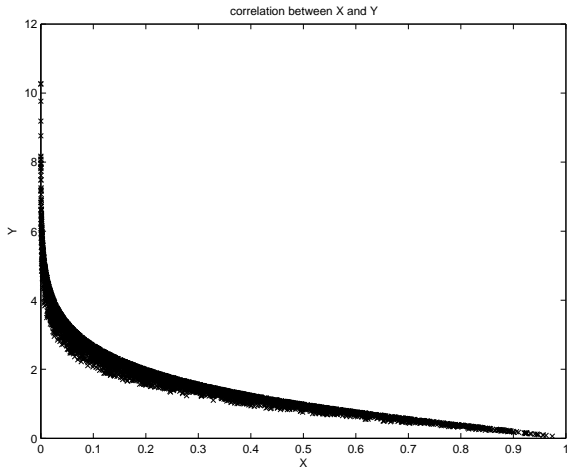
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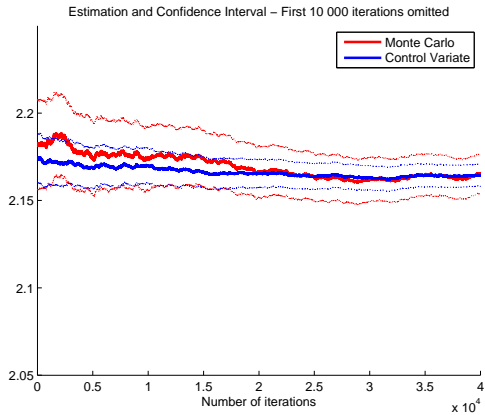
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- 3 Propose a control variate estimator of \mathcal{I} using $X = U_1 U_2$ as a control variate. in this case
- 4 On estimate : $\sqrt{1 - \rho_{X,Y}^2} \approx 0.55$.

Results



Results



Call put-parity

- For all x, K

$$(x - K)_+ - (K - x)_+ = x - K$$

- When $\mathbb{E}[S]$ is known, a natural control variable for $\mathbb{E}[(S - K)_+]$ is $X = (S - K)$.
- In financial engineering

$$\mathcal{I} \stackrel{\text{def}}{=} \mathbb{E} [\exp(-rT) (S_T - K)_+]$$

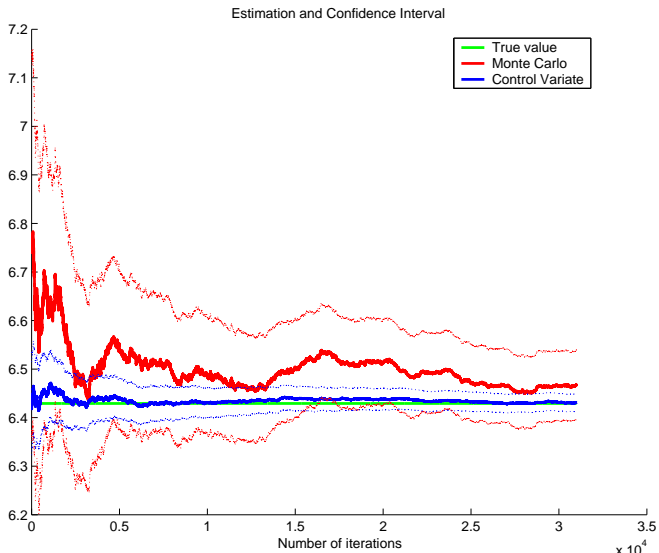
where S_T is the price of the underlying asset at the maturity T , modeled has a GBM $S_0 = x$ where

$$\mathbb{E} [\exp(-rT) (S_T - K)] = x - \exp(-rT)K$$

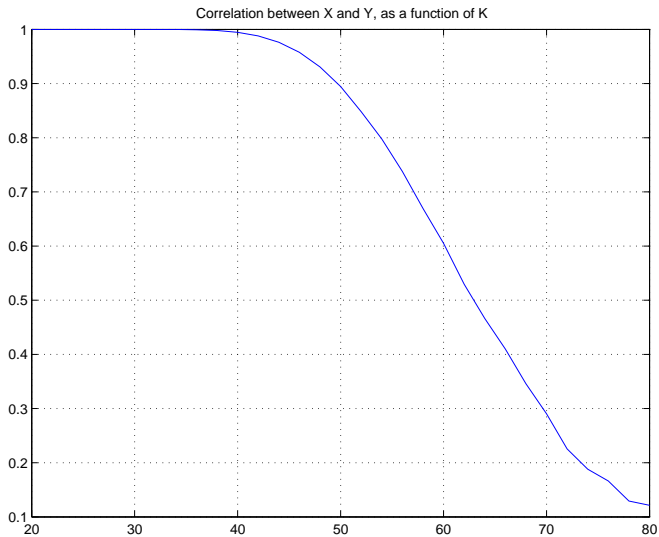
- In this case

$$Y = \exp(-rT) (S_T - K)_+ \qquad X = \exp(-rT) (S_T - K)$$

Numerical applications



Numerical applications



Asian Option

$$\mathcal{I} \stackrel{\text{def}}{=} \exp(-rT) \mathbb{E} \left[\left(M^{-1} \sum_{k=1}^M S_{t_k} - K \right)_+ \right]$$

where $t_k = kT/M$ and $\{S_t, t \geq 0\}$ is a GBM

$$S_t = S_0 \exp \left((r - 0.5\sigma^2)t + \sigma W_t \right).$$

Three possible control variates

$$X^{(1)} = \exp(-rT) S_T - S_0$$

$$X^{(2)} = \exp(-rT) (S_T - K)_+ - \mathbb{E} \left[\exp(-rT) (S_T - K)_+ \right]$$

$$X^{(3)} = \exp(-rT) \left(\exp \left(M^{-1} \sum_{k=1}^M \ln S_{t_k} \right) - K \right)_+ \\ - \mathbb{E} \left[\exp(-rT) \left(\exp \left(M^{-1} \sum_{k=1}^M \ln S_{t_k} \right) - K \right)_+ \right].$$

Numerical Applications

